

PASCAL MICHAILLAT
University of California, Santa Cruz

EMMANUEL SAEZ
University of California, Berkeley

$u^* = \sqrt{uv}$: *The Full-Employment Rate of Unemployment in the United States*

ABSTRACT This paper computes the unemployment rate u^* that is consistent with full employment in the United States. First, the paper argues that social efficiency is the most appropriate economic interpretation of the legal concept of full employment. Here efficiency means minimizing the nonproductive use of labor—both unemployment and recruiting. As it takes one worker to service one job vacancy, the nonproductive use of labor is measured by the number of job seekers and job vacancies, $u + v$. Through the Beveridge curve, the numbers of job seekers and vacancies are inversely related, $uv = \text{constant}$. With such symmetry the labor market is efficient when there are as many job seekers as vacancies ($u = v$), inefficiently tight when there are more vacancies than job seekers ($v > u$), and inefficiently slack when there are more job seekers than vacancies ($u > v$). Accordingly, the full-employment rate of unemployment (FERU) is the geometric average of the unemployment and vacancy rates: $u^* = \sqrt{uv}$. From 1930 to 2024, the FERU averages 4.1 percent and is stable, remaining between 2.5 percent and 6.7 percent. Unemployment has generally been above the FERU ($u > u^*$), especially during recessions. Unemployment has only been below the FERU ($u < u^*$) during major wars, as well as shortly before and in the aftermath of the pandemic.

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In the United States the federal government and central bank are mandated to maintain the economy at “full employment,” or “maximum employment.” This legislative mandate comes from the Employment Act of 1946, the Federal Reserve Reform Act of 1977, and the Full Employment and Balanced Growth Act of 1978 (Duboff 1977; Ginsburg 1979; M. Weir 1987; Steelman 2011; Bernanke 2013).¹ For instance, the Employment Act states that it is the “policy and responsibility of the federal government . . . to coordinate and utilize all its plans, functions, and resources . . . to promote maximum employment” (sect. 2). The Federal Reserve Reform Act of 1977 adds that it is the responsibility of the Federal Reserve “to promote effectively the goals of maximum employment, stable prices” (sect. 2A). Finally, the Full Employment and Balanced Growth Act of 1978 was written to “assert the responsibility of the Federal Government to use all practicable programs and policies to promote full employment” (preamble).² In this paper, we aim to compute the unemployment rate that characterizes a state of full or maximum employment. We denote it by u^* and, following Meade (1982), we refer to it as the full-employment rate of unemployment (FERU).

Our first task is to translate the legal notion of full employment into economic terms. Since the Employment Act and Full Employment and Balanced Growth Act clearly state that achieving full employment is a way to maximize social welfare, we translate full employment as social efficiency. Indeed, the Employment Act states that reaching full employment is designed “to foster . . . the general welfare” (sect. 2). The Full Employment and Balanced Growth Act adds that when the economy departs from full employment, it “is deprived of the full supply of goods and services, the full utilization of labor . . . and the related increases in economic well-being that would occur under conditions of genuine full employment” (sect. 2(a)(1)).

We therefore compute the FERU as the unemployment rate that achieves a socially efficient allocation of labor. This allocation maximizes social output by minimizing the uses of labor that are socially unproductive: both job seeking and recruiting. The goal is that workers spend as much time as

1. Employment Act of 1946, Pub. L. 79–304, 60 Stat. 23; Federal Reserve Reform Act of 1977, Pub. L. 95–188, 91 Stat. 1387; Full Employment and Balanced Growth Act of 1978, Pub. L. 95–523, 92 Stat. 1887. Records can be found at FRASER, <https://fraser.stlouisfed.org/>.

2. During the debate preceding the Employment Act, maximum employment was considered a less stringent goal than full employment (Duboff 1977). In 1978, the Full Employment and Balanced Growth Act amended the Employment Act and replaced maximum employment with the more ambitious target of full employment (M. Weir 1987).

possible producing socially useful things and waste as little time as possible searching for jobs or new hires. Of course, job seeking and recruiting are necessary for workers and firms to match with each other, but they do not generate any social welfare by themselves.

The FERU maximizes social output: goods and services produced in the market and at home that engender social welfare. In theory, unemployed workers might produce valuable goods and services at home while looking for jobs. But in practice, the benefits from home production are almost entirely offset by the psychological costs from being unemployed, so the social product of unemployed labor is minimal (Michaillat and Saez 2021a; Hussam and others 2022). Furthermore, not all employed workers produce social output. Many workers devote their time to recruiting instead of producing goods and services that add to social welfare. In fact, it takes about one full-time worker to service one job vacancy, so the number of recruiters can be counted by the number of vacancies (Gavazza, Mongey, and Violante 2018; Michaillat and Saez 2021a). Accordingly, the share of socially productive workers in the labor force is $1 - u - v$, where u is the unemployment rate and v is the vacancy rate. The FERU, therefore, minimizes the sum of the unemployment and vacancy rates, $u + v$.

A naive way to minimize $u + v$ would be to set the unemployment rate u and vacancy rate v to zero. But it is impossible to simultaneously reduce the numbers of job seekers and job vacancies because of the Beveridge curve. When the number of job seekers falls along the Beveridge curve, the number of vacancies necessarily rises; conversely, when the number of vacancies falls, the number of job seekers necessarily rises. In fact, the Beveridge curve is approximately a rectangular hyperbola: $uv = A$, where $A > 0$ is a constant (Michaillat and Saez 2021a). Hence, it is infeasible to set the unemployment and vacancy rates to zero, or even to reduce them simultaneously.

In sum, the efficient allocation minimizes $u + v$ subject to $uv = A$. Because of the symmetrical roles played by job seekers and vacancies, the efficient allocation must have as many job seekers as vacancies. This is equivalent to saying that the economy is at full employment when there are as many job seekers as vacancies ($u = v$). A further consequence is that the labor market is inefficiently tight when there are more vacancies than job seekers ($v > u$), and inefficiently slack when there are more job seekers than vacancies ($u > v$).

For policymakers seeking to communicate a single, clear indicator, the full-employment criterion can be expressed using labor market tightness, defined as the number of job vacancies per job seeker, v/u . Our analysis

shows that when tightness equals one, the economy is at full employment. When tightness exceeds one, the labor market is inefficiently tight. When tightness falls below one, it is inefficiently slack. Thus, tightness alone suffices to indicate whether the economy is at full employment, with the added advantage that the full-employment tightness takes an intuitive value: one.

Because we are used to thinking about unemployment rather than tightness, and because we have a better idea of the effects of stabilization policies on unemployment than on tightness, it is still useful to construct the rate of unemployment at full unemployment—the FERU. From the Beveridge curve and the equality of the efficient unemployment and vacancy rates, we deduce that the FERU is the geometric average of the unemployment and vacancy rates: $u^* = \sqrt{uv}$. As it only requires unemployment and vacancy rates, the FERU formula is easy to apply, even in real time. We derived a more general but also more complex formula in Michailat and Saez (2021a). Here we show that, empirically, the relevant parameters align so that the general formula can be greatly simplified. This provides an incredibly simple, easy-to-derive, and easy-to-use formula—which might be useful to policymakers.

Computing the FERU in the United States between 1930 and 2024, we find that the FERU averages 4.1 percent. The FERU is also quite stable: It remains between 2.5 percent and 6.7 percent, while the unemployment rate fluctuates between 1.0 percent and 25.3 percent.

Furthermore, the unemployment rate has generally been above the FERU, meaning that the US labor market has generally been inefficiently slack. The unemployment gap $u - u^*$ averages +2.3 percentage points. The gap is especially wide in recessions—as wide as +20.9 percentage points during the Great Depression and +5.9 percentage points during the Great Recession. The US labor market has only been inefficiently tight during major wars—World War II, the Korean War, and the Vietnam War—and around the coronavirus pandemic—from 2018:Q3 to 2020:Q1 and then from 2021:Q3 to 2024:Q2.

As the FERU formula can be applied in real time, we can use it to examine the US labor market during and after the coronavirus pandemic. We observe that the pandemic labor market has been extremely unusual. First, in 2020, the unemployment gap reached +6.3 percentage points. The last time the economy faced such slack was 1940, at the onset of World War II. Then, in 2022, the unemployment gap bottomed to −1.5 percentage points. The last time the economy became so tight was 1945, at the end of World War II.

I. Existing Unemployment Targets

Before beginning the analysis, we review unemployment targets used by US policymakers and argue that they do not align well with the US government's full-employment mandate.

I.A. Numerical Targets

In the early postwar period, right after the Employment Act established the full-employment mandate and created the Council of Economic Advisers (CEA) to enforce it, several numerical values were used as full-employment targets. From 1946 to 1956, the CEA used an unemployment rate of 3 percent as a marker of full employment (Duboff 1977). Then the CEA started raising their unemployment target. In 1962, the CEA wrote that an unemployment rate of 4 percent was "a reasonable and prudent full employment target for stabilization policy" (Duboff 1977, 10). Then, in 1969, Arthur Burns (1969, 280) reported that "since the [CEA] identified an unemployment rate of 4 percent with a condition of practically full employment, this figure served as a constant in the equation for computing the potential output."

A first issue with a numerical target is that it lacks a theoretical foundation. Hence, it is unclear what the target means or whether it accurately represents full employment. Policymakers recognized this limitation at the time. Even before becoming chairman of the Federal Reserve, Burns argued that the 4 percent target was not compelling because it did not incorporate information on job vacancies. For instance, Burns (1962, 17) wrote that "a serious need remains for strengthening the statistical machinery of the Employment Act. . . . In seeking to discriminate between structural changes as one possible cause of unemployment and deficiency of aggregate demand as another, we are still frustrated by an almost complete absence of statistics on job vacancies." Burns (1969, 284) added that "we need to develop comprehensive data on job vacancies, so that it will no longer be necessary to guess whether or when a deficiency in aggregate demand exists." In 2000, the US government started collecting data on job vacancies through the Job Openings and Labor Turnover Survey (JOLTS). In this paper, we combine data on job vacancies and unemployment to compute the unemployment rate consistent with full employment.

A second issue with a numerical target is that it is unclear when and how the target should change. Policymakers became aware of these limitations when the unemployment rate started rising in the 1970s. It was not clear whether the target should rise too, so the CEA moved away from a

numerical target for full employment. When testifying in front of Congress in 1975, Alan Greenspan, who was then chairing the CEA, was asked what the target for full employment was. He responded: “I do not think we should set a target” (Duboff 1977, 13).

I.B. NAIRU

In recent times, the US government has used the non-accelerating inflation rate of unemployment (NAIRU) as the full-employment target. For instance, the Joint Economic Committee (2019, 2) recently wrote that “today, full employment is considered by many to be synonymous with the non-accelerating inflationary rate of unemployment (NAIRU)—the rate of unemployment that neither stokes nor slows inflation.” Similarly, the CEA (2024, 24) described the concept of full employment as follows: “Modern economics has generally defined full employment by citing the theoretical concept of the lowest unemployment rate consistent with stable inflation, which is referred to as u^* , . . . the non-accelerating inflationary rate of unemployment (termed NAIRU).” These quotes are particularly meaningful because they come from the Joint Economic Committee and CEA, which were both created by the Employment Act of 1946 to ensure that the government achieved its employment mandate. Federal Reserve Chair Jerome Powell offered the same definition of full employment: “Most FOMC participants agree that labor market conditions are consistent with maximum employment in the sense of the highest level of employment that is consistent with price stability” (Federal Reserve Board 2022, 6).

The NAIRU is the unemployment rate at which inflation remains stable. It is measured by estimating Phillips curves (Staiger, Stock, and Watson 1997; Gordon 1997; Laubach 2001; Ball and Mankiw 2002; Orphanides and Williams 2002; Crump and others 2019).

Although the NAIRU contains information relevant to the Fed’s price-stability mandate, it does not represent the efficient rate of unemployment (Rogerson 1997). In modern models of the labor market, workers and firms meet through a matching function and form long-term employment relationships (Pissarides 2000). In these models, infinitely many real wages are acceptable in equilibrium (Hall 2005). However, only one of those wages yields the efficient rate of unemployment. There is no guarantee that the real wage arising under stable inflation coincides with this efficient real wage (Blanchard and Galí 2010). Accordingly, there is no guarantee that the unemployment rate prevailing under stable inflation—the NAIRU—is efficient. Since we have defined full employment as a socially efficient allocation of labor, the NAIRU cannot be a measure of full employment.

1.C. CBO's NRU

Another full-employment target used by the US government is the natural rate of unemployment (NRU)—which has been rebranded non-cyclical rate of unemployment since 2021—constructed by the Congressional Budget Office (CBO). For example, when he was president of the Federal Reserve Bank of Boston, Eric Rosengren (2014, 180) measured the departure of the Fed from its full-employment mandate by “the squared deviations of unemployment from an estimate of full employment utilizing the Congressional Budget Office assessment of the natural rate for each year.” Similarly, Powell (2018, 4) argues that policymakers should navigate using the natural rate of unemployment u^* as a guide; for instance, when “the unemployment rate is above u^* ,” the Federal Reserve should “lower the real federal funds rate . . . which will stimulate spending and raise employment.” To illustrate what u^* was from 1960 to 2000, and how it had fluctuated, Powell (2018, fig. 2) plots the CBO's NRU.

The CBO's NRU is a slow-moving trend of the unemployment rate computed by assuming that the labor market was at full employment in 2005 and then by incorporating changes in the demographic composition of the labor force over time (see appendix B in Shackleton 2018).

Although the NRU conveys information about the demographic forces exerted on the labor market, without a theory of full employment, it is impossible to know whether the US labor market really was at full employment in 2005, and by induction, whether the NRU in any year measures full employment. Thus, the CBO's NRU cannot be a satisfactory measure of full employment.

1.D. Daly and Others' (2012) NRU

Daly and others (2012) propose an alternative method to measure the NRU based on the Beveridge curve. They start from the empirical Beveridge curve, which they take as given. Then they estimate a long-run level of labor market tightness, which would prevail in the absence of business cycle shocks.³ Finally, they read the NRU at the intersection of the empirical Beveridge curve and the estimated long-run tightness (Daly and others 2012, fig. 4).

3. They actually estimate a long-run job creation curve. But in their theoretical framework, the job creation curve is just a line whose slope is labor market tightness (Daly and others 2012, fig. 1). So estimating a long-run job creation line is tantamount to estimating a long-run tightness.

The issue is that in the framework on which their analysis is based, there is no guarantee that long-run tightness is efficient, so there is no guarantee that this NRU is the efficient rate of unemployment (Pissarides 2000). As we have defined full employment as a socially efficient allocation of labor, the NRU cannot be a satisfactory measure of full employment. Instead, the NRU computed by Daly and others (2012) measures the noncyclical, structural rate of unemployment given the Beveridge curve.

I.E. Other Targets

In recent years, other full-employment targets have been developed to guide policymakers (Crump, Nekarda, and Petrosky-Nadeau 2020). These targets either guarantee stable prices, like the NAIRU, or reflect a slow-moving trend of unemployment, like the CBO's NRU and that of Daly and others (2012). They are not designed to measure the rate of unemployment that maximizes social welfare, as Crump, Nekarda, and Petrosky-Nadeau (2020) rightfully note, so they cannot satisfactorily measure full employment.

II. Derivation of the FERU Formula

Based on the legislation that introduced the full-employment mandate in the United States, we defined the FERU as the rate of unemployment that achieves a socially efficient allocation of labor. Therefore, the FERU is the solution to the problem of a social planner who allocates labor to maximize welfare. We now describe this problem and solve it to derive the FERU formula.

II.A. Social Welfare Function

The social planner allocates labor to maximize social output. Social output is the production of goods and services that generate social welfare. We have said that the social planner aims to maximize social welfare. But for simplicity, we leave out distributional considerations from the social welfare function, so social welfare is determined by social output.⁴ This perspective on full employment is consistent with the view expressed by Beveridge (1944, 20) that “the material end of all human activity is

4. Distributional considerations can be excluded by assuming that workers are risk neutral. If workers are risk averse and are not perfectly insured against unemployment, then the distribution of consumption influences welfare, and the efficient unemployment rate is given by a more complex formula that incorporates distributional elements (Landais, Michailat, and Saez 2018a).

consumption. Employment is wanted as a means to more consumption . . . as a means to a higher standard of life.”

II.B. Workers Available for Production

We assume that the social planner has the entire labor force at their disposal for production. This assumption aligns with the legislation behind the full-employment mandate, which intends to provide employment opportunities for all labor force participants. For instance, the Employment Act says that it aims to afford “useful employment opportunities, including self-employment, for those able, willing, and seeking to work” (sect. 2). The Full Employment and Balanced Growth Act uses similar language. Its goal is to “translate into practical reality the right of all Americans who are able, willing, and seeking to work to full opportunity for useful paid employment” (preamble). Thus, the labor force represents the pool of workers that the social planner can tap into for production. People out of the labor force may be in school or in training, may have retired, or may be looking after their family. They are not available to the planner for production.

Although the planner takes the labor force as given, they might have to account for changes in the size of labor force if that size systematically responds to the state of the labor market. In practice, however, the labor force participation rate is acyclical, so the planner takes the labor force size as fixed. Using US data covering 1946–1954, Rees (1957) does not find evidence of the discouraged-worker theory. More systematically, in US data covering 1960–2006, Shimer (2009) finds that the labor force participation rate is acyclical. Similarly, using US data spanning 1976–2009, Rogerson and Shimer (2011, 624–25) find that, over the business cycle, “the labor force participation rate is nearly constant.” Erceg and Levin (2014) also find that the labor force participation rate is acyclical in the United States between 1972 and 2007.⁵ Finally, using a vector autoregression run on US data for 1976–2016, Cairo, Fujita, and Morales-Jimenez (2022, fig. 1C) find that the impulse response of the labor force participation rate to a positive productivity shock (the typical shock in business cycle models) is zero for two years, and while it is slightly positive after two years, it is never significantly different from zero.⁶

5. Erceg and Levin (2014) argue that high unemployment during the Great Recession caused a drop in labor force participation. But as Aaronson and others (2014) and Krueger (2017) show, the decline in labor force participation was primarily caused by population aging and other trends that preceded the Great Recession.

6. In fact, in section IV.A we show that the FERU formula is not modified when we endogenize the labor force participation rate and allow it to respond to labor market conditions.

II.C. Social Product of Employed Labor

We have said that the planner has the entire labor force at their disposal for production. Among those are workers employed by firms and job seekers. We start by assessing the social product of employed workers.

Employed workers must spend some of their time recruiting new hires for their firms, so they are unable to spend their entire time contributing to social output. Recruiting takes work: designing and advertising job vacancies, screening and interviewing candidates, and negotiating contracts. Beside recruiting, employed workers might also spend time looking for a new job, which takes further time away from socially productive tasks.

There are two sources of information about the amount of labor devoted to recruiting in the United States. The first source is the National Employer Survey, which was conducted by the Census Bureau in 1997 (Villena Roldán 2010). The survey asked thousands of establishments across industries about their recruiting practices (Cappelli 2001). Using the survey, Michaillat and Saez (2021a) estimate that servicing a job vacancy requires 0.92 workers at any point in time.

A second source is the survey conducted by the consulting firm Bersin and Associates in 2011 (Gavazza, Mongey, and Violante 2018). The survey asked over four hundred firms with more than one hundred employees about their spending on all recruiting activities. Gavazza, Mongey, and Violante (2018) find that recruiting one worker costs 0.928 of a monthly wage. To translate this number into the labor cost of servicing a job vacancy, we assess the time it took to fill a vacancy in 2011. On an average month in 2011, there were 4.305 million hires and 3.430 million vacancies, as measured by the Bureau of Labor Statistics (BLS 2024c, 2024f). Therefore, vacancies were filled at a monthly rate of $q = 4.305/3.430 = 1.26$, and vacancies stayed open on average for $1/q = 1/1.26 = 0.79$ months. These results imply that it takes $0.928/0.79 = 1.17$ workers to service a vacancy.

The two surveys show that it takes about one full-time worker to service a job vacancy—maybe a bit less or maybe a bit more.⁷ In other words, the number of recruiters in the United States is well measured by the number of vacancies. So the number of workers diverted from producing and allocated to recruiting can be measured by the number of vacancies open at any point in time.

Employed workers might also be distracted from producing if they search for new jobs at work. However, the average time spent on job search

7. In section IV.B we show how to extend the FERU formula if the number of recruiters per vacancy is different from one.

by employed workers is only 31 seconds per day (Ahn and Shao 2021, table 1). So on-the-job search is a tiny amount taken away from production, and we abstract from it here.

II.D. Social Product of Unemployed Labor

Next, we assess the social product of unemployed labor. We consider three possible activities for unemployed workers. The first is looking for a job. Job seeking is required to find employment but—just like recruiting—it does not contribute to social output. The second is producing goods and services at home. Home production adds to social output and contributes to social welfare. The third is remaining idle when not looking for jobs or producing at home.

The value of job seekers' home production, net of the psychological cost of idleness, is estimated to be negligible. Using administrative data from the US military, Borgschulte and Martorell (2018) study how service members choose between reenlisting and leaving the military. The choices allow them to estimate the value of home production plus public benefits minus the psychological cost of idleness during unemployment. Subtracting the value of public benefits from these estimates, Michailat and Saez (2021a) find that the value of home production minus the psychological cost of idleness could be as low as 3 percent of the value of market production.

Given its minimal value, we set the social product of unemployed labor to zero.⁸ The unproductivity of unemployment was already noted by Robinson (1949, 11): “The most important aspect of unemployment is its wastefulness. It is the existence of unused productive resources side by side with unsatisfied human needs that is the intolerable condition.”

Where do the psychological costs of unemployment come from? The psychological costs associated with unemployment arise from various sources. First, depression, anxiety, and strained personal relations are common consequences of job loss (Eisenberg and Lazarsfeld 1938; Theodossiou 1998). Job loss is a traumatic event that can lead to a decline in an individual's self-esteem and sense of self-worth (Goldsmith, Veum, and Darity 1996). Joblessness also diminishes psychological well-being by creating a sense of helplessness: that one's life is no longer under their control (Goldsmith and Darity 1992). Furthermore, job search appears to reduce unemployed workers' life satisfaction (Krueger and Mueller 2011). In fact, Jahoda (1981) emphasizes numerous important benefits of work—which

8. In section IV.B we show how to extend the FERU formula if the social product of unemployed labor is nonzero.

are lost during unemployment. These benefits from work encompass a structured daily routine, regular interactions and shared experiences with individuals beyond the immediate family, the pursuit of overarching goals and purposes, a source of personal status and identity, and the engagement in regular activities. Collectively, the loss of these benefits contributes to the psychological burdens associated with unemployment.

That the idleness associated with unemployment can create psychological hardship goes against the idea—standard in economics—that unemployed workers enjoy leisure time. Yet, even though it is often neglected in economics, the psychological toll from unemployment has been understood for a long time. Robinson (1949, 11) for instance observed that “the most striking aspect of unemployment is the suffering of the unemployed and their families—the loss of health and morale that follows loss of income and occupation.” At this point, the detrimental effects of unemployment on mental and physical health are well documented (Dooley, Fielding, and Levi 1996; Platt and Hawton 2000; Frey and Stutzer 2002; Wanberg 2012; Brand 2015).

A field experiment in Bangladesh by Hussam and others (2022) illustrates just how large the psychosocial cost of unemployment is. The experiment shows that paid employment raises psychosocial well-being substantially more than the same amount of cash alone. In fact, two-thirds of employed workers would be willing to forgo cash payments and continue working for free.

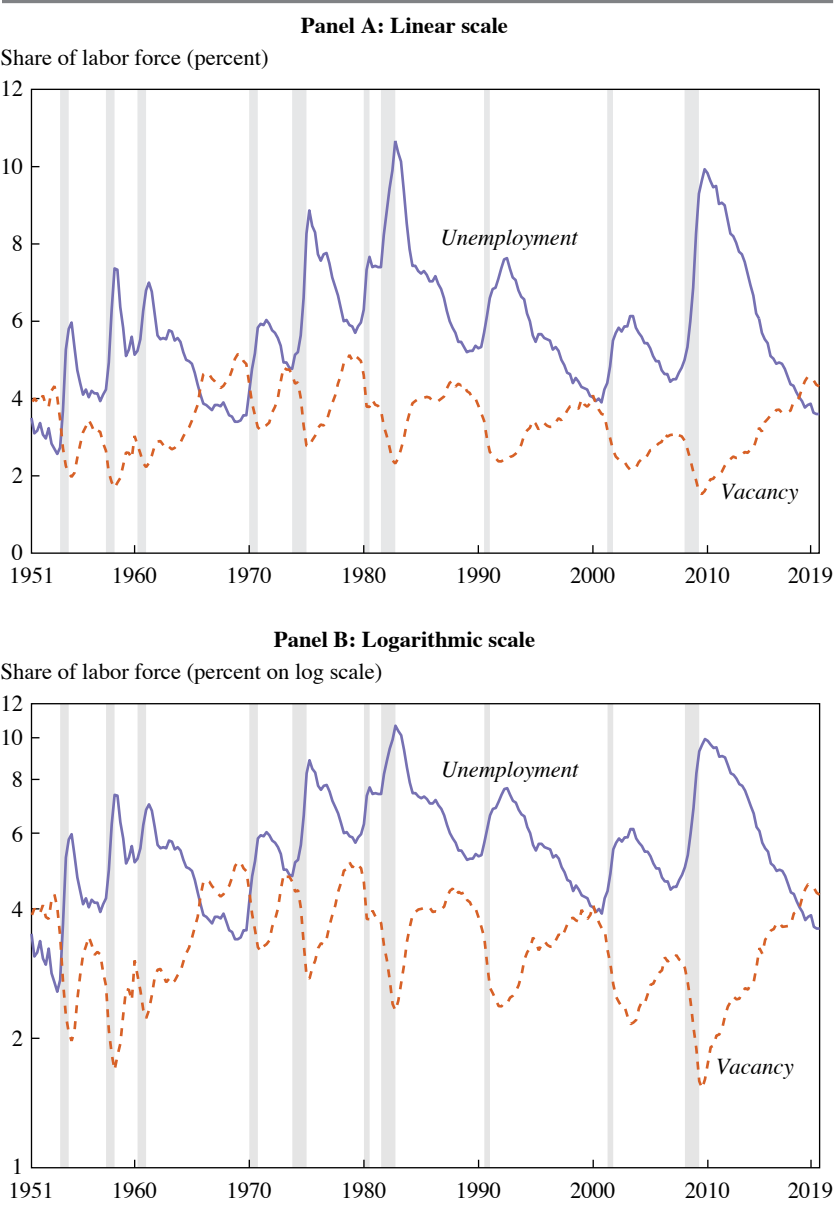
II.E. Shape of the Beveridge Curve

Given that both unemployed workers and vacant jobs are socially costly, the social planner would want to reduce both. This is not feasible, however, because of the Beveridge curve, which imposes that the numbers of job seekers and job vacancies are negatively related. When the economy is in a slump, there are a lot of job seekers and few vacancies. Conversely, when the economy is in a boom, there are few job seekers and many vacancies.

Looking at labor market data for Great Britain, Beveridge (1944) first observed that the numbers of job vacancies and job seekers move in opposite directions. Dow and Dicks-Mireaux (1958, figs. 1 and 2) confirmed Beveridge’s observation by plotting unemployment and vacancy data for Great Britain, 1946–1956. The data reveal that, over time, the unemployment rate increases whenever the vacancy rate declines, and vice versa.

The Beveridge curve holds remarkably well in the United States (Blanchard and Diamond 1989; Elsby, Michaels, and Ratner 2015). Figure 1, panel A, depicts the unemployment and vacancy rates in the

Figure 1. Unemployment and Vacancy Rates in the United States, 1951:Q1–2019:Q4



Source: The unemployment rate is measured by BLS (2024k). Between 1951:Q1 and 2000:Q4, the vacancy rate is constructed by Barnichon (2010); between 2001:Q1 and 2019:Q4, the vacancy rate is the number of job openings divided by the civilian labor force, both measured by BLS (2024a, 2024f).

Note: Unemployment and vacancy rates are quarterly averages of monthly series. The vertical gray areas are recessions dated by the National Bureau of Economic Research (NBER 2023).

United States between 1951 and 2019. The unemployment rate is the number of job seekers divided by the size of the labor force. The vacancy rate is the number of vacancies divided by the size of the labor force. The figure shows that unemployment and vacancy rates move in opposite directions: The unemployment rate is sharply countercyclical, while the vacancy rate is clearly procyclical. The unemployment and vacancy data plotted here are produced by BLS (2024a, 2024f, 2024k) and Barnichon (2010), and they are widely used (Daly and others 2012; Diamond and Şahin 2015; Elsby, Michaels, and Ratner 2015; Barnichon and Figura 2015; Ahn and Crane 2020; Petrosky-Nadeau and Zhang 2021; Barlevy and others 2024; Michaillat and Saez 2021a, 2024b). We will come back to the construction of these data in section III.A.

In fact, unemployment and vacancy rates appear not only to be negatively related but to be the inverse of each other. So doubling the unemployment rate cuts the vacancy rate in half, and conversely, doubling the vacancy rate cuts the unemployment rate in half. Figure 1, panel B, displays again unemployment and vacancy rates, but now on a logarithmic scale. The fluctuations of the unemployment and vacancy rates are close to a mirror image of each other, indicating that unemployment and vacancy rates are inversely related.

Mathematically, the property that the unemployment rate $u \in [0, 1]$ and vacancy rate $v \in [0, 1]$ are inversely related implies that the Beveridge curve is a rectangular hyperbola:

$$vu = A,$$

where $A \in (0, 1/4)$ is a constant.⁹

We can formally establish that the Beveridge curve is a rectangular hyperbola by estimating the elasticity of the vacancy rate with respect to the unemployment rate, $d \ln(v)/d \ln(u)$. An elasticity of -1 corresponds to a hyperbola. Using the algorithm developed by Bai and Perron (1998) and the data displayed in figure 1, Michaillat and Saez (2021a, figs. 5 and 6) estimate the structural breaks of the US Beveridge curve and the elasticity of the

9. We impose the condition $A < 1/4$ so the equation $vu = A$ admits at least a solution (u, v) such that $u + v \leq 1$. The condition $u + v \leq 1$ must hold because the number of job seekers and recruiters is less than the number of labor force participants. To see where the upper bound $1/4$ comes from, consider the point on the Beveridge curve such that $u = v$. That point satisfies $u^2 = A$ or $u = \sqrt{A}$, and $v = u = \sqrt{A}$. The constraint $u + v \leq 1$ translates into $2\sqrt{A} \leq 1$, which is equivalent to $A \leq 1/4$. By imposing $A < 1/4$, we ensure that parts of the Beveridge curve satisfy the constraint $u + v \leq 1$.

Beveridge curve between these breaks. They find that over the 1951–2019 period, the Beveridge elasticity remains between -0.84 and -1.02 , so never far from -1 . This finding confirms that the US Beveridge curve is close to a rectangular hyperbola.¹⁰

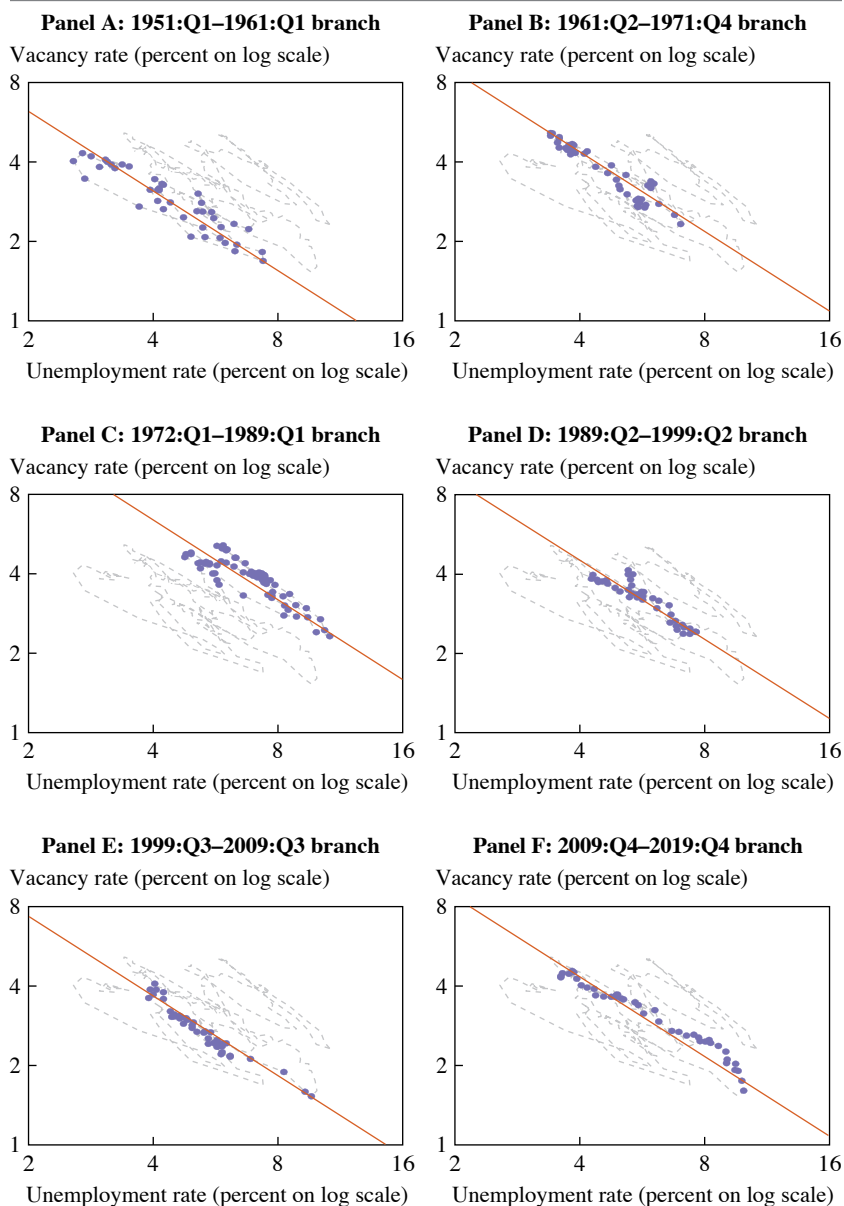
Graphically, it is evident that the US Beveridge curve closely resembles a rectangular hyperbola. The branches of the Beveridge curve identified by Michaillat and Saez (2021a, fig. 5) are plotted in figure 2. The US labor market typically stays on one branch for a decade or more before the Beveridge curve abruptly shifts to a new location (a sudden change in A). In each panel, the solid straight line represents a rectangular hyperbola. Since the panels plot the unemployment and vacancy rates on logarithmic scales, the hyperbola appears as a downward-sloping line with a slope of -1 . Across all panels, the Beveridge curve aligns closely with the rectangular hyperbola.

It is quite natural that the Beveridge curve takes the shape of a rectangular hyperbola, since it is the shape that arises in a basic matching model of the labor market. In the model, the Beveridge curve is the locus of points such that labor market flows are balanced: The number of workers who lose or quit their jobs equals the number of workers who find a job. The employment rate $1 - u$ is approximately constant at one since the unemployment rate u is an order of magnitude less than one. The job separation rate λ is also constant, so the number of job separations $\lambda(1 - u)$ is approximately constant at λ . So along the Beveridge curve, the number of workers who find a job is constant at λ . With the standard symmetric Cobb-Douglas matching function, $m = \omega \sqrt{uv}$, the number of workers who find a job at any point in time is proportional to \sqrt{uv} .¹¹ Hence, along the Beveridge curve, \sqrt{uv} and thus uv must be constant: The Beveridge curve is a rectangular hyperbola.

We have just provided a foundation for the hyperbolic Beveridge curve based on a basic matching model, but the analysis is in no way limited to that model. Our analysis only presumes that the Beveridge curve exists—it does not put additional restrictions on the structure of the labor market. For instance, the basic matching model only features labor flows between employment and unemployment. In practice, there are vast labor flows in and out of the labor force, and from employment to employment (Blanchard

10. In section IV.B we show how to extend the FERU formula if the Beveridge curve is an isoelastic curve with an elasticity different from -1 .

11. The US matching function appears to have a Cobb-Douglas form with exponents of 0.5 on unemployment and vacancies (Michaillat and Saez 2021a).

Figure 2. The US Beveridge Curve Approximates a Rectangular Hyperbola

Source: Unemployment and vacancy rates come from figure 1. The structural breaks between branches of the Beveridge curve are estimated by Michaillat and Saez (2021a, fig. 5) using the algorithm of Bai and Perron (1998).

Note: The solid, straight lines are rectangular hyperbolas $uv = A$, where the constant A is specific to each branch.

and Diamond 1990, fig. 1). Our analysis applies to all models with such labor flows as long as they feature a Beveridge curve.¹²

Similarly, the basic matching model assumes that firms recruit workers only from unemployment. In reality, firms also recruit workers from other employers and from outside the labor force. Fortunately, our analysis extends to more sophisticated labor market models, where hires come from various sources. The only requirements are that vacancies reflect firms' recruiting effort and that vacancies are related to unemployment through a Beveridge curve. The only relevant consideration for welfare is that when unemployment falls, firms allocate more resources to recruitment.

Finally, we assume that the labor market is always on the Beveridge curve. A potential concern is that labor market dynamics outside of the Beveridge curve may be important. Indeed, in matching models, unemployment evolves through a dynamic process driven by the difference between inflows into unemployment and outflows from unemployment, so unemployment is not always on the Beveridge curve. What can allay this concern is that in the United States, the inflows and outflows are extremely large, so unemployment dynamics converges extremely quickly to the Beveridge curve. Michailat and Saez (2021a) show that 50 percent of the deviation of the US unemployment rate from the Beveridge curve evaporates within one month, and 90 percent within one quarter. Thus, the US unemployment rate is always near the Beveridge curve. This explains why many matching models assume that the Beveridge curve holds at all times, as we do here (Hall 2005; Pissarides 2009; Landais, Michailat, and Saez 2018b; Michailat 2024).

II.F. Full-Employment Criterion

Using the social product of employed and unemployed labor and the shape of the Beveridge curve, we now formally describe and solve the social planner's problem. The solution to the planner's problem will give us the full-employment criterion based on unemployment and vacancy rates.

The planner aims to maximize social output by minimizing the sum of the unemployment and vacancy rates, $u + v$. Indeed, since unemployment and recruiting are socially wasteful and the labor force is given,

12. An implicit assumption is that all workers have the same productivity across all firms. Therefore, job-to-job and labor force transitions do not affect the output of transitioning workers or overall welfare. (Since the labor force has constant size, any worker exiting the labor force is replaced by a new worker entering it. For example, a worker going on parental leave is replaced by one returning from parental leave.)

maximizing social output is tantamount to minimizing labor in unemployment or recruiting. And since the number of recruiters can be counted by the number of vacancies, the objective is to minimize the number of job seekers plus vacancies. Equivalently, the labor force size being fixed, the objective is to minimize the unemployment rate plus the vacancy rate.

This minimization is subject to the Beveridge curve constraint, $uv = A$. Because of the Beveridge curve, it is not possible to reduce unemployment and vacancies at the same time, so the planner must trade off unemployment and vacancies. The planner takes the Beveridge curve as given because the Beveridge curve does not seem to respond to monetary or fiscal stabilization policy. Indeed, in many business cycle models with unemployment, the Beveridge curve is unaffected by monetary and fiscal policy (Blanchard and Galí 2010; Ravenna and Walsh 2011; Michaillat 2014; Michaillat and Saez 2019, 2022, 2024a). In these models the Beveridge curve is determined by the matching function and job separation rate. Neither responds to monetary or fiscal policy, so the Beveridge curve is unaffected by policy.¹³

The planner minimizes nonproduction $u + v$ subject to the Beveridge curve $uv = A$, with $u \in [0, 1]$ and $v \in [0, 1]$. To simplify the problem, we substitute the Beveridge curve, $v = A/u$, into the objective function. Then the problem simply is to minimize $u + A/u$ over $u \in [A, 1]$.¹⁴ The function $u \mapsto u + A/u$, defined over the interval $[A, 1]$, is continuous and strictly convex.¹⁵ Therefore, the function admits a unique minimum on $[A, 1]$.

As we have just seen, the minimization of $u + v$ subject to $uv = A$, with $[u, v] \in [0, 1]^2$ admits a unique solution. In addition, the minimization

13. Other policies do influence the Beveridge curve. For example, reducing unemployment insurance bolsters job seekers' search efforts, which shifts the Beveridge curve inward (Landais, Michaillat, and Saez 2018b; Hochmuth and others 2021). The effect of such policies on welfare can be split into two components (Landais, Michaillat, and Saez 2018a). The first component is the direct effect on welfare, assuming labor market tightness remains fixed. This includes the costs and benefits associated with shifting the Beveridge curve. The second component is the effect on welfare through tightness, which is the product of the effect of tightness on welfare (holding the policy constant) and the effect of the policy on tightness. The effect of tightness on welfare holds the policy constant, so it leaves the Beveridge curve unchanged, and it can be computed just as the effect of unemployment on welfare in this paper. Consequently, the unemployment and tightness gaps derived here remain central to optimal policy design, though they might need to be supplemented by additional elements specific to the policy in question.

14. With $u \in [A, 1]$, we ensure that $v = A/u$ is in $[0, 1]$. In fact, $v \in [A, 1]$, just like u .

15. To see that the function is strictly convex, note that its second derivative is strictly positive: $2A/u^3 > 0$ for any $u > 0$.

problem is perfectly symmetric in u and v . Therefore, the minimum must be reached when $u = v$. To see why, imagine that the minimum was reached for $u = u_0$ and $v = v_0$ with $u_0 \neq v_0$. Then, because of the symmetry of the problem, setting the unemployment and vacancy rates to $u = v_0$ and $v = u_0$ would also minimize the objective function while respecting all the constraints. So the solution to the minimization problem would not be unique. We reach a contradiction here, which means that it cannot be that $u_0 \neq v_0$.

That is, full employment prevails when the unemployment and vacancy rates are equal ($u = v$). When they are not equal, the labor market is operating inefficiently. The labor market is inefficiently tight when there are more vacancies than job seekers ($v > u$). In that case, increasing u and reducing v would increase social output. The labor market is inefficiently slack when there are more job seekers than vacancies ($u > v$). Then, reducing u and increasing v would increase social output.

We can also solve the planner's problem by first-order condition. Recall that the planner aims to minimize $u + A/u$ over $u \in [A, 1]$. Since the function $u \mapsto u + A/u$ is strictly convex, the first-order condition is sufficient to find the function's minimum over the interval $[A, 1]$. We take the function's derivative with respect to u and set it to zero. We obtain $1 - A/u^2 = 0$, or equivalently $u = \sqrt{A}$. We verify that $\sqrt{A} \in [A, 1]$, because $0 < A < 1$. Therefore, the function's minimum occurs when $u = \sqrt{A}$. By the Beveridge curve we have $v = A/u$, so at the minimum $v = A/\sqrt{A} = \sqrt{A}$.¹⁶ Accordingly, at full employment, the unemployment and vacancy rates are equal and satisfy

$$(1) \quad u^* = v^* = \sqrt{A}.$$

Equation (1) shows that full employment occurs when unemployment and vacancy rates are equal. The equation also shows that the location of the Beveridge curve, A , solely determines these rates at full employment. This location summarizes everything we need to know for our welfare analysis—it serves as the key sufficient statistic (Chetty 2009). In basic matching models the Beveridge curve's position is determined by the job separation rate and the efficacy of the matching function. Any change in

16. Technically, because the number of job seekers and recruiters cannot exceed the number of labor force participants, the planner's problem should include the constraint $u + v \leq 1$. But the constraint is satisfied at the minimum, so it does not alter the problem's solution. Indeed, we have $A < 1/4$, so $\sqrt{A} < 1/2$, which implies that at the minimum, $u + v = 2 \times \sqrt{A} < 1$.

either parameter shifts the curve, affecting the FERU. However, which parameter causes the shift is irrelevant; only the shift itself matters for welfare.

We have expressed the full-employment criterion in terms of two separate variables: unemployment rate u and vacancy rate v . But we can reformulate the full-employment criterion in terms of one single variable: labor market tightness v/u . Tightness represents the number of vacancies per job seeker. It is a core variable in matching models of the labor market (Pissarides 2000; Shimer 2005; Hall 2005; Michaillat 2012). We have seen that the economy is at full employment when $v = u$, so it is at full employment when tightness equals one. The economy is inefficiently tight when $v > u$, so when tightness exceeds one. Finally, the economy is inefficiently slack when $v < u$, so when tightness falls below one.

II.G. FERU Formula

Although we have established that tightness at full employment is one, it is still useful to construct the rate of unemployment at full unemployment—the FERU. This is because researchers and policymakers more commonly think about unemployment than about tightness and because the effects of stabilization policies on unemployment are better understood than those on tightness (Ramey 2013, 2016).

To derive an expression for the FERU, we start from equation (1) and substitute A out of it by using the Beveridge curve $A = uv$. We find that the FERU is the geometric average of the unemployment and vacancy rates:

$$(2) \quad u^* = \sqrt{uv}.$$

Since $uv = A > 0$, expression (2) implies that the FERU is strictly positive. Hence, full employment should not be interpreted as zero unemployment.

A first reason why full employment does not mean zero unemployment is that zero unemployment is infeasible. Indeed, the Beveridge curve prevents unemployment from ever reaching zero. Because each vacancy requires a recruiter, the vacancy rate v is at most one. Accordingly, the Beveridge curve $u = A/v$ prevents the unemployment rate to fall below $A > 0$.

The fact that labor market flows impose a minimum level of unemployment—and therefore that full employment cannot be zero unemployment—has been known for a long time. Beveridge (1944, 125) realized that “however great the unsatisfied demand for labour, there is an irreducible minimum of unemployment, a margin in the labour force required to make change and movement possible”; as a result, “even under full employment, there

will be some unemployment. . . . On each day some men able and willing to work will not be working.” Robinson (1946, 169–70) made the same observation: “In a changing world there are always bound to be, at any moment, some workers who have left one job and have not yet found another. . . . Changes in occupation for personal reasons will always be going on. So long as such shifts in employment are taking place there is always likely to be some unemployment even when the general demand for labour is very high.”

A second reason why full employment does not mean zero unemployment is that zero unemployment is undesirable. Unemployment is clearly a waste of economic resources as people who would like to work are not able to be productive. Yet, reducing the unemployment rate to zero is not desirable because it would require diverting a vast amount of labor toward recruiting. In fact, it is not efficient to reduce the unemployment rate below the vacancy rate. Reducing the unemployment rate by 1 percent requires raising the vacancy rate by 1 percent, due to the hyperbolic Beveridge curve. When the unemployment rate is less than the vacancy rate, the increase in vacancy rate is more than the decrease in unemployment rate. Hence, overall, although the unemployment rate falls, the sum of the unemployment and vacancy rates increases—which means that social output falls.¹⁷

II.H. Application to the Diamond-Mortensen-Pissarides Model

We now apply our approach to the most common model of the labor market: the Diamond-Mortensen-Pissarides (DMP) model (Diamond 1982; Mortensen 1982; Pissarides 1985). The concept of efficiency used here is the same as in the DMP model. The model features both unemployed workers and job vacancies, each inducing output losses. More unemployment means fewer people at work so less output; more vacancies mean more labor devoted to recruiting and less output. The efficient allocation maximizes output by minimizing the combined losses from unemployment and recruiting. As the DMP model features a Beveridge curve, our results easily apply.

17. Zero unemployment is not desirable here because of the resources absorbed by recruiting. Robinson (1946, 170) agreed that “no-one regards 100% employment as a desirable objective.” Her logic was different, however. She argued that “the attainment of full employment, in this absolute sense, would require strict controls, including direction of labour” and that it would “involve great sacrifices of liberty,” even the “complete conscription of labour.”

We consider the model presented by Pissarides (2000, chap. 1). The labor force is composed of $L > 0$ workers with linear utility function. The $1 - u$ employed workers have a productivity $p > 0$. The u unemployed workers engage in home production and their productivity is $z < p$. And firms incur a flow recruiting cost $pc > 0$ for each vacancy. Hence, flow social welfare is

$$\left[p(1 - u) + zu - pcv \right] L.$$

We have argued that in the United States, it is accurate to set $z = 0$ and $c = 1$. Hence, flow welfare simplifies to

$$(3) \quad p \left[1 - (u + v) \right] L.$$

Maximizing flow welfare (3) is equivalent to minimizing $u + v$.

Next, we turn to the Beveridge curve. The Beveridge curve is the locus of points such that the number of workers who lose or quit their jobs equals the number of workers who find a job. The job separation rate is λ , so the number of workers who lose or quit their jobs is $\lambda(1 - u)$. With the standard symmetric Cobb-Douglas matching function, the number of workers who find a job is $m = \omega \sqrt{uv} = (\omega \sqrt{v/u}) u$. Along the Beveridge curve, $\lambda(1 - u) = (\omega \sqrt{v/u}) u$, so the Beveridge curve satisfies

$$u = \frac{\lambda}{\lambda + \omega \sqrt{v/u}}.$$

However, in the United States the job separation rate, λ , is more than an order of magnitude smaller than the job-finding rate, $\omega \sqrt{v/u}$ (Barnichon and Shapiro 2024). Therefore, the Beveridge curve can be approximated by $u = \lambda/(\omega \sqrt{v/u})$, which is a rectangular hyperbola:

$$(4) \quad uv = \left(\frac{\lambda}{\omega} \right)^2.$$

Formally, because the DMP model is dynamic, the social planner maximizes the present-discounted sum of flow social welfare, subject to the law of motion of unemployment (Pissarides 2000). To simplify, we follow Hosios (1990) and assume that the discount rate is zero. Under this assumption, the

social planner maximizes steady-state welfare. That is, the planner maximizes flow welfare (3) subject to the Beveridge curve (4). Equivalently, the planner minimizes $u + v$ subject to $uv = (\lambda/\omega)^2$. This is the exact problem studied here, so all the results apply: Efficient unemployment and vacancy rates are $u^* = v^* = \sqrt{uv} = \lambda/\omega$; efficient tightness is $v^*/u^* = 1$.¹⁸

III. FERU in the United States

We compute the FERU in the United States across three distinct periods: the standard postwar era (1951–2019), the Great Depression and World War II (1930–1950), and the coronavirus pandemic (2020–2024). We find that generally the US economy is not at full employment but is inefficiently slack.

III.A. Postwar Period

We first focus on the postwar period, 1951–2019. This is a standard period in the macro-labor literature, for which the unemployment and vacancy data are well known and well understood (Shimer 2005, 2007; Daly and others 2012; Diamond and Şahin 2015; Michailat and Saez 2021a). We stop at the end of 2019 to avoid incorporating the pandemic, which is an extremely unusual period that we will discuss in section III.C.

The unemployment rate u and vacancy rate v used in our analysis are displayed in figure 1 above. The unemployment rate is constructed by BLS (2024k) from the Current Population Survey (CPS). This is the standard, official measure of unemployment, labelled U3 by BLS (2023). This measure only includes job seekers who want a job, are available to start a job, and have been actively searching for a job in the past four weeks.¹⁹

The vacancy rate is derived from two different sources because there is no continuous vacancy series over the period. For 1951–2000, we use the vacancy rate constructed by Barnichon (2010). This series is based on the Conference Board's help wanted advertising index, adjusted to account

18. The Hosios (1990) condition gives the bargaining power required for the market unemployment rate to coincide with the efficient unemployment rate under Nash bargaining. Instead, we determine the unemployment rate that prevails when the labor market operates efficiently.

19. In section IV.C we repeat the analysis with two broader measures of unemployment that include job seekers with lower search effort: U4 and U5. These measures add to U3 people who want a job, are available to start a job, and have been actively searching for a job in the past twelve months but not in the past four weeks.

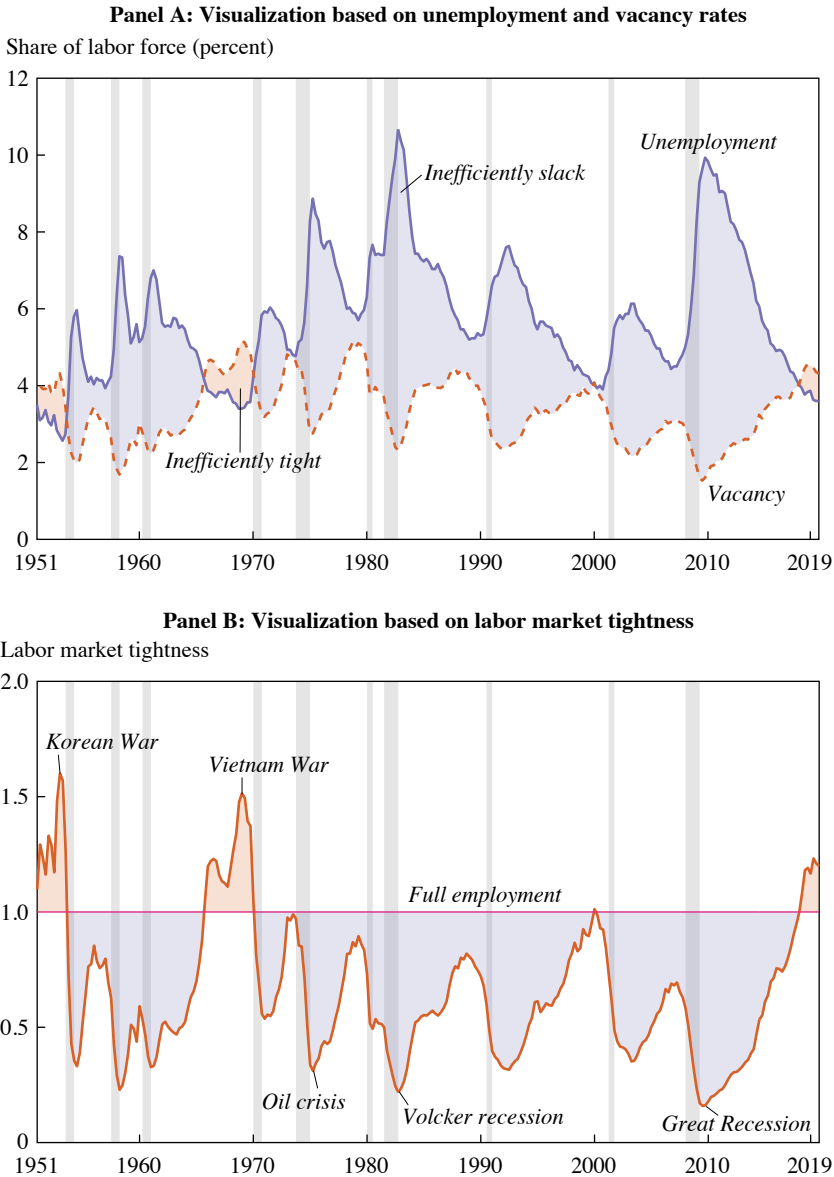
for the shift from print advertising to online advertising in the 1990s. The Conference Board index aggregates help wanted advertising in major metropolitan newspapers in the United States. It serves as a reliable proxy for job vacancies (Abraham 1987; Shimer 2005). For 2001–2019, we use the number of job openings measured by BLS (2024f) from the JOLTS, divided by the civilian labor force constructed by BLS (2024a) from the CPS.²⁰ We then splice the two series to create a continuous vacancy rate for 1951–2019. The two series are perfectly aligned because Barnichon (2010) used the JOLTS data to scale the Conference Board index so as to translate it into a vacancy rate (which was possible because the Conference Board and JOLTS series overlap in the early 2000s).

Next, we use the unemployment and vacancy rates to assess the state of the US labor market between 1951 and 2019 (figure 3, panel A). The labor market is inefficiently slack whenever the unemployment rate is above the vacancy rate; it is inefficiently tight whenever the unemployment rate is below the vacancy rate. Over the period, the unemployment rate averages 5.8 percent, while the vacancy rate only averages 3.4 percent. So on average, the unemployment rate is markedly higher than the vacancy rate, which indicates that the labor market is inefficiently slack. In fact, the labor market is persistently inefficiently slack except in three episodes when it turns inefficiently tight: the Korean War (1951:Q1–1953:Q3), the Vietnam War (1965:Q4–1970:Q1), and the end of the first Trump presidency (2018:Q2–2019:Q4).

The state of the US labor market can also be visualized by plotting labor market tightness v/u (figure 3, panel B). The labor market is inefficiently slack whenever tightness is below one, inefficiently tight whenever tightness is above one, and at full employment when tightness equals one—when there is one vacancy per job seeker. Tightness averages 0.65 between 1951 and 2019, well below one, which is another manifestation that the labor market is inefficiently slack on average. Tightness peaked at 1.60

20. To best align vacancy and labor force data, we shift forward by one month the number of job openings from JOLTS. For instance, we assign to December 2023 the number of job openings that the BLS assigns to November 2023. The motivation for this shift is that the number of job openings from the JOLTS refers to the last business day of the month (Thursday, November 30, 2023), while the labor force from the CPS refers to the Sunday–Saturday week including the 12th of the month (Sunday, December 10, 2023, to Saturday, December 16, 2023) (BLS 2020a, 2024e). So the number of job openings refers to a day that is closer to next month’s CPS reference week than to this month’s CPS reference week.

Figure 3. Deviation from Full Employment in the United States, 1951:Q1–2019:Q4



in 1953:Q1, during the Korean War, and it bottomed at 0.16 in 2009:Q3, during the Great Recession. Twice, the labor market reached full employment just before entering a recession. This happened before the 1973–1975 recession (tightness peaked at 0.99 in 1973:Q3) and before the 2001 dot-com recession (tightness peaked at 1.01 in 2000:Q1).

We then compute the FERU using the formula $u^* = \sqrt{uv}$ (figure 4, panel A). The FERU is stable: It remains between 3.1 percent and 5.5 percent, with an average value of 4.3 percent. The Beveridge curve shifts in and out during the postwar period (Michaillat and Saez 2021a), but the shifts are not large enough to produce noteworthy changes in the FERU.

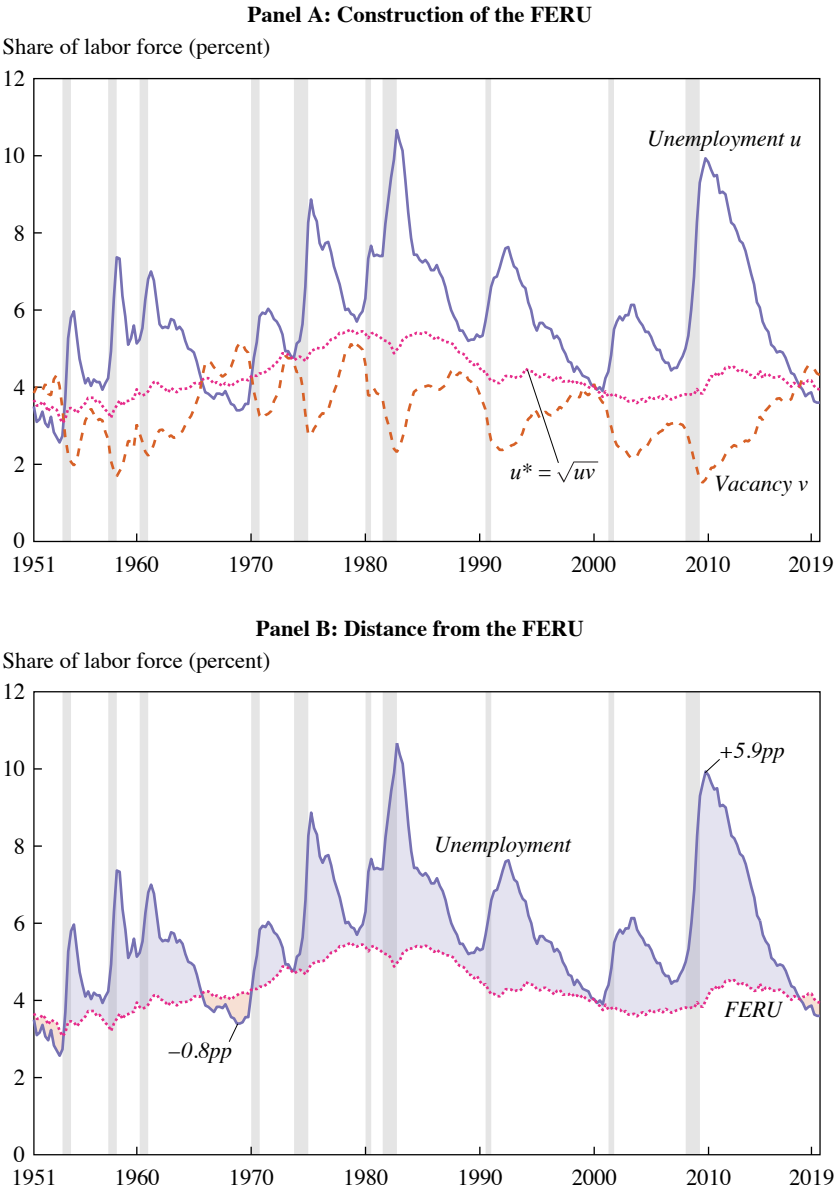
Of course, what is key to designing stabilization policy is not the FERU alone but the unemployment gap—the difference between unemployment rate and FERU, $u - u^*$. The unemployment gap indicates the distance from full employment at any given time. We compute the unemployment gap and find that it is generally positive and sharply countercyclical (figure 4, panel B). The unemployment gap averages +1.5 percentage points between 1951 and 2019. The gap peaked at +5.9 percentage points in 2009:Q4, during the Great Recession. At the end of the Volcker recession, in 1982:Q4, the gap reached the slightly lower value of +5.7 percentage points. The lowest value taken by the unemployment gap is −0.8 percentage points, in 1969:Q1, during the Vietnam War. During the Korean War, the unemployment gap was almost as low, reaching −0.7 percentage points in 1953:Q1. Hence, the economy is generally not at full employment. It is especially far from full employment in recessions.

III.B. The Great Depression and World War II

Next, we apply our full-employment criterion and FERU formula to the period 1930–1950, which covers both the Great Depression and World War II. Due to its simplicity, the FERU formula can easily be applied to such historical data.

The unemployment and vacancy rates for 1930–1950 are constructed by Petrosky-Nadeau and Zhang (2021). For 1930–1947, the unemployment rate is constructed by extrapolating David Weir’s (1992) annual unemployment series to a monthly series using monthly unemployment rates compiled by the National Bureau of Economic Research (NBER). For 1948–1950, the unemployment rate comes from BLS (2024k). The 1930–1950 vacancy rate is based on the help wanted index created by the Metropolitan Life Insurance Company (MetLife). This index aggregates help wanted advertisements from newspapers across major US cities, and it is regarded as a reliable proxy for job vacancies (Zagorsky 1998). The MetLife index

Figure 4. FERU in the United States, 1951:Q1–2019:Q4



Source: Unemployment rate u and vacancy rate v come from figure 1.
Note: The FERU is $u^* = \sqrt{uv}$. The vertical gray areas are NBER-dated recessions.

is scaled to align with Barnichon's (2010) vacancy rate at the end of 1950, effectively translating the index into a vacancy rate.²¹

Between 1930 and 1950, it remains true that unemployment and vacancy rates move in opposite directions (figure 5, panel A). In fact, using a logarithmic scale, it appears that unemployment and vacancy rates are inversely related (figure 5, panel B). These fluctuations indicate that just as in the postwar era, the Beveridge curve is close to a rectangular hyperbola in 1930–1950. To confirm this observation, we compute the elasticity of the 1930–1950 Beveridge curve by running an ordinary least squares (OLS) regression of log vacancy rate on log unemployment rate. We find an elasticity of -0.79 , which is not far from the elasticity of -1 for a rectangular hyperbola and is close to the elasticity of -0.84 for the 1951–1961 Beveridge curve (Michaillat and Saez 2021a, fig. 6). The 1930–1950 period saw vast fluctuations in unemployment and vacancy rates: The unemployment rate fluctuated between 1.0 percent and 25.3 percent; the vacancy rate fluctuated between 0.7 percent and 6.7 percent. Yet the hyperbolic shape of the Beveridge curve held well.

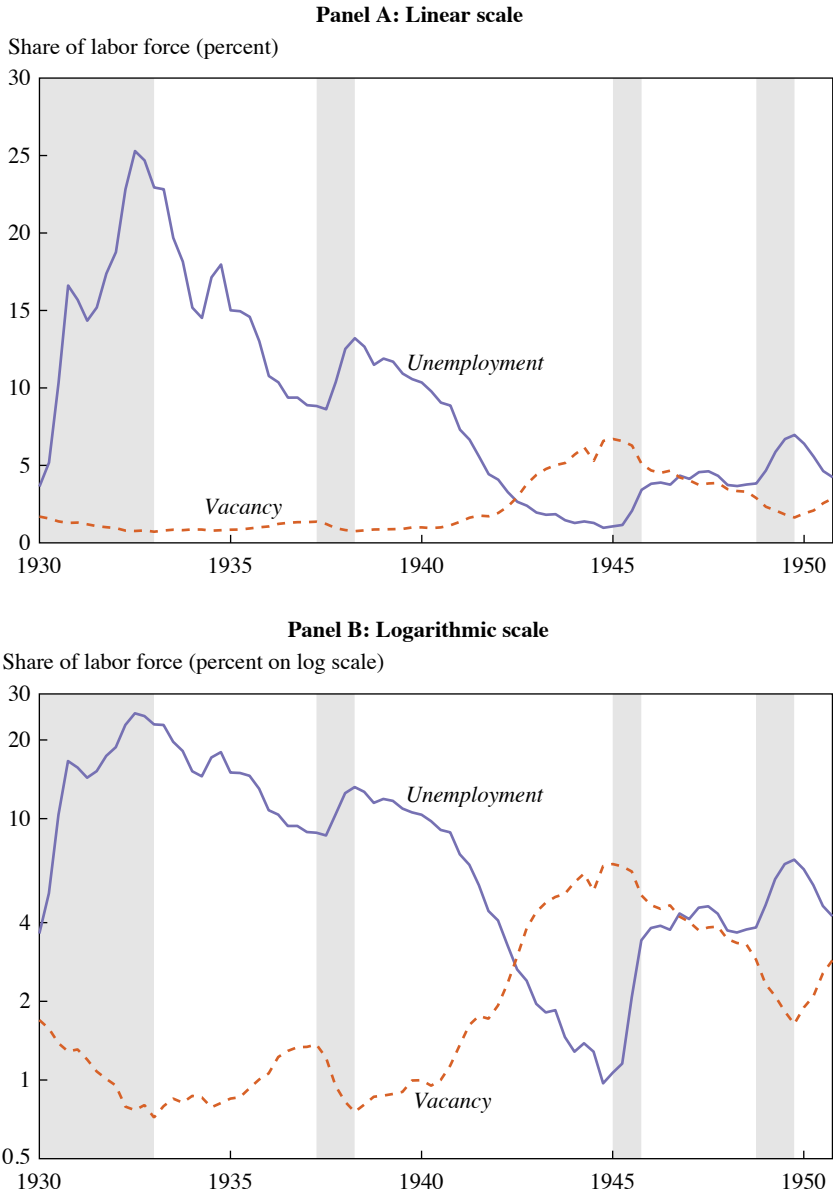
We compare the unemployment and vacancy rates to assess the state of the US labor market between 1930 and 1950 (figure 6, panel A). The unemployment rate averages 9.0 percent over the period, while the vacancy rate only averages 2.3 percent. So on average, the unemployment rate is markedly higher than the vacancy rate, which indicates that the US labor market is inefficiently slack. In fact, the US labor market is always inefficiently slack between 1930 and 1950 except during and right after World War II (1942:Q3–1946:Q3), when it was inefficiently tight.

The state of the labor market can also be visualized by plotting labor market tightness (figure 6, panel B). Tightness averages $0.85 < 1$ between 1930 and 1950, which confirms that the US labor market is inefficiently slack on average. Tightness is extremely volatile during the period. It plunged to 0.03 in 1932:Q3, during the Great Depression, and peaked at 6.8 in 1944:Q4, toward the end of World War II.

We then compute the FERU using $u^* = \sqrt{uv}$ (figure 7, panel A). Despite significant macroeconomic volatility during the period, the FERU is stable: It remains between 2.5 percent and 4.6 percent, with an average value of 3.5 percent.

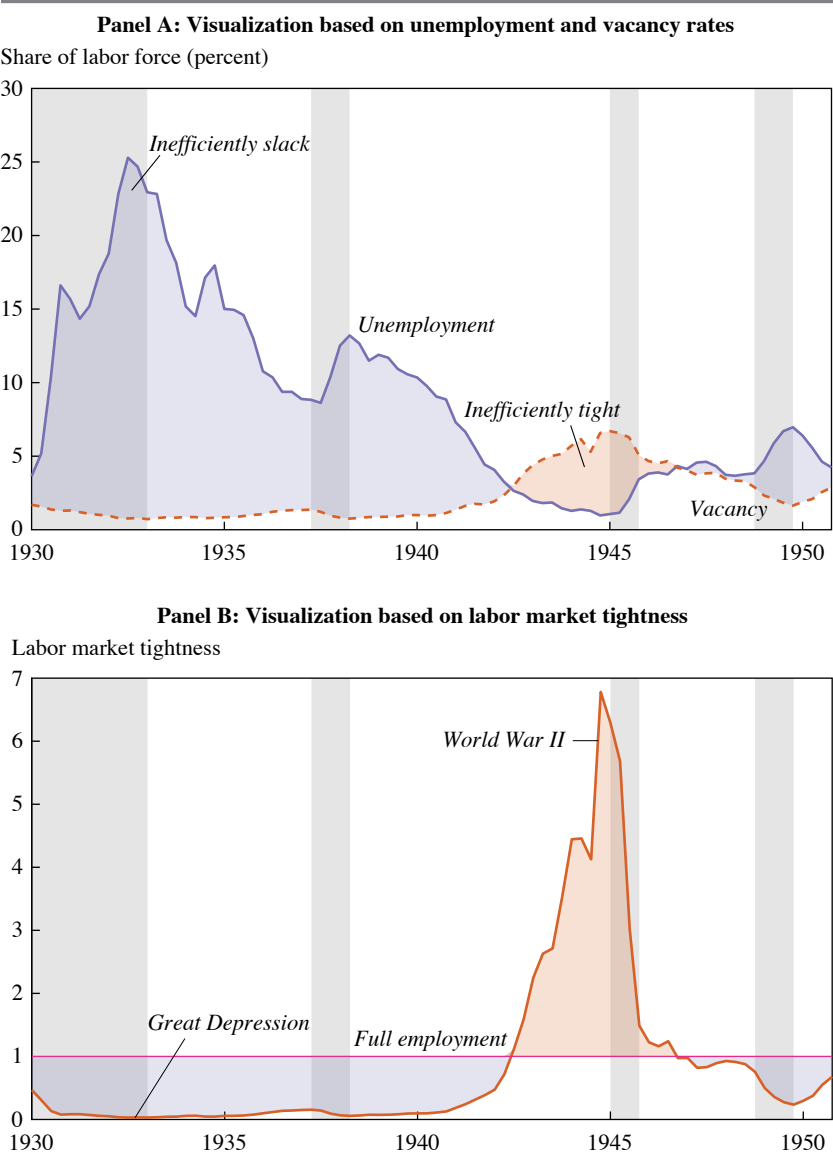
21. Petrosky-Nadeau and Zhang (2021) produce a vacancy series that starts in 1919 and an unemployment series that starts in 1890. Zagorsky (1998) argues, however, that the vacancy numbers are unreliable for 1919–1923 because some important newspaper data were missing during that time. Moreover, there is no monthly measure of unemployment between 1890 and 1929. Instead, the monthly unemployment fluctuations are inferred from the spread between the yields of bonds of different quality. Given these limitations, we begin our analysis in 1930.

Figure 5. Unemployment and Vacancy Rates in the United States, 1930:Q1–1950:Q4



Source: Unemployment and vacancy rates are quarterly averages of the monthly series constructed by Petrosky-Nadeau and Zhang (2021).
Note: The vertical gray areas are NBER-dated recessions.

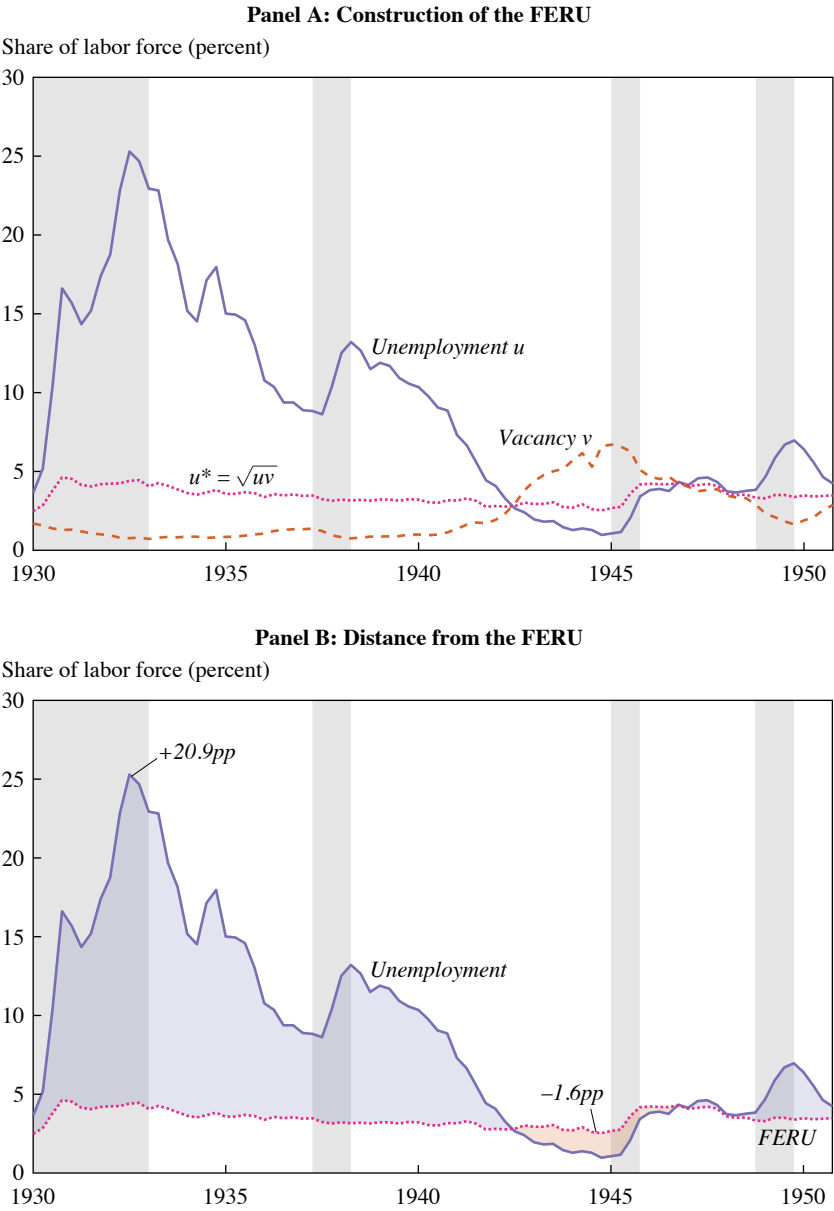
Figure 6. Deviation from Full Employment in the United States, 1930:Q1–1950:Q4



Source: Unemployment and vacancy rates come from figure 5.

Note: Labor market tightness is the ratio of job vacancies to job seekers. The vertical gray areas are NBER-dated recessions. The labor market is at full employment when the unemployment rate equals the vacancy rate, inefficiently slack when the unemployment rate exceeds the vacancy rate, and inefficiently tight when the unemployment rate is below the vacancy rate. Equivalently, the labor market is at full employment when tightness equals one, inefficiently slack when tightness is below one, and inefficiently tight when tightness exceeds one.

Figure 7. FERU in the United States, 1930:Q1–1950:Q4



Finally, we compute the unemployment gap $u - u^*$ (figure 7, panel B). The unemployment gap averages +5.5 percentage points between 1930 and 1950. The unemployment gap was, of course, positive and very large during the Great Depression: The labor market was much too slack then. The unemployment gap reached +20.9 percentage points in 1932:Q3. The economy recovered only slowly from the depression. The economy reached full employment in 1942:Q3, a few quarters after the United States had entered World War II. The unemployment gap kept falling during the war; it reached -1.6 percentage points in 1945:Q1. The unemployment gap turned positive again during the 1948–1949 recession.

III.C. Coronavirus Pandemic

Last, we apply our full-employment criterion and FERU formula to the coronavirus pandemic and its aftermath, from 2020:Q1 to 2024:Q2. Here the simplicity of the FERU formula allows us to apply it to real-time data and assess the current state of the US labor market.

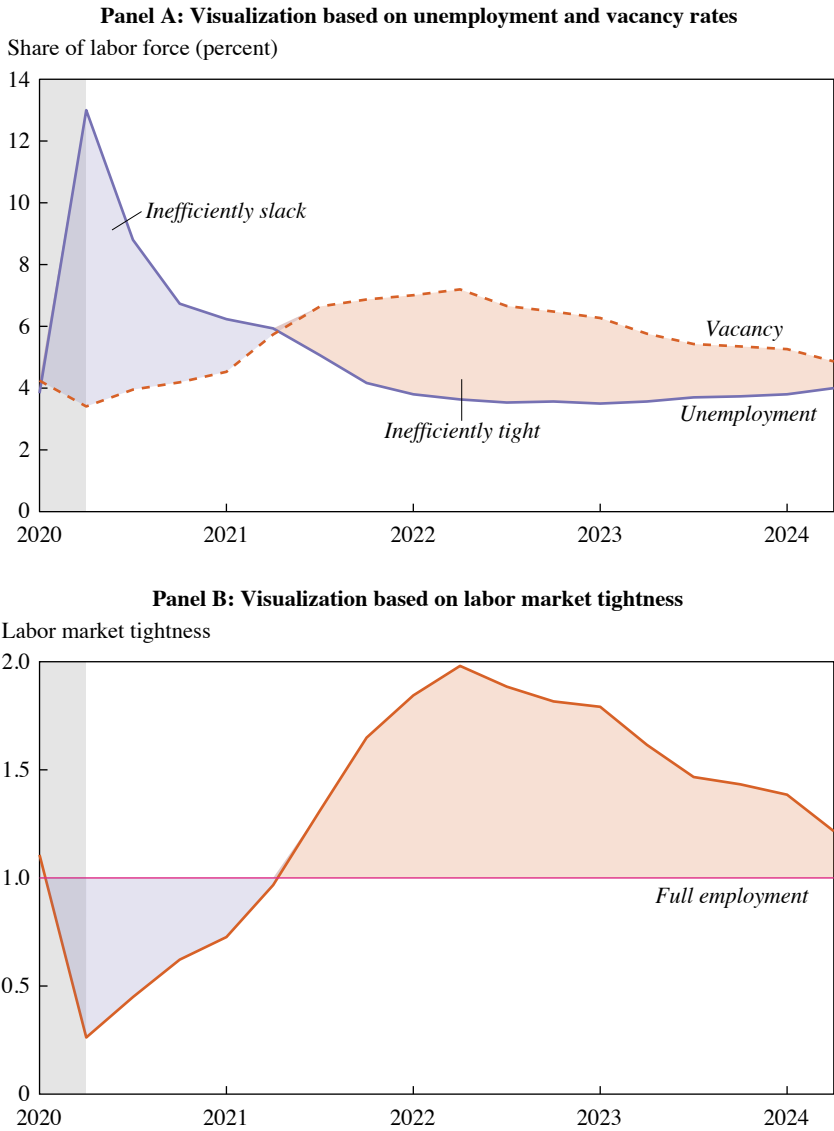
The unemployment rate is measured by BLS (2024k) from the CPS.²² The vacancy rate is calculated as the number of job openings measured by BLS (2024f) from the JOLTS, divided by the civilian labor force measured by BLS (2024a) from the CPS.²³ Both series are displayed on figure 8, panel A. Over 2020:Q1–2024:Q2, the unemployment rate averages 5.0 percent, and the vacancy rate averages 5.5 percent.

We compare the unemployment and vacancy rates to assess the state of the US labor market after the pandemic (figure 8, panel A). Between 2020:Q2 and 2021:Q2, the unemployment rate exceeds the vacancy rate, so the labor market is inefficiently slack. Then, between 2021:Q3 and 2024:Q2, the vacancy rate surpasses the unemployment rate, so the labor market is inefficiently tight.

22. At the start of the pandemic, many people in the CPS were misclassified as employed instead of unemployed (BLS 2020b). Their responses were recorded incorrectly, categorizing them as employed but absent from work when they should have been classified as unemployed on temporary layoff. This misclassification likely caused the reported unemployment rate to be lower than the true rate in March, April, and May 2020. In April and May, the true rate may have been up to 5 percentage points higher than reported (Barnichon and Yee 2020). The error was corrected from June 2020 onward, but the BLS lacked sufficient information to adjust the earlier rates. Here we follow their approach and use the official unemployment rate, though the 2020:Q2 rate may be underestimated.

23. The response rate to the JOLTS dropped markedly during the pandemic (BLS 2024d). It fell from 58 percent in December 2019 to 31 percent in September 2022. It has only recovered to 33 percent in April 2024. Hence, during this period, our measure of the vacancy rate might be surrounded by more uncertainty than usual.

Figure 8. Deviation from Full Employment in the United States, 2020:Q1–2024:Q2



Source: The unemployment rate is measured by BLS (2024k). The vacancy rate is the number of job openings divided by the civilian labor force, both measured by BLS (2024a, 2024f).

Note: Unemployment and vacancy rates are quarterly averages of monthly series. Labor market tightness is the ratio of job vacancies to job seekers. The vertical gray area is the NBER-dated pandemic recession. The labor market is at full employment when the unemployment rate equals the vacancy rate, inefficiently slack when the unemployment rate exceeds the vacancy rate, and inefficiently tight when the unemployment rate is below the vacancy rate. Equivalently, the labor market is at full employment when tightness equals one, inefficiently slack when tightness is below one, and inefficiently tight when tightness exceeds one.

Deviations from full employment can also be visualized by plotting labor market tightness (figure 8, panel B). Tightness averages 1.31 between 2020:Q1 and 2024:Q2. Tightness cratered to 0.26 in 2020:Q2, so the labor market was much too slack at the beginning of the pandemic. The labor market then recovered and passed the point of full employment (tightness of one) in the middle of 2021. Tightness then steadily rose to reach 1.98 in 2022:Q2. At that point, the labor market was much too tight. After peaking in 2022:Q2, tightness slowly fell down to 1.22 in 2024:Q2. So in 2024 tightness has returned to its pre-pandemic level (1.23 in 2019:Q2 and 1.21 in 2019:Q3). While the labor market remains too tight in 2024, it is nearing full employment.

Between 2020 and 2024, the FERU averages 5.1 percent (figure 9, panel A). The FERU was 4.0 percent in 2020:Q1, at the onset of the pandemic, but it sharply increased to 6.7 percent in the next quarter. It hovered around 6.0 percent during the rest of 2020–2021 and slowly decreased to 4.4 percent in 2024:Q2.

We also compute the unemployment gap $u - u^*$ (figure 9, panel B). While the unemployment gap averages zero over the period, the labor market experienced sharp departures from full employment. The unemployment gap was initially positive and large: The labor market was much too slack in the first year of the pandemic. The unemployment gap peaked at +6.3 percentage points in 2020:Q2. But the economy recovered quickly and reached full employment in the middle of 2021. The unemployment gap turned negative after that, reaching –1.5 percentage points in 2022:Q2. The gap then shrank to –0.4 percentage points in 2024:Q2. So during 2022–2024, the labor market was well beyond full employment.

The FERU increased by almost 3 percentage points at the onset of the pandemic (from 4.0 percent in 2020:Q1 to 6.7 percent in 2020:Q2). Such a sharp increase is unprecedented. It can be explained by the gigantic outward shift of the Beveridge curve that took place in the spring of 2020. Graphically, the FERU appears at the intersection of the Beveridge curve and the identity line (figure 10). In 2020:Q1, at the onset of the pandemic, the labor market was close to full employment, and the unemployment rate was at 3.8 percent. A year later, in 2021:Q2, the labor market had returned to the vicinity of full employment, but the unemployment rate was now 5.9 percent. This rise was caused by the outward shift of the Beveridge curve that occurred in the spring of 2020. Mathematically, the FERU is determined by the location of the Beveridge curve—equation (1)—so only a sharp outward shift of the curve can raise the FERU.

Figure 9. FERU in the United States, 2020:Q1–2024:Q2

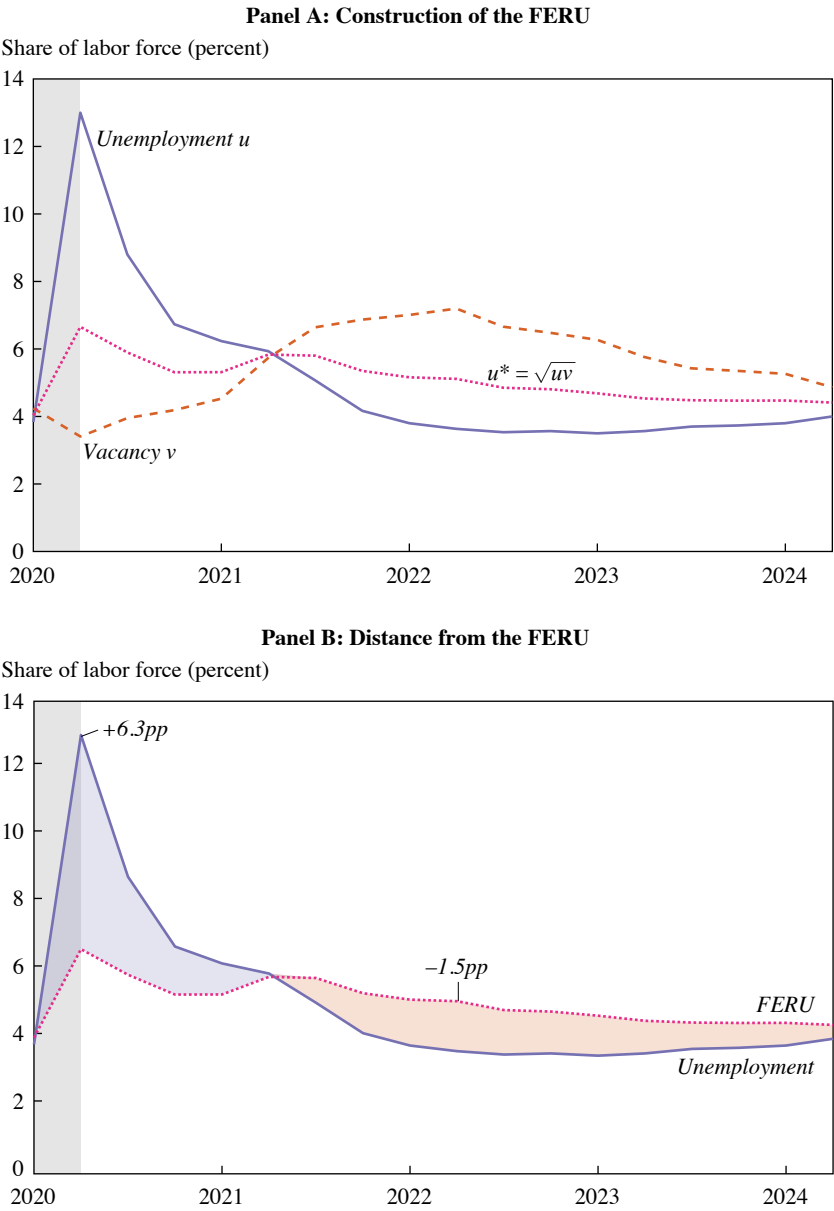
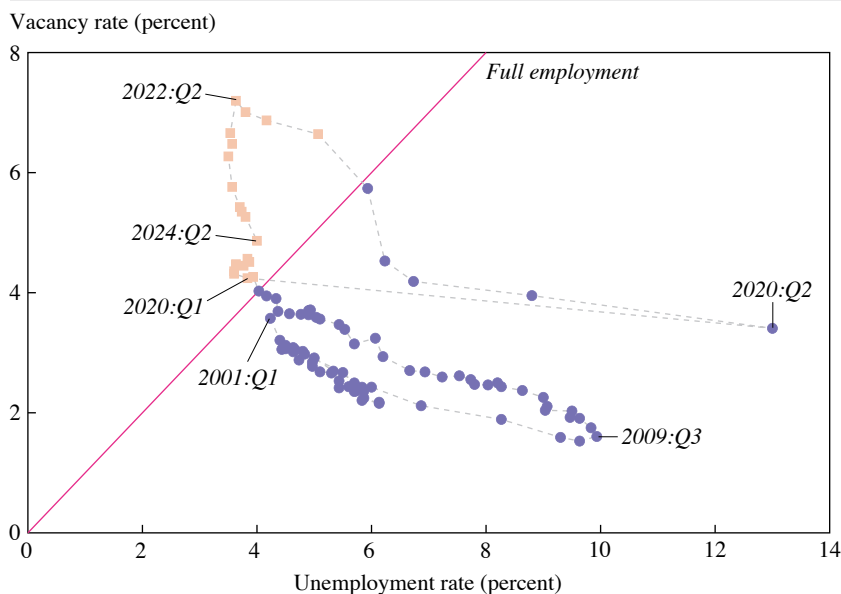


Figure 10. Beveridge Curve in the United States, 2001:Q1–2024:Q2

Source: The unemployment rate is measured by BLS (2024k). The vacancy rate is the number of job openings divided by the civilian labor force, both measured by BLS (2024a, 2024f).

Note: Each marker gives the unemployment and vacancy rates in a quarter between 2001 and 2024. The labor market is at full employment when the unemployment rate equals the vacancy rate, inefficiently slack when the unemployment rate exceeds the vacancy rate (circles), and inefficiently tight when the vacancy rate exceeds the unemployment rate (squares).

III.D. Complete 1930–2024 Period

To conclude, we combine the US unemployment and vacancy rates from 1930:Q1 to 2024:Q2. Given that the US labor market experienced extreme fluctuations during the entire period, especially in the first two decades, we plot the unemployment and vacancy rates, as well as labor market tightness and FERU, on logarithmic scales. Besides improving the readability of the figures, logarithmic scales have several advantages. First, the symmetry of the unemployment and vacancy movements on a logarithmic scale makes it clear that the Beveridge curve is a rectangular hyperbola. Second, the FERU is particularly easy to construct on a logarithmic scale: It is just the midpoint of the unemployment and vacancy rates.²⁴

24. Since $u^* = \sqrt{uv}$, then $\ln(u^*) = (\ln(u) + \ln(v))/2$.

A first finding is that, over almost a century, the unemployment rate is generally above the vacancy rate, and this gap is exacerbated in recessions (figure 11). This means that the labor market does not generally operate at full employment. Instead, it is generally inefficiently slack, especially during recessions. Over the period, the unemployment rate averages 6.4 percent, whereas the vacancy rate averages only half of that, 3.2 percent.

The labor market is not always inefficiently slack, however. There are several episodes when it becomes inefficiently tight. And these episodes do not appear at random. Before 2018, the labor market had only been inefficiently tight during major wars—World War II, the Korean War, the Vietnam War. Keynes (1936) doubted that an economy could reach full employment in peacetime. He was essentially right: Before 2018, the US economy had only reached full employment in wartime.

Since 2018, the labor market has been inefficiently tight just before the coronavirus pandemic (2018:Q3–2020:Q1) and in the aftermath of the pandemic (2021:Q3–2024:Q2). The state of the labor market around the pandemic is therefore a rarity: It is the only peacetime episode of an inefficiently tight labor market in the United States.

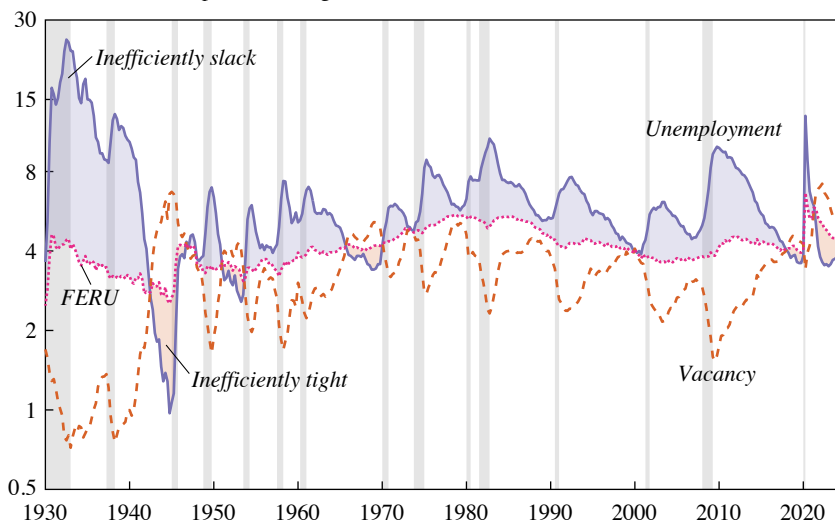
Over 1930–2024, the FERU averages 4.1 percent (figure 11). The FERU is stable over time, remaining between 2.5 percent and 6.7 percent over almost a century. It hovered around 4 percent between 1930 and 1970. It rose to about 5 percent in the 1970s and stayed there in the 1980s. It then remained around 4 percent again between 1990 and 2020. Finally, it temporarily rose above 6 percent during the pandemic, before falling back down below 5 percent after 2023.

Accordingly, over 1930–2024, the unemployment gap averages +2.3 percentage points. The unemployment gap reached its highest level on record, +20.9 percentage points, during the Great Depression. The unemployment gap then reached its lowest level on record, –1.6 percentage points, at the end of World War II. During and after the pandemic, the unemployment gap reached its highest and lowest levels since 1945. First, the unemployment gap peaked at +6.3 percentage points in the middle of the pandemic; then, the unemployment gap fell to –1.5 percentage points when the economy was recovering from the pandemic.

The state of the labor market can also be visualized by plotting labor market tightness (figure 12). Over 1930–2024, labor market tightness averages 0.73. Tightness is extremely volatile before the end of World War II. Tightness reached its most extreme values during that period: Tightness plunged to 0.03 during the Great Depression and climbed all the way to 6.8

Figure 11. FERU and Unemployment Gap in the United States, 1930:Q1–2024:Q2

Share of labor force (percent on log scale)



Source: The unemployment rate u and vacancy rate v are obtained by splicing the unemployment and vacancy rates from panel A in figures 1, 5, and 8.

Note: The FERU is $u^* = \sqrt{uv}$, so on a logarithmic scale it is the midpoint between the unemployment and vacancy rates. The vertical gray areas are NBER-dated recessions. The labor market is at full employment when the unemployment rate is equal to the FERU, or equivalently to the vacancy rate. The labor market is inefficiently slack when the unemployment rate is above the FERU, or equivalently above the vacancy rate. The labor market is inefficiently tight when the unemployment rate is below the FERU, or equivalently below the vacancy rate.

at the end of World War II. In the aftermath of the pandemic, the US labor market has become historically tight. In 2022:Q2, tightness reached 1.98, a value which it had last reached in 1945.

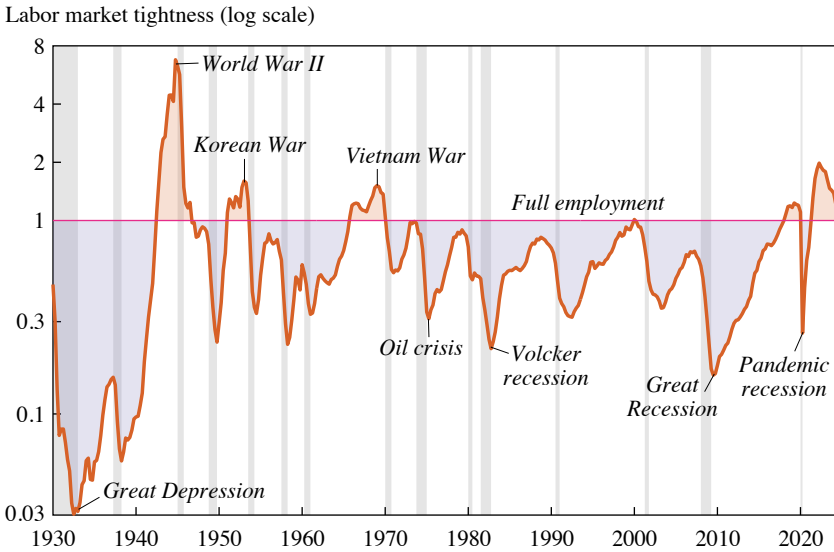
IV. Robustness of the US FERU

In this section, we demonstrate that the value of the FERU computed in section III is robust to an endogenous labor force, alternative calibrations of the model parameters, and alternative measures of unemployment.

IV.A. Endogenous Labor Force

We derive the formula $u^* = \sqrt{uv}$ by assuming that the labor force is exogenous. This assumption is motivated by evidence that labor force participation in the United States is acyclical. The formula continues to

Figure 12. Labor Market Tightness in the United States, 1930:Q1–2024:Q2



Source: Labor market tightness v/u is obtained by splicing labor market tightness from panel B in figures 3, 6, and 8.

Note: The vertical gray areas are NBER-dated recessions. The labor market is at full employment when tightness equals one, inefficiently slack when tightness is below one, and inefficiently tight when tightness exceeds one.

hold, however, if we endogenize labor force participation and allow it to be cyclical.

We normalize the size of the population to one, and we denote the size of the labor force by $h \in (0, 1)$. Each person $i \in [0, 1]$ has linear utility over consumption $c(i)$. If they do not participate in the labor force, they enjoy utility ηi^ϕ . The parameter $\eta > 0$ governs the utility from nonparticipation relative to consumption. The parameter $\phi \geq 0$ ensures that the utility from nonparticipation is increasing in i . People with high i enjoy nonparticipation very much. The utility from nonparticipation may come from home production or recreation.

A person's only decision is whether to participate in the labor force or not. If person i refuses to participate, they get utility ηi^ϕ . If they decide to participate, they receive utility from their expected labor income $(1 - u)w$, where $(1 - u)$ is the probability of finding a job, and w is the real wage that they receive if they find a job. We assume that unemployed workers do not

receive any income, but the analysis would be unchanged if they received unemployment benefits.²⁵

The participation decision is simple. Anyone with sufficiently high utility from nonparticipation relative to the expected labor income remains outside the labor force. Anyone with sufficiently low enough utility from nonparticipation participates. Formally, people opt to participate when $\eta i^\phi \leq (1 - u) w$, and they refuse to participate when $\eta i^\phi > (1 - u) w$. Accordingly, the size of the labor force is implicitly defined by

$$(5) \quad \eta h^\phi = (1 - u) w.$$

This equation says that the marginal labor force participant ($i = h$) is indifferent between participating and not, because their nonparticipation utility ηh^ϕ equals their expected labor income $(1 - u) w$.

Next, we compute the real wage w . We assume that firms have linear production functions and normalize labor productivity to one. In aggregate, firms employ $(1 - u) h$ workers, and among those, $(1 - u - v) h$ are producers and vh are recruiters. So firms produce $(1 - u - v) h$ goods and services. On the other hand, the aggregate real wage bill is $w (1 - u) h$. Under the usual assumption that firms make no profits because of free entry, the aggregate production and real wage bill must be equal, so

$$(6) \quad w = \frac{1 - u - v}{1 - u}.$$

Notice that $w < 1$: Producers are paid strictly less than their marginal product. This is because firms must make some profits on producers to cover recruiting costs.

Combining equations (5) and (6), we find that the labor force participation rate is an implicit function $h(u)$ of the unemployment rate:

$$(7) \quad h(u) = \left[\frac{1 - u - v(u)}{\eta} \right]^{1/\phi}.$$

25. Assume that the government provides unemployment benefits b to all job seekers and that the benefits are financed by a payroll tax t levied on all employees. Then the income of the $(1 - u) h$ employees becomes $(1 - t) w$ while the income of the uh job seekers becomes b . The expected income from participating therefore becomes $(1 - u) (1 - t) w + ub = (1 - u) w + [ub - (1 - u) tw]$. Since the unemployment insurance's budget must be balanced, the income provided to job seekers through unemployment benefits, buh , must equal the income taken away from employees through payroll taxes, $tw (1 - u) h$. The budget constraint requires $ub - (1 - u) tw = 0$, so the expected income from participation is unchanged at $(1 - u) w$.

The participation rate depends solely on the unemployment rate. This is because the unemployment rate determines the job-finding probability, $1 - u$, and the real wage, given by equation (6), and thus the expected labor income.

We now turn to the welfare function in this generalized framework. Social welfare is just the sum of individual utilities:

$$W = \int_0^1 c(i) di + \int_h^1 \eta i^\phi di.$$

The first term is social welfare from consumption, which is just aggregate consumption, $[1 - u - v(u)] h(u)$. The second term is social welfare from nonparticipation for everyone who decides to stay out of the labor force. Hence social welfare is a function of the unemployment rate:

$$W(u) = [1 - u - v(u)] h(u) + \int_{h(u)}^1 \eta i^\phi di.$$

The social planner chooses the unemployment rate u to maximize welfare $W(u)$. Using the Leibniz rule, we compute the first-order condition for the maximization problem:

$$0 = W'(u) = [-1 - v'(u)] h(u) + h'(u) [1 - (u + v) - \eta h(u)^\phi].$$

Critically, we learn from equation (7) that $\eta h(u)^\phi = 1 - (u + v)$, so the second term in the equation is zero. Therefore, the first-order condition reduces to $v'(u) = -1$, just as in the baseline case with fixed labor force. Since the Beveridge curve is a rectangular hyperbola, $v'(u) = -v/u$, we recover the result that welfare is maximized when $u = v$.

In sum, endogenizing labor force participation does not change the analysis at all. The FERU remains given by $u^* = \sqrt{uv}$. This result stems from an envelope theorem logic that is classic in public economics. Even if the social planner alters the labor force by changing the unemployment rate, welfare is unaffected because the workers who move in or out of the labor force are indifferent between participating or not.²⁶ For them, expected labor income equals utility from nonparticipation.

26. If workers strictly preferred participation to nonparticipation, they would move into the labor force. Conversely, if they strictly preferred nonparticipation, they would move out of the labor force. Participation only requires searching for a job, so nothing prevents a willing worker from participating. This differs from employment, which requires securing a job. Thus, nonparticipation is voluntary whereas unemployment is involuntary.

Our analysis also explains why the labor force participation rate seems acyclical. At full employment, $u + v(u)$ is minimized, so the derivative of $u + v(u)$ with respect to u is zero. Equation (7) then implies that the derivative of $h(u)$ with respect to u is zero around full employment. That is, unemployment has no first-order effect on the labor force participation rate around full employment.

Away from full employment, things are different, but no clear cyclical-ity emerges. When the unemployment rate is inefficiently low ($u < u^*$), then $u + v(u)$ is decreasing in u , so that $h(u)$ is increasing in u . When the unemployment rate is inefficiently high ($u > u^*$), the opposite occurs: $u + v(u)$ is increasing in u , so $h(u)$ is decreasing in u . Thus, the labor force participation rate is countercyclical when the labor market is inefficiently tight and procyclical when the labor market is inefficiently slack.

Given that the US labor market is generally inefficiently slack, we would expect the participation rate to be mildly procyclical. However, we would expect the fluctuations to be small because at the extensive margin, labor supply is quite inelastic, which means that the Frisch elasticity of labor supply $1/\phi$ is quite small (Chetty and others 2012). Through equation (7), this inelasticity implies that the participation rate does not respond much to the unemployment rate.

IV.B. Alternative Calibrations of the Model Parameters

We derive the formula $u^* = \sqrt{uv}$ by assuming that the elasticity of the Beveridge curve is -1 , each vacancy requires one recruiter, and the social product of unemployed labor is zero. This calibration is based on evidence for the United States.

Yet, the FERU formula can also be derived using a more general calibration of the parameters (Michaillat and Saez 2021a). We now assume that the Beveridge elasticity is $-\epsilon \neq -1$, the recruiting cost is $\kappa \neq 1$, and the social product of unemployed labor is $\zeta \neq 0$. Then social output becomes

$$(1 - u - \kappa v) + \zeta u = 1 - \left[(1 - \zeta)u + \kappa v \right].$$

Hence, the social planner minimizes $(1 - \zeta)u + \kappa v$, subject to the Beveridge curve $u^\epsilon v = A$.

Under this general calibration, the FERU becomes

$$(8) \quad u^* = \left(\frac{\kappa \epsilon}{1 - \zeta} u^\epsilon v \right)^{1/(1+\epsilon)},$$

as shown by Michailat and Saez (2021a). Of course, by setting the parameters to their baseline values, $\epsilon = 1$, $\kappa = 1$, and $\zeta = 0$, formula (8) reduces to $u^* = \sqrt{uv}$.

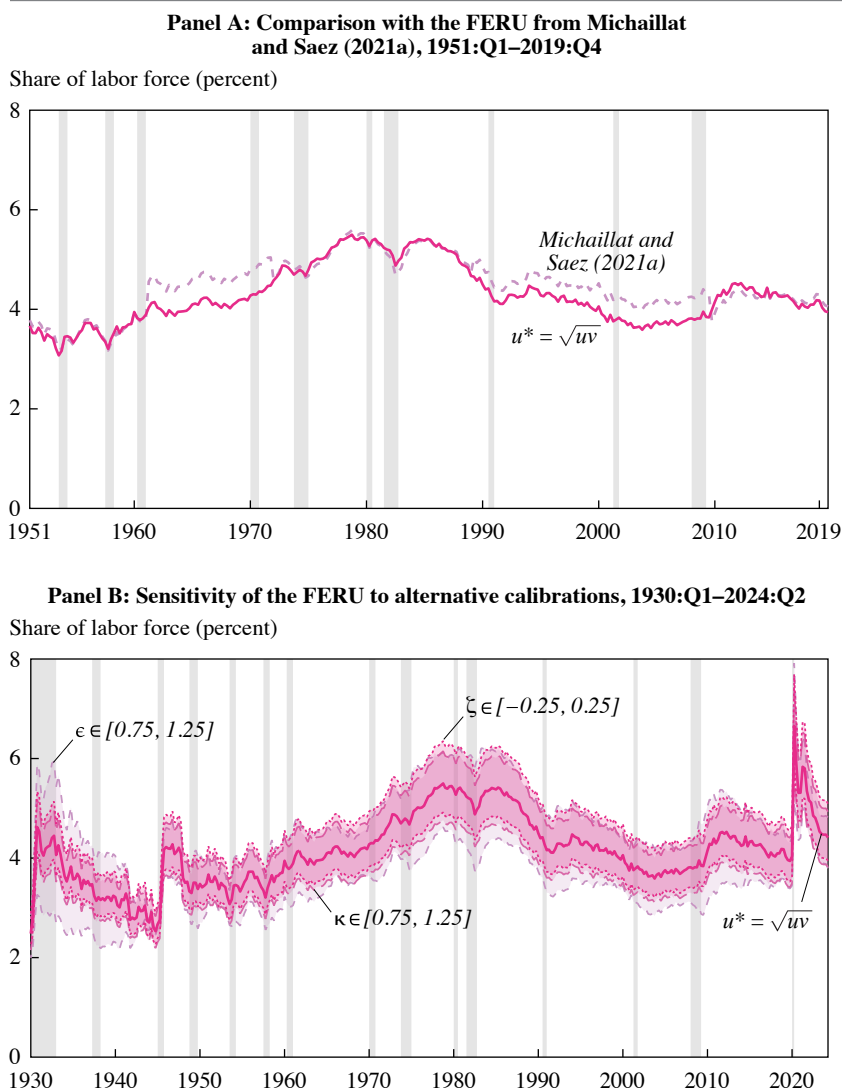
While the generalized formula is more flexible, it requires tracking three parameters (ϵ , ζ , κ), so it is harder to compute than $u^* = \sqrt{uv}$. The generalized formula is especially difficult to use in real time because it requires tracking the slope of the Beveridge curve, which is difficult when the curve shifts. By setting the parameters to reasonable but fixed values, we obtain $u^* = \sqrt{uv}$, which is simpler and thus better suited to measure full employment in real time.

In the United States, however, the two formulas yield almost identical FERUs (figure 13, panel A). Following Michailat and Saez (2021a, fig. 7B), we apply formula (8) with the Beveridge elasticity $-\epsilon$ estimated by Michailat and Saez (2021a, fig. 6), a recruiting cost $\kappa = 0.92$, and a social product of unemployed labor $\zeta = 0.26$. Between 1951 and 2019, that FERU only differs from $u^* = \sqrt{uv}$ by 0.2 percentage points on average; they never differ by more than 0.6 percentage points.

Naturally, there is some uncertainty about the true values of the parameters ϵ , κ , and ζ . However, the FERU produced by the generalized formula is fairly insensitive to these values. To demonstrate this, we compute the FERUs given by formula (8) when the parameters range from 25 percent below to 25 percent above their baseline values (figure 13, panel B). Specifically, the Beveridge elasticity $-\epsilon$ ranges from -1.25 to -0.75 , the recruiting cost κ ranges from 0.75 to 1.25 , and the social product of unemployed labor ζ ranges from -0.25 to 0.25 .²⁷ Between 1930:Q1 and 2024:Q2, the average width of the FERU band generated by $\epsilon \in [0.75, 1.25]$, $\kappa = 1$, and $\zeta = 0$ is 1.4 percentage points. The average width of the FERU band generated by $\kappa \in [0.75, 1.25]$, $\epsilon = 1$, and $\zeta = 0$ is 1.0 percentage points. And the average width of the FERU band generated by $\zeta \in [-0.25, 0.25]$, $\epsilon = 1$, and $\kappa = 1$ is 1.1 percentage points. Thus, all the FERUs remain close to the baseline value, located at the center of the band. For example, in 2024:Q2, the baseline FERU is 4.4 percent, while the FERUs are 3.8 percent for $\epsilon = 0.75$, 3.8 percent for $\kappa = 0.75$, 3.9 percent for $\zeta = -0.25$, 4.8 percent for $\epsilon = 1.25$, 4.9 percent for $\kappa = 1.25$, and 5.1 percent for $\zeta = 0.25$.

Overall, the value of the FERU in the United States is robust to a broad range of parameter calibrations. Nevertheless, to measure the FERU in

27. In formula (8), the relevant statistic is the social product of employed labor relative to unemployed labor, $1 - \zeta$, rather than ζ directly. For $1 - \zeta$ to range from 0.75 to 1.25 , ζ must range from -0.25 to 0.25 .

Figure 13. US FERU Under Alternative Calibrations of the Model Parameters

Source: For panel A, the solid line is the FERU from figure 11; the dashed line is the FERU from Michaillat and Saez (2021a, fig. 7B), which is obtained from formula (8) with the Beveridge elasticity $-\epsilon$ estimated by Michaillat and Saez (2021a, fig. 6), the recruiting cost $\kappa = 0.92$, and the social product of unemployed labor $\zeta = 0.26$. For panel B, the solid line is the FERU from figure 11. The area between the dot-dashed lines represents the FERUs obtained from formula (8) with $\epsilon \in [0.75, 1.25]$, $\kappa = 1$, and $\zeta = 0$. The area between the dashed lines represents the FERUs obtained from formula (8) with $\epsilon = 1$, $\kappa \in [0.75, 1.25]$, and $\zeta = 0$. The area between the dotted lines represents the FERUs obtained from formula (8) with $\epsilon = 1$, $\kappa = 1$, and $\zeta \in [-0.25, 0.25]$.

Note: The vertical gray areas are NBER-dated recessions.

other countries accurately, it might be necessary to use the generalized formula and adjust the calibration of the parameters. This is because the parameters might not take the same values abroad as in the US economy. The shape of the Beveridge curve is particularly likely to be different (Gäddnäs and Keränen 2023, table 1).

IV.C. Alternative Measures of Unemployment

When we apply the FERU formula to the US economy in section III, we stick to the official definition of unemployment. Here we recompute the FERU using broader definitions of unemployment—replacing the unemployment rate $U3$ by the broader unemployment rates $U4$ and $U5$ and adjusting the size of the labor force accordingly. To clarify that our baseline measures of u and v are based on the concept $U3$ of unemployment, we denote them by $u3$ and $v3$ here.

Unemployment comprises people able, willing, and seeking to work. The empirical challenge is to determine who is seeking a job. People search with different intensity and methods. Ideally anyone searching in any way would be counted as unemployed. However, the standard unemployment rate $U3$ only counts as unemployed people who have been actively searching in the past four weeks. There are workers who have been searching for a job in the past year but not in the past month who are not counted as unemployed, although in theory they belong there.

The unemployment concept $U4$ includes all the workers in the standard unemployment concept $U3$, plus workers who want a job, are available to start a job now, have been actively searching for a job in the past twelve months, but have not been searching in the past four weeks because they became discouraged about their job prospects. When asked why they did not look for work during the last four weeks, these workers respond, for instance, “There are no jobs available,” or “They have been unable to find work in the past” (BLS 2023). These additional workers are labeled discouraged workers. They are not counted in $U3$ because they did not actively search for work in the last four weeks.

The unemployment concept $U5$ includes all the workers in $U4$ plus workers who want a job, are available to start a job now, have been actively searching for a job in the past twelve months, but have not been searching in the past four weeks for reasons other than discouragement about their job prospects (BLS 2023). When asked why they did not look for work during the last four weeks, these workers respond, for instance, that they could not search because of family responsibilities, childcare problems, or ill health. These additional workers are not classified in $U3$ because they

did not actively search for work in the last four weeks; they are not classified in U4 because they were not discouraged about their job prospects. Together with the discouraged workers, these workers compose the marginally attached workers.

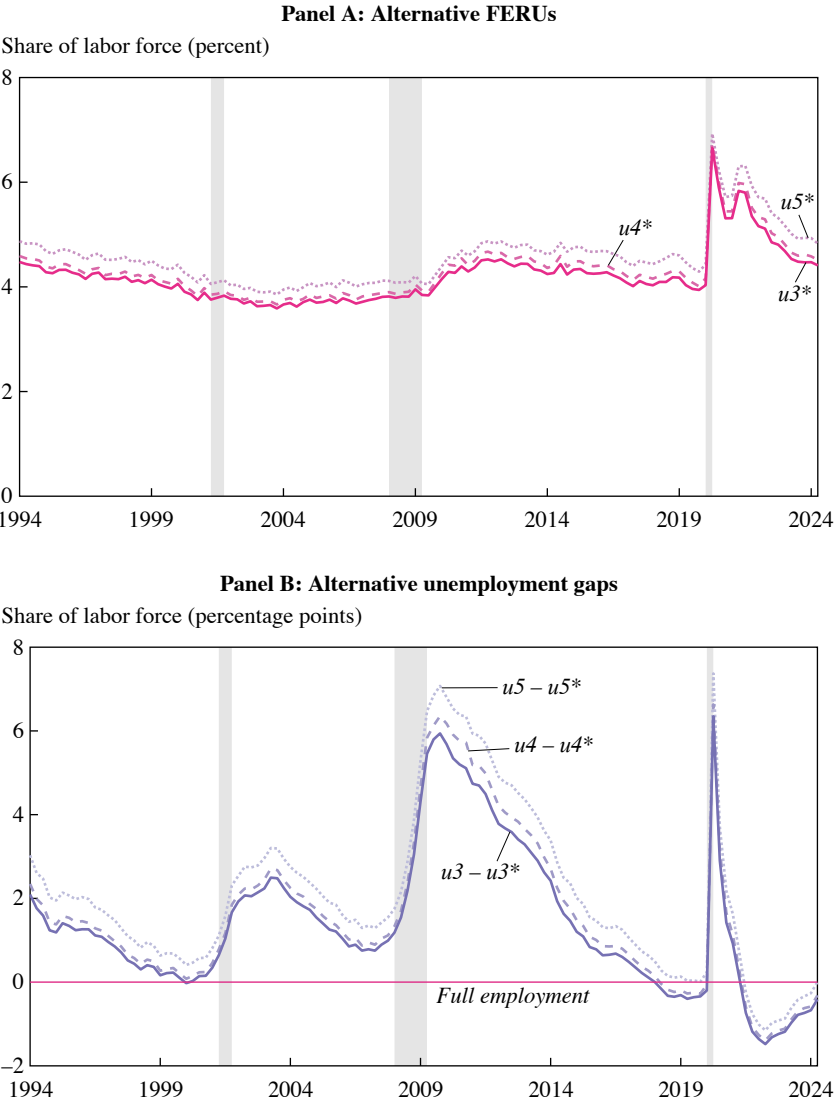
To be consistent with the concepts U4 and U5 of unemployment, we introduce two broader labor force sizes, which we use to compute unemployment and vacancy rates consistent with the U4 and U5 concepts. The unemployment rate $u4$, constructed by BLS (2024i), is the unemployment level U4 divided by a broader labor force made of the standard labor force plus discouraged workers, both measured by BLS (2024a, 2024g). We construct the vacancy rate $v4$ as the vacancy level from figure 11 divided by this broader labor force. In that way, the rates $u4$ and $v4$ have the same denominator. Similarly, the unemployment rate $u5$, constructed by BLS (2024j), is the unemployment level U5 divided by an even broader labor force, made of the standard labor force plus marginally attached workers, both measured by BLS (2024a, 2024h). We construct the vacancy rate $v5$ as the vacancy level from figure 11 divided by this even broader standard labor force. The unemployment concepts U4 and U5 were only introduced in 1994, so we can only measure $u4$, $v4$, $u5$, and $v5$ between 1994:Q1 and 2024:Q2.

Over 1994–2024, the standard unemployment rate, $u3$, averages 5.6 percent. By comparison, $u4$ averages 5.9 percent and $u5$ averages 6.7 percent. So the discouraged workers make up less than 0.5 percent of the labor force, and the marginally attached workers make up about 1 percent of the labor force. All the vacancy rates are extremely close, averaging 3.4 percent over the period.

Using these broader measures of unemployment, we construct broader measures of the FERU: $u4^* = \sqrt{u4 \times v4}$ and $u5^* = \sqrt{u5 \times v5}$ (figure 14, panel A). We compare these measures to the standard FERU: $u3^* = \sqrt{u3 \times v3}$. Between 1994:Q1 and 2024:Q2, $u3^*$ averages 4.2 percent, $u4^*$ averages 4.3 percent, and $u5^*$ averages 4.6 percent. So the three FERU measures are close to each other—closer, in fact, than the three measures of unemployment. We also see that all three measures follow the same pattern: The largest distance between $u3^*$ and $u4^*$ is only 0.2 percentage points and the largest distance between $u3^*$ and $u5^*$ is only 0.6 percentage points.

Finally, we construct broader measures of the unemployment gap: $u4 - u4^*$ and $u5 - u5^*$ (figure 14, panel B). The three unemployment gaps all move together. But the unemployment gaps constructed with the concepts U4 and U5 of unemployment are larger than the baseline unemployment gap, because the unemployment rates $u4$ and $u5$ are larger than $u3$. Over the

Figure 14. US FERU Under Alternative Measures of Unemployment, 1994:Q1–2024:Q2



Source: The FERUs are given by $u3^* = \sqrt{u3 \times v3}$, $u4^* = \sqrt{u4 \times v4}$, and $u5^* = \sqrt{u5 \times v5}$. The unemployment rates $u3$, $u4$, and $u5$ are measured by BLS (2024i, 2024j, 2024k). The vacancy level comes from figure 11. The vacancy rate $v3$ is the vacancy level divided by the number of labor force participants, measured by BLS (2024a). The vacancy rate $v4$ is the vacancy level divided by the number of labor force participants and discouraged workers, both measured by BLS (2024a, 2024g). The vacancy rate $v5$ is the vacancy level divided by the number of labor-force participants and marginally attached workers, both measured by BLS (2024a, 2024h).

Note: Unemployment and vacancy rates are quarterly averages of monthly series. The vertical gray areas are NBER-dated recessions.

1994–2024 period, the gap $u3 - u3^*$ averages +1.4 percentage points, the gap $u4 - u4^*$ averages +1.6 percentage points, and the gap $u5 - u5^*$ averages +2.1 percentage points. Under the largest unemployment gap, $u5 - u5^*$, the economy appears at full employment in 2019—and not inefficiently tight. The final readings of the unemployment gaps in 2024:Q2 are $u3 - u3^* = -0.4$ percentage points, $u4 - u4^* = -0.3$ percentage points, and $u5 - u5^* = 0.0$ percentage points. So in 2024, the labor market is inefficiently tight under U3 and U4, but it is back to full employment under U5.

V. Explaining Deviations from Full Employment in the United States

Despite the US government’s full-employment mandate, the US labor market has consistently fallen short of full employment in the past century. Here we attempt to explain why the US labor market has deviated from full employment in different periods.

V.A. *The Great Depression and Its Aftermath*

During the Great Depression and its aftermath, the US economy was exceedingly slack. From the beginning of 1930, when our data begin, to the end of 1941, when the United States entered World War II, the unemployment gap averages +9.6 percentage points (figure 7, panel B). So the US economy was extremely far from full employment.

Three factors may explain this large amount of slack. The first is that the US government and the Federal Reserve did not have a full-employment mandate at the time. The mandate was introduced with the Employment Act of 1946, as a result of the Great Depression. A second factor is that the Federal Reserve was committed to the gold standard. The gold standard generated a deep deflation in the early 1930s, with dramatic consequences (Eichengreen and Temin 2000). A third factor is that the Fed failed to curb recurrent banking panics in the 1930s (Friedman and Schwartz 1963). Overall, as former Fed Chair Ben Bernanke (2022, xvii) writes, “Blaming the Depression entirely on the Fed is an exaggeration, but the relatively new and unseasoned central bank did perform poorly.”

V.B. *World War II, the Korean War, the Vietnam War*

The US labor market was pulled out of its Great Depression slackness by World War II (figure 12). In fact, the labor market became inefficiently tight during the war, with tightness averaging 4.0 over the 1943–1945 period. The labor market was once again inefficiently tight during the Korean War,

with tightness averaging 1.3 over the 1951–1953 period, and during the Vietnam War, with tightness averaging 1.3 over the 1966–1969 period.

Why was tightness so high during the wars? Part of the reason is that government spending was substantial during these three major wars (Ramey and Shapiro 1998). Such expenditure boosts aggregate demand, which increases tightness (Michaillat and Saez 2019, fig. 2). Another part of the reason is that millions of potential labor force participants were sent abroad on military duty (US Department of Veteran Affairs 2023). Such drastic reduction in labor force reduces labor supply, which raises tightness and reduces the unemployment rate among the workers who stayed in the United States (Michaillat and Saez 2022, fig. 4).

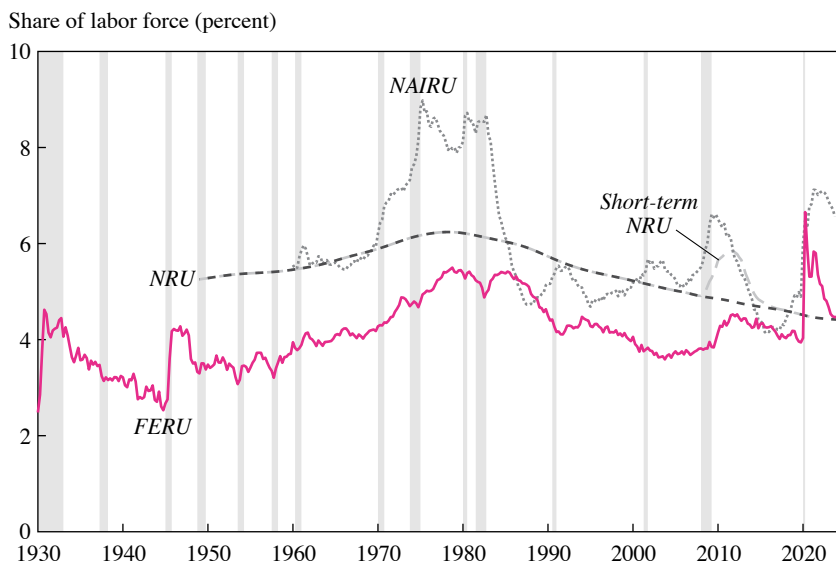
So why didn't the Fed tighten monetary policy to reduce tightness in wartime? Indeed, a high real interest rate curbs aggregate demand, which reduces tightness and raises unemployment. An appropriate increase in interest rates could have brought tightness back to its full-employment level of one (Michaillat and Saez 2022, figs. 5 and 7). In the case of World War II, there is a simple answer. As Bernanke (2022, xviii) explains, during and shortly after World War II, "At the Treasury's request, the Fed held interest rates at low levels to reduce the government's cost of financing the war."

The same happened at the beginning of the Korean War, when "facing new hostilities in Korea, President Truman pressed the Fed to keep rates low" (Bernanke 2022, xviii). The Fed did rebel and was allowed to phase out the low interest rate peg that had been in place. But the phasing out came too late to cool down the labor market.

The situation during the Vietnam War was different. The Fed raised interest rates by half a percentage point at the end of 1965, at the exact time when the economy had reached full employment (Bernanke 2022). However, President Johnson was furious that the Fed tightened monetary policy. He needed low rates to help finance the war. Despite the pressure exerted by Johnson, the Fed continued increasing rates in 1966, which rapidly cooled the labor market. Worried about a possible recession, the Fed reversed its previous tightening. Under pressure from the White House, and facing a chaotic political situation, the Fed continued swinging between tightening and loosening until 1970. The lack of decisive tightening explains why the labor market remained so hot from 1966 to 1969.

V.C. Postwar Period

In the postwar period, the US labor market was generally inefficiently slack (figure 4, panel B). A manifestation of such pervasive slack is that the unemployment gap averaged 1.4 percentage points between 1946 and 2019.

Figure 15. FERU, NRU, and NAIRU in the United States, 1930:Q1–2024:Q2

Source: The FERU comes from figure 11. The NRU is constructed by CBO (2024) for 1949:Q1–2024:Q2. The short-term NRU is constructed by CBO (2021) for 1949:Q1–2020:Q4. The NAIRU is constructed by Crump and others (2024, fig. 2) for 1960:Q1–2023:Q4.

Note: The vertical gray areas are NBER-dated recessions.

Another manifestation is that the labor market remained below full employment for almost half a century: It was inefficiently slack from 1970 to 2017, only touching full employment in two quarters, in 1973:Q3 and 2000:Q1.

A first possible reason to explain this slackness is that the Fed and other policymakers often use the NRU computed by the CBO (2024) to measure full employment. Over 1949–2019, the NRU averages 5.5 percent (figure 15). This is 1.2 percentage points above the average FERU between 1949 and 2019. So policymakers might have targeted an unemployment rate that was just too high. The average distance between FERU and NRU by itself explains almost the entire average postwar unemployment gap.

Another measure of full employment that policymakers sometimes use is the NAIRU—although there is no standardized time series for it. Just like the NRU, the NAIRU appears to be significantly higher than the FERU. For instance, the NAIRU computed by Crump and others (2024, fig. 2) using state-of-the-art techniques averages 5.9 percent over 1960–2019 (figure 15). This is 1.5 percentage points more than the average FERU over the same period. Once again, by using the NAIRU, policymakers would have targeted an unemployment rate that was too high.

A second reason that might explain the slackness of the US labor market in the postwar period, especially after 1970, is that the Fed prioritized inflation at the expense of unemployment. Thornton (2011) reviews policy directives by the Federal Open Market Committee (FOMC) and finds that it made no reference to unemployment or full employment between 1979 and 2008—despite the dual mandate introduced in 1977. Instead, Thornton (2011, 1) finds that the FOMC preferred “to state its objectives in terms of price stability and economic growth.” This changed at the end of 2008, when the FOMC started mentioning its dual objective of “maximum employment and price stability” in policy directives and statements (Thornton 2011, 2). Kaya and others (2019) also detect this focus on inflation in FOMC transcripts. They find that from 1960 to 2010 FOMC discussions increasingly emphasized inflation relative to unemployment and that this shift occurred during the Volcker era and continued even as inflation declined. They conclude that “the emphasis on inflation has become entrenched and disconnected from actual inflation” (Kaya and others 2019, 641).

The prioritization of inflation might be due to a change in the Fed’s preferences or in macroeconomic theory. It might also come from Congress. Hess and Shelton (2016) examine legislative activity to determine when Congress pressures the Fed and whether this pressure affects monetary policy. They find that by “the late 1980s Congress shifted from threatening [the Fed] when unemployment was high to threatening when inflation was high” (Hess and Shelton 2016, 603). This finding is consistent with Margaret Weir’s (1987, 377) view that “by the mid-1980s full employment had been all but erased as a major political issue in the United States.” In fact, Weir (1987, 395) argues that although the Kennedy CEA identified an unemployment rate of 4 percent as full employment, in the following decades “more conservative economists [offered] ever-increasing rates of unemployment as the ‘true’ definition of full employment.”

V.D. The Great Recession

The Great Recession saw the highest unemployment gap of the 1946–2019 period, at +5.9 percentage points, and it presented new challenges to the Fed (figure 4, panel B). Although the unemployment gap skyrocketed in 2008–2009, the Fed was unable to respond because it ran against the zero lower bound on nominal interest rates from the end of 2008 until 2015 (Federal Reserve Board 2024). The Fed could not stimulate aggregate demand through lower interest rates because it was constrained by the zero lower bound, so it could not boost tightness and lower unemployment

(Michaillat and Saez 2022, fig. 8). Hence, unemployment remained inefficiently high until 2018.

The Fed did resort to unconventional monetary policy, including forward guidance and quantitative easing, to reduce long-term interest rates in the aftermath of the Great Recession (Kuttner 2018). But the effectiveness of such policies is debatable (Greenlaw and others 2018; Michaillat and Saez 2021b). Moreover, the Fed may not have used these policies aggressively enough because once again they targeted an unemployment rate that was too high. The Fed commonly uses the CBO's NRU to indicate full employment. During the Great Recession the CBO (2021) adjusted the NRU upward by 1 percentage point because they believed that structural factors temporarily kept the unemployment rate high. As a result, in 2011:Q4, the short-term NRU reached 5.8 percent (figure 15). We do find that the outward shift of the Beveridge curve after the Great Recession raised the FERU by 0.5 percentage points, but the FERU stood at only 4.5 percent in 2011:Q4—1.3 percentage points below the short-term NRU.

V.E. Coronavirus Pandemic

The coronavirus pandemic sharply slowed down economic activity. In 2020, the US economy reached the largest unemployment gap since the Great Depression, at +6.3 percentage points (figure 9, panel B). As during the Great Recession, the Fed could not respond more aggressively to the slackness of the economy because of the zero lower bound (Federal Reserve Board 2024).

Thanks to aggressive expansionary fiscal policy, however, the US economy recovered rapidly from the pandemic (Romer 2021). The US economy reached full employment in 2021:Q2 and continued tightening after that (figure 8, panel B). In 2022:Q2, labor market tightness reached 1.98, a level it had not seen since the end of World War II. It is only then, in spring of 2022, that the Fed started tightening monetary policy (Federal Reserve Board 2024). It is unclear why the Fed did not start tightening monetary policy earlier. After the labor market became too tight (2021:Q2), an entire year passed before the Fed increased rates (2022:Q2). This delay is all the more surprising since inflation was also above its target of 2 percent at the time. Core inflation was 3.7 percent in 2021:Q2 and rose to 6.3 percent in 2022:Q1 (BLS 2024b). Combined with the two years required by monetary policy to be fully effective (Coibion 2012), this delay explains well why the labor market remained inefficiently tight until 2024:Q2.

VI. Conclusion

To conclude, we summarize our findings, review the policy prescriptions that emerge from the analysis, and explore how unemployment and vacancy data can also be leveraged to detect recessions.

VI.A. Summary

In the United States, the federal government and its central bank are mandated to stabilize the economy at full employment. However, there is no agreed-upon measure of the FERU, which makes it difficult for them to design policy to achieve full employment and for observers to assess their performance (Duboff 1977).

In this paper, we argue that the US FERU is given by $u^* = \sqrt{uv}$, where u is the unemployment rate and v is the vacancy rate. Between 1930 and 2024, the FERU is stable, hovering around 4 percent. The FERU has generally been below the unemployment rate, so the US economy has generally fallen short of full employment.

VI.B. How to Achieve Full Employment?

Since \sqrt{uv} can be measured in real time, the US government and Fed could use u^* as their full-employment target. But which policies can bring the economy to full employment?

The most natural choice is monetary policy. Empirically, we know that reducing the federal funds rate lowers unemployment (Bernanke and Blinder 1992; Christiano, Eichenbaum, and Evans 1999; Coibion 2012; Ramey 2016). A midrange estimate is that lowering the nominal interest rate by 1 percentage point decreases the unemployment rate by 0.5 percentage points (Michaillat and Saez 2022). Theoretically, the mechanism is simple. Reducing the federal funds rate lowers the real interest rate, which makes consumption more appealing than saving and boosts aggregate demand. A higher aggregate demand raises market tightness and lowers unemployment (Michaillat and Saez 2022, fig. 5).

If the zero lower bound on nominal interest rates becomes binding, conventional monetary policy cannot restore full employment. But other policies, such as government spending, can bring the economy closer to full employment. Empirically, it is clear that government spending reduces unemployment (Monacelli, Perotti, and Trigari 2010; Ramey 2013). A midrange estimate is that raising government spending by 1 percent of GDP decreases the unemployment rate by 0.5 percentage points (Michaillat and Saez 2019). Here again the theoretical mechanism is simple: Increasing

government spending boosts aggregate demand, which raises market tightness and lowers unemployment (Michaillat and Saez 2019, fig. 2).

VI.C. When Is It Optimal to Maintain the Economy at Full Employment?

Targeting the FERU would not only satisfy the Fed's legal mandate, it would also be the optimal monetary policy in a range of models built around the Beveridge curve. For instance, in models in which inflation is fixed, the optimal monetary policy is to adjust interest rates in order to maintain unemployment at the FERU (Michaillat and Saez 2022, fig. 7). In such models, monetary policy does not affect inflation, so it is optimal to keep unemployment at the FERU.

Of course, fixed inflation is a strong assumption. But maintaining unemployment at the FERU is also optimal in models with endogenous inflation, as long as the divine coincidence holds (Michaillat and Saez 2024a). In such models, lower unemployment leads to higher inflation, but when the unemployment rate is efficient, inflation is on target. Therefore, there is no trade-off between inflation and unemployment: Maintaining unemployment at the FERU also maintains inflation on target.

How big should adjustments in interest rates be to keep the economy at full employment? Starting from a federal funds rate $r > 0$ and an inefficient unemployment rate $u \neq u^*$, the Fed should set the federal funds rate to r^* so that

$$(9) \quad r - r^* = \frac{u - u^*}{du/dr},$$

as shown by Michaillat and Saez (2022). The statistic $r - r^*$ indicates the change in interest rate required to reach full employment. The statistic $u - u^*$ is the prevailing unemployment gap. And the statistic $du/dr > 0$ is the monetary multiplier: the decrease in unemployment rate achieved by lowering the nominal interest rate by 1 percentage point. With a monetary multiplier of 0.5, formula (9) shows that the Fed should cut its interest rate by 2 percentage points for each positive percentage point of unemployment gap and raise its interest rate by 2 percentage points for each negative percentage point of unemployment gap.

If the zero lower bound becomes binding, conventional monetary policy is unable to achieve full employment (Michaillat and Saez 2022, fig. 8). Government spending can bring the economy to full employment, but that might not be optimal.

VI.D. *How to Use the FERU in More Complicated Situations*

The FERU remains a useful statistic to design optimal monetary and fiscal policy even in models in which targeting the FERU is suboptimal. For instance, when government spending is not a perfect substitute for private spending, it is not optimal to use government spending to eliminate the unemployment gap (Michaillat and Saez 2019). Optimal government spending deviates from the Samuelson (1954) rule to reduce—but not eliminate—the unemployment gap. Yet, the FERU and unemployment gap remain key determinants of optimal government spending. The optimal level of government spending is determined by the unemployment gap, together with the elasticity of substitution between public and private consumption and the fiscal multiplier (Michaillat and Saez 2019).

The same logic applies to optimal monetary policy in models without divine coincidence. When the divine coincidence fails, monetary policy faces a trade-off between closing the unemployment gap and closing the inflation gap, so targeting the FERU is no longer optimal. Nevertheless, the optimal interest rate will depend on the FERU: It will be determined by weighing the unemployment gap against the inflation gap.

To see this, consider an extension of our framework in which social welfare depends not only on unemployment but also on inflation. Let's denote the efficient inflation rate by π^* . The welfare loss around the efficient allocation (u^*, π^*) admits the following quadratic approximation:

$$(10) \quad \mathcal{L}(u, \pi) = (\pi - \pi^*)^2 + \alpha(u - u^*)^2,$$

where $\alpha > 0$ measures the importance of unemployment relative to inflation in the social welfare function.

Additionally, the social planner faces a Phillips curve that relates inflation to unemployment. Around the efficient allocation, the Phillips curve admits the following linear approximation:

$$(11) \quad \pi - \pi^* = -\beta(u - u^*) + \gamma,$$

where $\beta > 0$ gives the slope of the downward-sloping Phillips curve, and $\gamma \neq 0$ is introduced to break the divine coincidence. Indeed, if $\gamma = 0$, then $\pi = \pi^*$ when $u = u^*$: Inflation is on target at full employment, so the divine coincidence holds. If $\gamma > 0$, then $\pi > \pi^*$ when $u = u^*$: Inflation is too high at full employment, so unemployment must be above the FERU to bring

inflation to target, just like in the 1970s. Conversely, if $\gamma < 0$, then $\pi < \pi^*$ when $u = u^*$: Inflation is too low at full employment.

The Beveridgean Phillips curve developed by Michaillat and Saez (2024a) links the unemployment and inflation gaps, but it also guarantees that the divine coincidence holds, so $\gamma = 0$ in equation (11). More work is required to understand why the divine coincidence might fail and to identify the economic forces driving $\gamma \neq 0$.

The social planner minimizes the welfare loss (10) subject to the Phillips curve (11). Inflation can be substituted out of the welfare loss using the Phillips curve. Then, the planner's problem is to find $u \in [0, 1]$ to minimize

$$\left[-\beta(u - u^*) + \gamma \right]^2 + \alpha[u - u^*]^2.$$

This objective function is strictly convex in u , so the first-order condition is sufficient to find its minimum. We take the function's derivative with respect to u and set it to zero:

$$0 = -2\beta \left[-\beta(u - u^*) + \gamma \right] + 2\alpha[u - u^*].$$

Rearranging terms, we express the optimal unemployment gap as a function of the parameters:

$$(12) \quad u - u^* = \frac{\beta\gamma}{\alpha + \beta^2}.$$

Plugging equation (12) into the Phillips curve (11), we express the optimal inflation gap as a function of the same parameters:

$$(13) \quad \pi - \pi^* = \frac{\alpha\gamma}{\alpha + \beta^2}.$$

From these two expressions, we also obtain a simple expression for the ratio between the optimal unemployment and inflation gaps:

$$(14) \quad \frac{u - u^*}{\pi - \pi^*} = \frac{\beta}{\alpha}.$$

The trade-off between inflation and unemployment in the absence of divine coincidence clearly appears in these equations. Consider the case

where inflation is too high at the FERU ($\gamma > 0$). Then equations (12) and (13) say that it is optimal to keep unemployment above the FERU ($u > u^*$) and inflation above its target ($\pi > \pi^*$). Furthermore, equation (14) shows that the optimal gaps are determined by the welfare cost of unemployment relative to inflation (α) and the response of inflation to unemployment (β). If unemployment is particularly costly (high α) or if the Phillips curve is flat (low β), the optimal unemployment gap is small relative to the optimal inflation gap. In such cases, the optimal unemployment rate remains close to the FERU, either due to the high marginal cost of unemployment (high α) or the limited marginal benefit of unemployment (low β).

The equations provide two additional insights. First, if the divine coincidence holds ($\gamma = 0$), then it is optimal to keep unemployment at the FERU, which guarantees that inflation hits its target: $u = u^*$ in equation (12) and $\pi = \pi^*$ in equation (13). Second, it is never optimal for the unemployment and inflation gaps to have opposite signs: $(u - u^*)/(\pi - \pi^*) > 0$ in equation (14). This is because welfare can be improved if the gaps take opposite signs. For example, if the unemployment gap is negative and the inflation gap is positive (as in 2021–2024), then raising unemployment reduces the unemployment gap and, by cooling inflation, reduces the inflation gap—thereby improving welfare.

VI.E. Using Unemployment and Vacancy Data to Detect Recessions

Finally, the combination of vacancy and unemployment data has other interesting business cycle applications. Vacancy data are the black sheep of business cycle data: They are not well known, not well understood, and not widely trusted. Yet, when combined with unemployment data, they are extremely powerful for understanding business cycles. This paper provides an example of the normative power of the vacancy-unemployment combination. Michailat and Saez (2024b) show that the vacancy-unemployment combination has predictive power too.

Michailat and Saez (2024b) develop a new Sahm (2019)-type recession rule that combines vacancy and unemployment data. The new rule has greater foresight than the Sahm rule—which only uses unemployment data. It detects recession starts with a lag of 0.8 months on average, while the Sahm rule detects them with a lag of 2.1 months (Michailat and Saez 2024b, table 1). The new rule also has a better historical track record. It perfectly identifies the fifteen recessions that have occurred since 1929, without any false positives, while the Sahm rule breaks down before 1960.

In the present context, the recession rule proposed by Michailat and Saez (2024b) says that the US economy may have entered a recession as

early as March 2024. As of August 2024, the probability that the US economy is in recession is 48 percent.

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Comments and Discussion

COMMENT BY

HIE JOO AHN¹ Michaillat and Saez develop a theoretical framework to estimate the unemployment rate at full employment, referred to as the FERU. The FERU represents the unemployment rate that achieves a socially efficient allocation of labor. Specifically, it is the unemployment rate at which social output is maximized by minimizing the unproductive uses of labor, such as job searching and recruiting. The FERU thus represents the solution to a social planner's optimization problem, where labor is allocated to maximize social welfare.

To derive the FERU, the authors make the following assumptions. First, the social planner allocates labor to maximize social welfare, with social output as the determinant of social welfare. Second, the social planner has the entire labor force at their disposal for production, assuming a fixed or acyclical labor force participation rate. Third, filling a job vacancy requires one full-time worker. Fourth, the net value of job seekers' home production, accounting for the psychological cost of idleness, is negligible. Fifth, the Beveridge curve has a rectangular hyperbolic shape, implying that unemployed job seekers and recruiters contribute equally to forming job matches. Sixth, the social planner seeks to maximize social output by minimizing the sum of the unemployment and vacancy rates, subject to the trade-off between the two rates as characterized by the rectangular hyperbolic Beveridge curve. Under these assumptions, the optimal market

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tightness—the ratio of vacancy to unemployment—is one and the FERU is the geometric average of the unemployment and vacancy rates. Between 1930 and 2024, the FERU is fairly stable and suggests a generally slack labor market.

I appreciate the authors' work in developing a rigorous conceptual framework for an indicator of full employment. The formula for the FERU is simple and straightforward, making it accessible for policymakers and forecasters as a starting point for discussions on economic slack. While the validity of the FERU may be questioned due to several bold assumptions used to compute it, the transparent derivation presented in this paper, together with the authors' earlier paper (Michaillat and Saez 2021), provides a solid foundation for evaluating risks associated with the FERU estimate. With this valuable feature in mind, I examine the paper's assumptions and discuss their implications for assessing economic slack.

MATCHING FUNCTION ELASTICITY The assumptions about the matching function underlying the observed Beveridge curve are critical for estimating the FERU, as the optimal degree of labor market tightness depends primarily on the elasticity parameters of the matching function. Based on the observed relationship between the vacancy and unemployment rates, the authors posit that the structural Beveridge curve implied by the matching function is a rectangular hyperbola. This assumption implies that job seekers and recruiters contribute equally to job matching, with the elasticity parameters equal to 0.5 in a constant returns to scale (CRS) matching function. However, different matching functions can generate the observed rectangular hyperbolic Beveridge curve, as demonstrated below.

Consider a two-state model of unemployment with a time-varying separation rate and a CRS matching function of the following form:

$$H = mv^{1-\alpha}u^{\alpha},$$

where H is hires as a fraction of the labor force, v is the vacancy rate, u is the unemployment rate, α is the matching elasticity with respect to unemployed job seekers, $1 - \alpha$ is the matching elasticity with respect to vacancy postings or recruiters, and m is the matching productivity. For notational convenience, the time subscript is suppressed.

Dividing both sides by u , the job-finding rate f , is written as follows:

$$(1) \quad f = m \left(\frac{v}{u} \right)^{1-\alpha}.$$

Note that $\frac{v}{u}$ is the market tightness (θ). With the job separation rate (s), the unemployment rate is approximated with the following formula:

$$(2) \quad u = \frac{s}{s + f}.$$

Plugging equation (1) into equation (2), the unemployment rate is expressed as:

$$(3) \quad u = \frac{s}{s + m \left(\frac{v}{u} \right)^{1-\alpha}}.$$

One can solve for the equilibrium market tightness and unemployment rate given the vacancy and separation rates when the elasticity parameter and matching productivity are fixed. Equation (3) implies there exists a set of $[\alpha, m, s]$ values that satisfy the rectangular hyperbolic Beveridge curve relationship, $uv = A$.² For instance, even when the structural Beveridge curve is not a rectangular hyperbola ($\alpha \neq 0.5$) and hence optimal market tightness is not one, this model with a time-varying s can still produce a rectangular hyperbolic relationship between v and u .

For example, consider an economy with $\alpha = 0.7$ and $m = 0.3$. In period 1, the vacancy rate is 4 percent and the separation rate is 1.25 percent, which results in a job-finding rate of 30 percent, an unemployment rate of 4 percent, and market tightness of one. Suppose that a contractionary shock hits the labor market in period 2, lowering the vacancy rate to 2 percent and raising the separation rate to 1.75 percent. The job-finding rate drops from 30 percent to 20 percent, and the unemployment rate rises to 8 percent. Note that the relationship between v and u still lies along a rectangular hyperbola, as the vacancy rate doubles when the unemployment rate is halved. In this economy, the social planner's problem is to minimize $(v + u)$ subject to $H_0 = mv^{0.3} u^{0.7}$, where H_0 is a constant. Since unemployed workers

2. Intuitively, plugging in $v = \frac{A}{u}$ into equation (3) gives the unemployment rate as a function of $[\alpha, m, s]$. Previous studies point to cyclicity in the job separation rate (e.g., Fujita and Ramey 2009; Ahn and Hamilton 2020) and in matching productivity (e.g., Barnichon and Figura 2015; Mukoyama, Patterson, and Şahin 2018).

contribute more to job matches than recruiters, the social planner wants a higher u relative to v , resulting in optimal market tightness of $\theta^* = \frac{3}{7}$ and hence a higher FERU than the baseline. Note that market tightness at full employment is essentially determined by the elasticity parameters. The countercyclical separation rate helps fit a rectangular hyperbolic empirical Beveridge curve.³

This example demonstrates that the level of the FERU and the optimal level of labor market tightness can vary significantly depending on the structural characterization of the Beveridge curve. In this context, a more generalized matching function could offer additional insights into the determination of the FERU. The authors address some of these generalized cases in their 2021 paper and provide an interval of FERU estimates based on the structural parameterization in this paper, which I find very useful.

FULL EMPLOYMENT AND CYCLICALITY OF LABOR FORCE PARTICIPATION Using the FERU estimate, one can directly estimate the level of full employment, in particular the employment-to-population (EPOP) ratio at full employment. The EPOP ratio—a key cyclical indicator frequently referenced by policymakers and forecasters (Abraham and Kearney 2020)—is defined as follows:

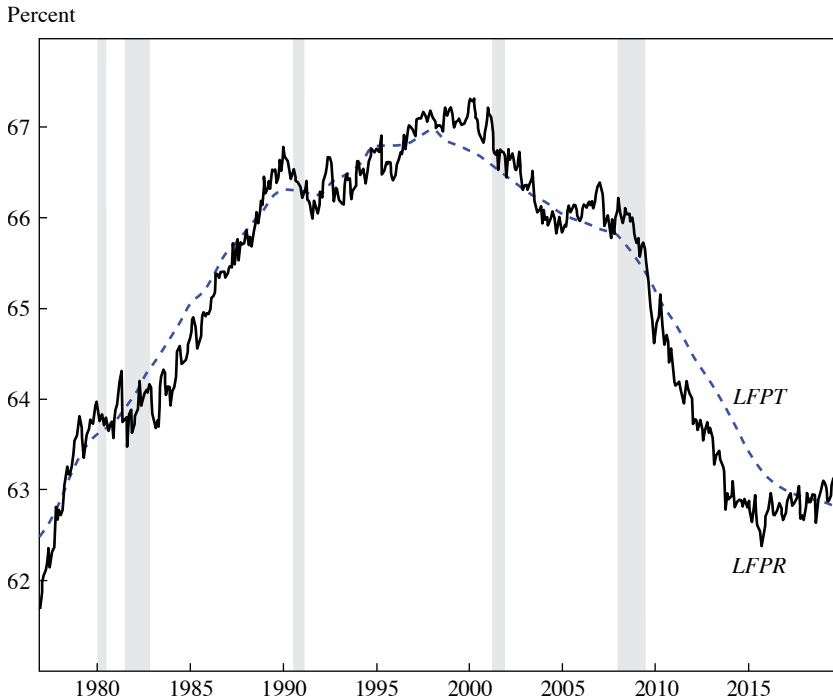
$$(4) \quad EPOP = (1 - u) \times LFPR,$$

where u is the unemployment rate and the LFPR denotes the labor force participation rate. If the LFPR is acyclical, the EPOP ratio at full employment is primarily determined by the FERU. However, if the LFPR moves with the business cycle, the EPOP ratio at full employment could imply a different conclusion about the economy's cyclical position compared to what the FERU indicates.

Recent empirical research suggests that the LFPR is procyclical, though its cyclicality is muted relative to that of the unemployment rate (Cajner, Coglianesi, and Montes 2021; Hobijn and Şahin 2021). Therefore, it is important to distinguish between the cyclical component of the LFPR and its long-term structural trend that reflects demographic or socioeconomic

3. This example benefited from valuable insights by Andrew Figura and is similar to a case discussed in Figura and Waller (2024). Ahn and Crane (2020) report an elasticity estimate that closely aligns with the one considered in this example.

Figure 1. LFPR and LFPR Trend (1976:M12–2019:M12)



Source: Current Population Survey and author's calculations.

Note: The solid line is the LFPR and the dashed line is the LFPT—the trend labor force participation rate from Hornstein and Kudlyak (2019) interpolated as a monthly series. Shaded areas denote NBER recessions.

changes in the population. Figure 1 displays the LFPR alongside its trend estimate (hereafter, LFPT) from Hornstein and Kudlyak (2019).⁴ The LFPR tends to rise above its trend in the mature phase of an economic expansion.

To estimate the cyclically neutral portion of the EPOP ratio, I replace the LFPR with the LFPT and u with the FERU (u^*) to calculate the full-employment EPOP ratio, referred to as the FEFP.

$$(5) \quad FEFP = (1 - u^*) \times LFPT.$$

4. The trend estimate from Hornstein and Kudlyak (2019) is available as an annual series. I interpolate the annual estimate into a monthly time series.

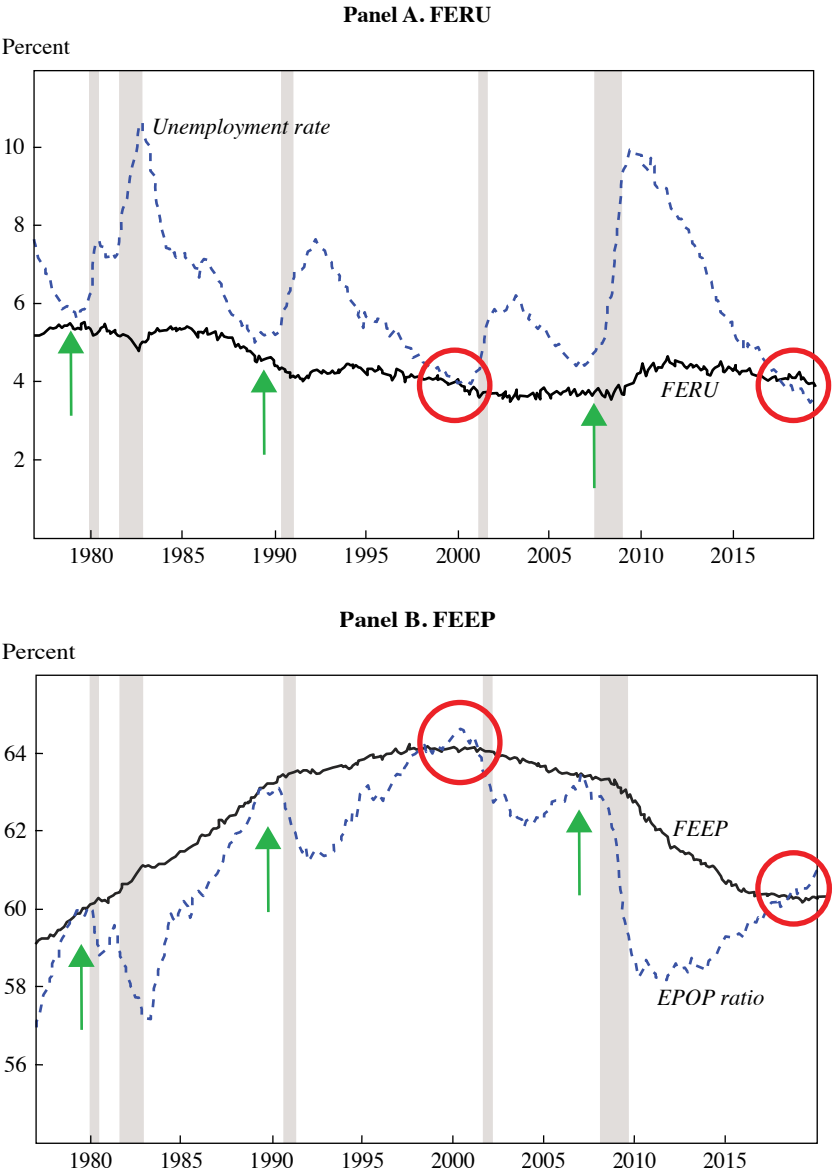
Figure 2 compares the cyclical position in the labor market as measured by the FERU (panel A) and the FEEP (panel B). The circles represent periods when both the FERU and the FEEP indicate full employment, while the arrows mark periods when the FEEP signals full employment but the FERU does not. As shown by the arrows, the FEEP indicates that the labor market reaches full employment toward the end of the expansions, while the FERU identifies such episodes. This implies that the FERU may overestimate overall economic slack compared to the FEEP.

The growth of the FEEP can serve as a benchmark for payroll employment gains, a key indicator of cyclical momentum in the labor market. The cyclically neutral pace of employment gains offers a valuable insight into the economy's cyclical position. The FEEP gains are defined as the first difference in the level of full employment, calculated by multiplying the FEEP by the civilian noninstitutional population.⁵ These gains represent the employment increases needed to sustain full employment when the LFPR is cyclically neutral and given population growth. Figure 3 compares the FEEP gains to total payroll gains, revealing a notable pattern: Payroll gains exceed the FEEP gains during economic expansions but fall below them during downturns. As an expansion matures and the economy approaches a recession, payroll gains converge toward the FEEP gains. This pattern suggests that the FEEP gains can indicate both the economy's cyclical position and the risk of a recession. Looking at the COVID-19 era in figure 4, the FEEP suggests that the economy was at full employment from 2022 (panel A), but the narrowed gap between payroll gains and FEEP gains in 2023 (panel B) signals an increased risk of an economic downturn in recent years.

Highlighting the benefits of the simple and transparent FERU formula and its practical utility, I have discussed how the authors' FERU can be extended to estimate the level of full employment. The cyclicity of the LFPR plays a crucial role in assessing full employment. While the authors demonstrate that their FERU formula remains robust to LFPR cyclicity, an extension of FERU that accounts for the cyclicity of participation would offer a more comprehensive evaluation of whether *employment* has reached full employment.

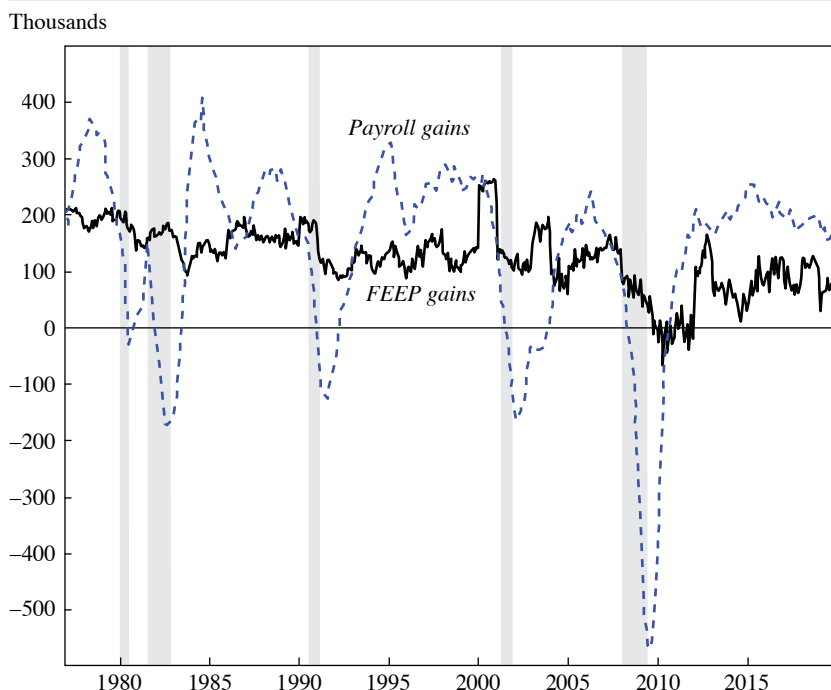
5. I do not explicitly address the difference between employment gains reported in the household survey and payroll gains from the establishment survey, as this difference varies significantly over time.

Figure 2. FERU and FEFP (1976:M12–2019:M12)



Source: Current Population Survey and author's calculations.

Note: In panel A, the dashed line represents the published unemployment rate, while the solid line shows the FERU. In panel B, the dashed line represents the published EPOP ratio, and the solid line depicts the FEFP. The circles highlight the periods when both measures indicate full employment, and the arrows mark the periods when the FEFP suggests full employment but the FERU does not. Shaded areas denote the NBER recessions.

Figure 3. FEEP Gains (1976:M12–2019:M12)

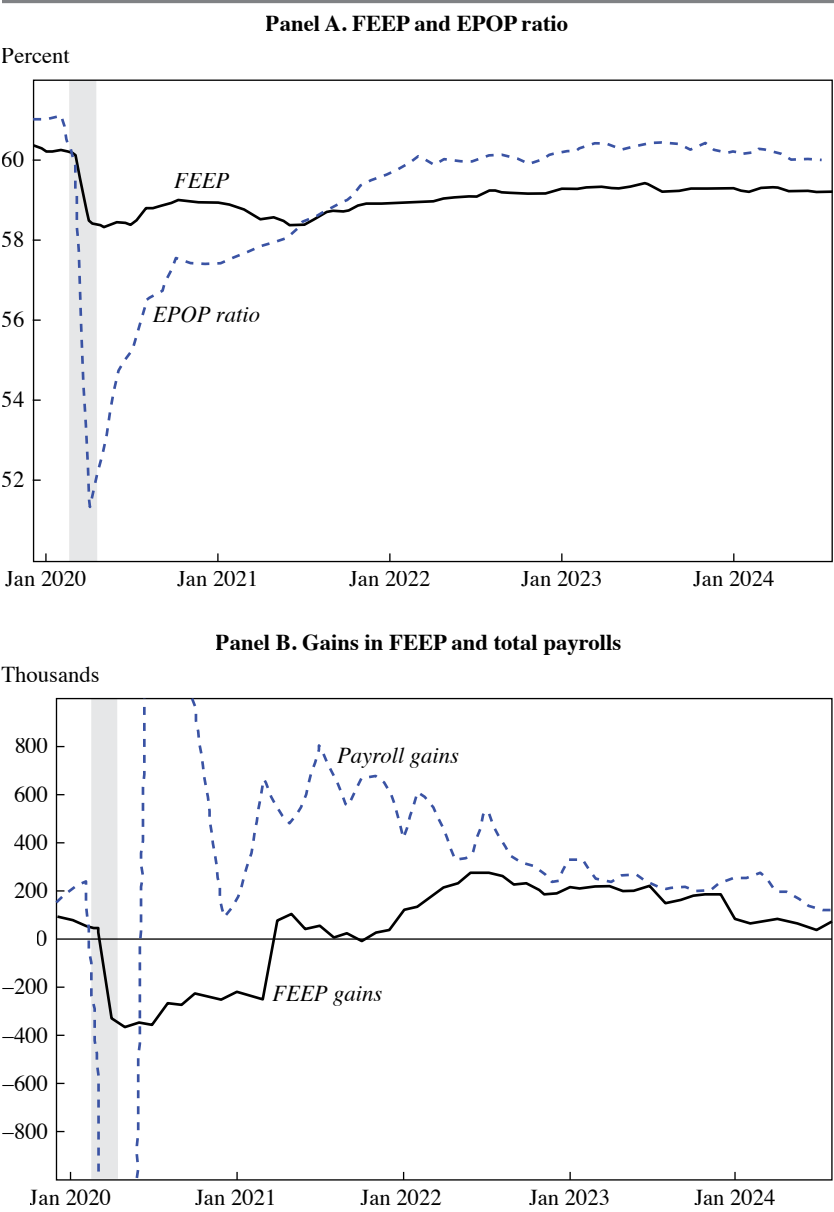
Source: Current Establishment Survey and author's calculations.

Note: The dashed line represents total payroll gains, while the solid line depicts the FEEP gains. Both series are presented as twelve-month moving averages. Shaded areas denote NBER recessions.

BROAD-BASED AND INCLUSIVE FERU Let me now discuss whether the FERU can serve as a benchmark for achieving broad-based and inclusive full employment in the labor market.⁶ Reaching this goal requires policymakers to adopt an unemployment benchmark that comprehensively reflects labor market outcomes across diverse worker groups. Incorporating worker heterogeneity into an unemployment rate benchmark is not straightforward, as the authors acknowledge. In this section, I demonstrate that the FERU offers a valuable starting point for assessing the unemployment rate at full employment across different groups of workers.

6. The Federal Reserve updated its monetary policy framework in 2020, redefining maximum employment as a “broad-based and inclusive goal” (Federal Reserve Board 2020). This shift aims to ensure that monetary policy supports economic growth that benefits a wider and more diverse segment of the population.

Figure 4. FEED and FEED Gains During the COVID-19 Era (2020:M1–2024:M8)



Source: Current Establishment Survey, Current Population Survey, and author's calculations.

Note: In panel A, the dashed line is the EPOP ratio and the solid line is the FEED. In panel B, the dashed line is total payroll gains and the solid line is the FEED gains. The three-month moving average of total payroll gains and the twelve-month moving average of FEED gains are plotted. The shaded areas depict the COVID-19 recession.

Let me begin by discussing race. Notably, the unemployment rates for both white and Black workers, as well as the overall vacancy rate, follow a rectangular hyperbolic Beveridge curve relationship (not shown).⁷ Assuming that the core assumptions underlying the FERU hold for both racial groups, I estimate the FERU separately for white and Black workers.⁸ Later, I discuss the caveats of this analysis and explore lessons for estimating a broad-based and inclusive FERU.

Figure 5 presents the FERU and unemployment rates by race, highlighting notable differences in the experience of full employment between the two racial groups. Since 1976, the unemployment rate for white workers has consistently reached its FERU during nearly all mature economic expansions except the one preceding the Great Recession. In contrast, the unemployment rate for Black workers did not reach its FERU level until 2019. This disparity suggests that the experience of full employment is significantly different between white and Black workers, and a policy that attempts to reduce the employment shortfall for Black workers may imply inflationary pressure for the overall economy.

However, there is an important caveat to this conclusion: Black workers have a significantly higher job separation rate compared to white workers, and the elasticity parameters in the matching function may differ between the two racial groups.⁹ If the matching elasticity with respect to unemployment for Black workers is larger than that for white workers—if unemployed Black workers contributed more (relative to vacancies) to job match creation than unemployed white workers—then given Black workers’ higher job separation rate relative to white workers’, the FERU for Black workers could be higher than that for white workers.

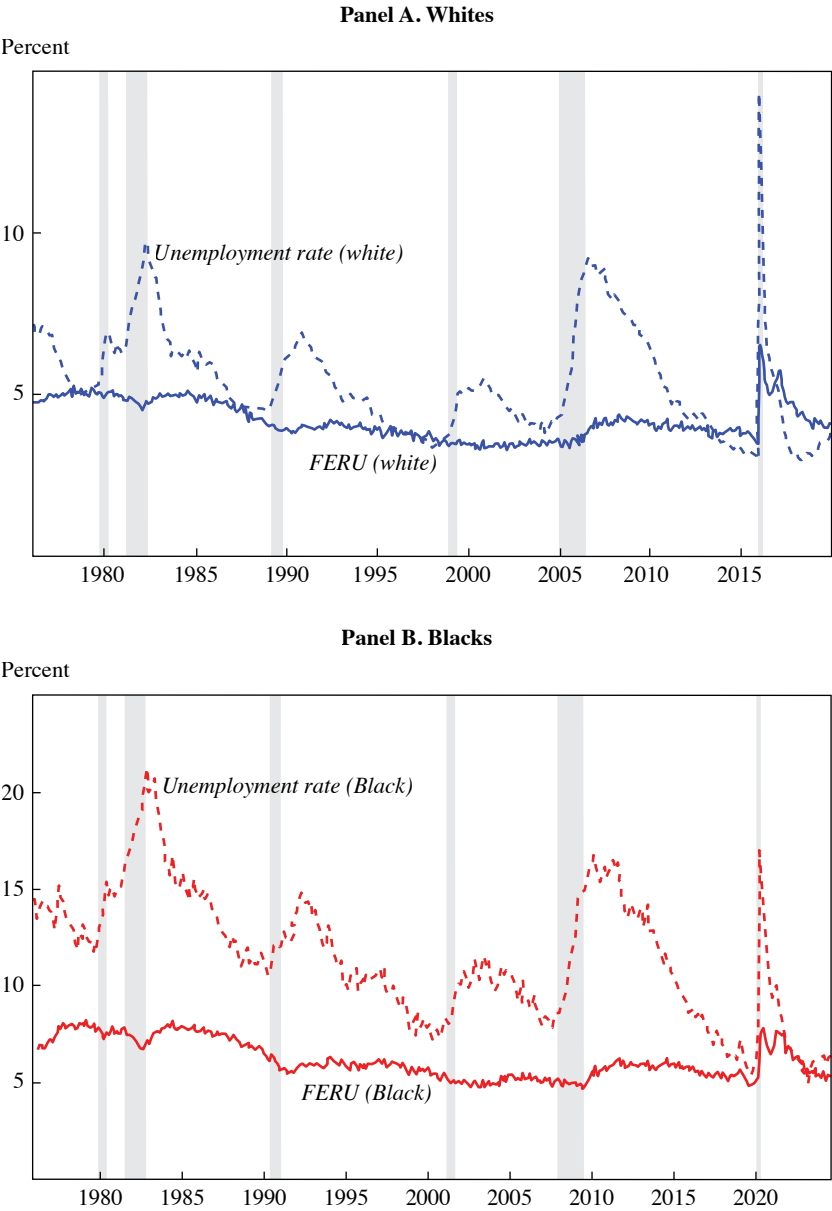
Of course, heterogeneity in labor market outcomes extends beyond race. Indeed, recent research has highlighted that latent heterogeneity not well captured with conventional data has an important effect on labor market dynamics and inequality. For instance, Ahn, Hobijn, and Şahin (2023) demonstrate that the US labor market has the structure of a dual labor market

7. If we take the natural logarithm of the unemployment rates for white and Black workers and compare these values to the logarithm of the vacancy rate, the two racial unemployment rates essentially mirror the logged vacancy rate—a characteristic of a rectangular hyperbola.

8. One assumption for this calculation is that there are no separate job vacancies for white workers and Black workers.

9. Cajner and others (2017) show that the job loss rate is higher among Black workers than white workers, with a significant portion of this disparity remaining unexplained even after accounting for factors such as demographics, education, industry, and occupation. Cairó and Lipton (2023) provide a structural interpretation of the empirical finding from Cajner and others (2017) and assess the effects of monetary policy.

Figure 5. FERU by Race (1976:M1–2024:M8)



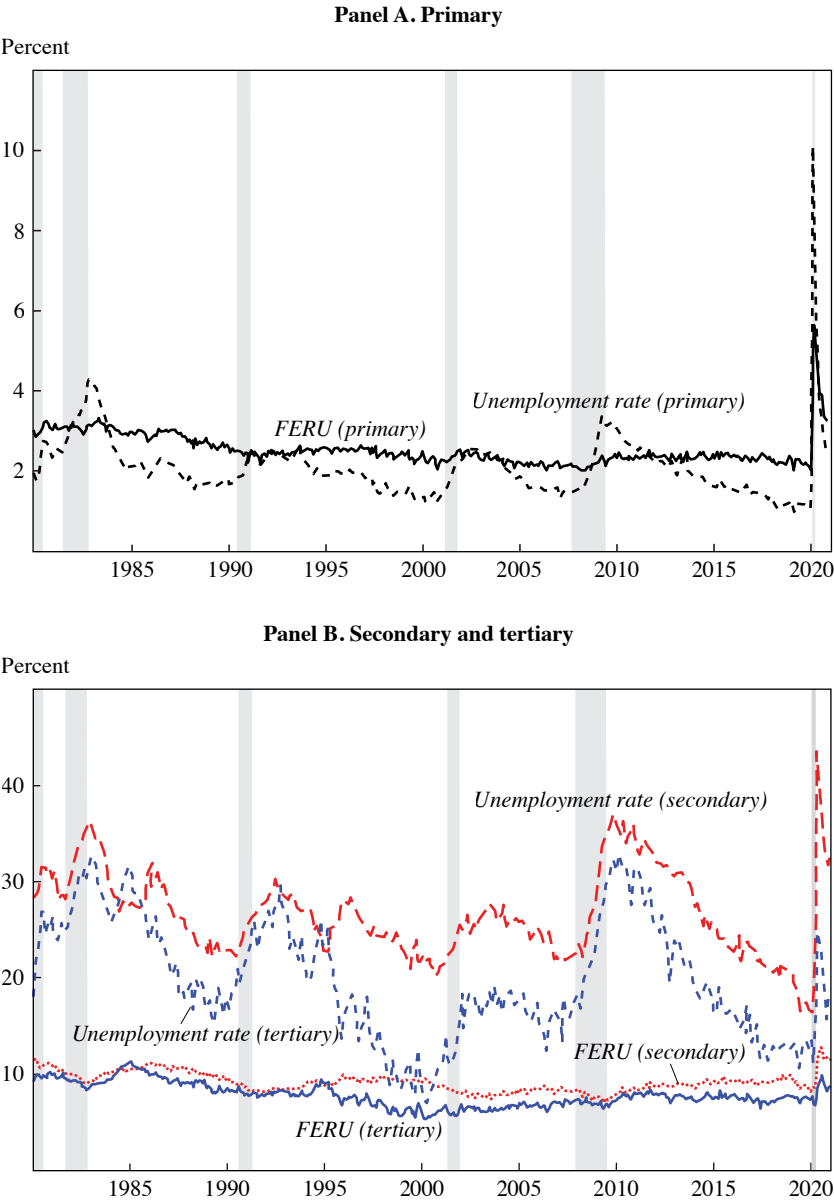
supplemented with a home production segment and hence can be approximated by three latent market segments: the primary market, characterized by job stability and an unemployment rate of around 2 percent; the secondary market, defined by job instability and an unemployment rate of around 25–30 percent; and the tertiary market, which essentially represents a home production segment with low labor force participation and a high unemployment rate of approximately 20 percent. While the unemployment rate of the primary segment is fairly stable, those of the secondary and tertiary markets exhibit strong countercyclicality. Let me make the very strong assumption that the FERU formula can be applied uniformly to all three segments, with workers in each market being exposed to the same aggregate vacancy rate (I will revisit this assumption later). I can then estimate the FERU for each market segment.¹⁰

Figure 6 displays the FERU estimates for the three segments. The full-employment experiences are very different across the segments. The primary market is almost always in full employment and workers in this segment rarely experience a slack market except in a few severe recessions. But the secondary and tertiary segments are always slack. If the assumptions made to obtain these results are close to the truth, the estimates essentially suggest that the baseline FERU may not represent anyone’s full-employment experience.

I am not suggesting that the baseline FERU is not a useful guideline; rather, I want to emphasize that the strong assumptions underlying the analysis in the context of the dual labor market highlight several areas that need further consideration if we want to advance the FERU as a benchmark for broad-based and inclusive labor market growth. One area is heterogeneity in the functioning of labor market. The matching function can vary significantly across different groups of workers with different matching function elasticities among worker groups. Relatedly, job vacancies may be segmented by sector, and recruiting costs can vary across different types of jobs. In addition, the costs of unemployment and the value of home production also differ among workers, further contributing to heterogeneity in what would be full employment for a particular group. A more generalized approach, such as the one proposed in the authors’ earlier paper (Michaillat and Saez 2021), could provide a valuable foundation for developing a conceptual framework that yields a broad-based and inclusive FERU.

10. The unemployment rates of the primary and tertiary segments exhibit a roughly rectangular hyperbolic relationship with the vacancy rate.

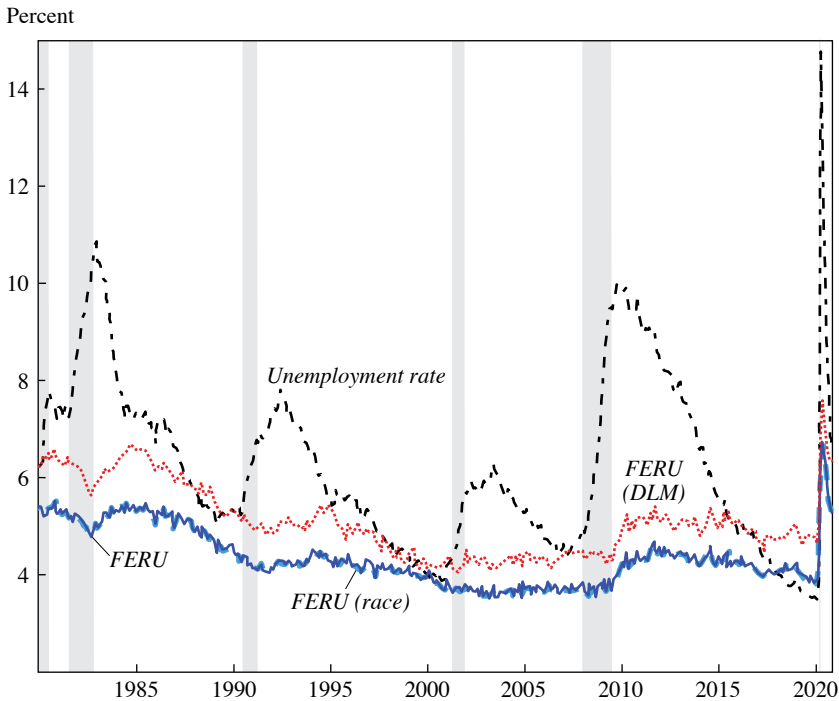
Figure 6. FERU by Segment (1980:M1–2020:M12)



Source: Author's calculations.

Note: In panel A, the solid line is the primary FERU, and the dashed line is the primary unemployment rate. In panel B, the dotted line is the secondary FERU, the long-dashed line is the secondary unemployment rate, the solid line is the tertiary FERU, and the short-dashed line is the tertiary unemployment rate. Shaded areas denote NBER recessions.

Figure 7. FERU and Heterogeneity (1980:M1–2020:M12)



Source: Author's calculations.

Note: The dash-dot line represents the unemployment rate, while the thick solid line depicts the baseline FERU. The dashed line shows the racial FERU, and the dotted line represents the DLM FERU. Shaded areas denote NBER recessions.

To evaluate the aggregate implications of omitted heterogeneity in the calculation of the FERU, I aggregate the FERU of racial groups and that of latent market segments using their respective population weights and then compare these estimates to the baseline.¹¹ I refer to the former as the racial FERU and the latter as the dual labor market (DLM) FERU. As shown in figure 7, the racial FERU essentially matches the baseline estimate, while the DLM FERU is notably higher. This finding suggests that the FERU could be higher than the baseline if labor market heterogeneity is comprehensively accounted for.

11. For this calculation, I consider three racial groups—whites, Blacks, and others—ensuring that the populations of these groups collectively sum to the total population.

MEASUREMENT ERRORS AND BIASES IN THE FERU Last, I discuss effects of data measurement errors on the estimation of the FERU. Previous research, including studies by Abowd and Zellner (1985), Feng and Hu (2013), and Ahn and Hamilton (2022), suggests that the unemployment rate is understated because of the presence of measurement errors.¹² These studies provide bias-adjusted estimates of the unemployment rate, which also exhibit a rectangular hyperbolic relationship with the total vacancy rate (not shown). Building on this empirical observation, I estimate the FERU using each bias-adjusted unemployment rate and examine the implications for assessing labor market slack.

Figure 8 shows the differences between the bias-adjusted unemployment rates and the corresponding FERU estimates, referred to here as the bias-adjusted unemployment rate gaps. These gaps are larger than those calculated using the baseline FERU and the published unemployment rate, especially during periods of high unemployment rate, reflecting the effects of countercyclical measurement errors. This observation suggests that labor market slack may be greater than what is indicated by the baseline FERU when measurement errors are accounted for and highlights the importance of considering measurement errors in the data when assessing the cyclical position of the labor market.

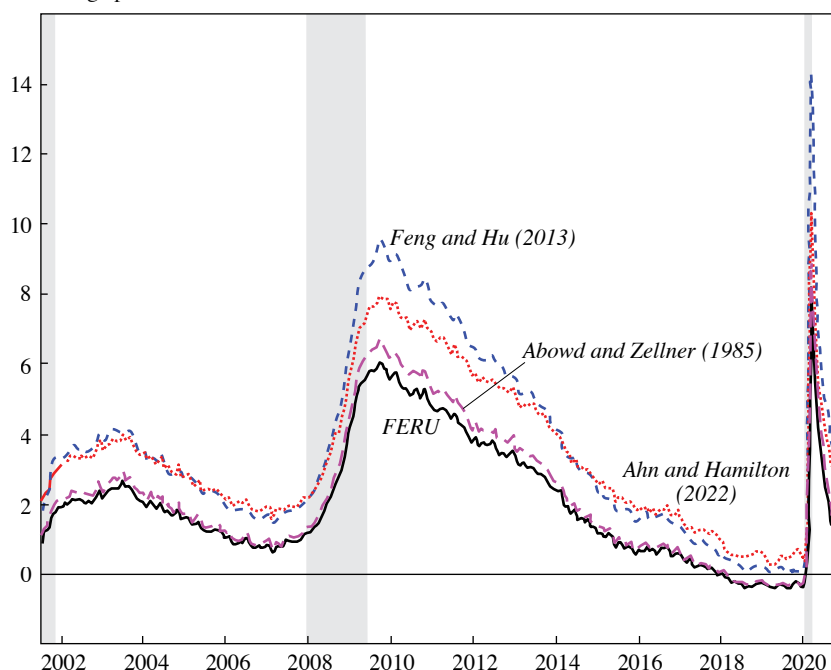
SUMMARY AND CONCLUSION There is much to like about this paper. The simple formula and the transparent assumptions behind the model allow forecasters and policymakers to not only estimate the FERU but also to assess risks associated with the estimate, which is crucial for the risk management aspect of monetary policy. In this discussion, I provide such an assessment of the FERU. The measurement errors imply that the FERU may understate the slack, but all the other cases considered in this discussion suggest that the FERU may overstate the slack.

One final point: The authors have treated the FERU as conceptually distinct from the non-accelerating inflation rate of unemployment and natural rate of unemployment. However, these three measures can empirically coincide with each other. Identifying the theoretical conditions under which they align will be crucial for using the FERU to design and describe optimal monetary policy.

12. The declining response rates in both household and establishment surveys, and particularly in the Job Openings and Labor Turnover Survey (JOLTS), have raised concerns about potential biases in measured unemployment and vacancy rates.

Figure 8. Bias-Adjusted Unemployment Rate Gap (2001:M7–2020:M12)

Percentage points



Source: Author's calculations.

Note: The solid line at the bottom depicts the baseline FERU. The long-dashed line shows the FERU estimated using Abowd and Zellner's (1985) bias-adjusted unemployment rate; the dotted line above is the FERU estimated with Ahn and Hamilton's (2022) bias-adjusted unemployment rate; the top short-dashed line is the FERU estimated with Feng and Hu's (2013) bias-adjusted unemployment rate. Shaded areas indicate NBER recessions.

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COMMENT BY

BART HOBIJN¹ In this paper, with an Einsteinian title, Pascal Michaillat and Emmanuel Saez argue that a policymaker who aims for the efficient combination of unemployment and vacancies on a fixed Beveridge curve should choose that combination where $u^* = v^* = \sqrt{uv}$, that is, a vacancy-unemployment ratio equal to one and each rate equal to the square root of the product of the current unemployment and vacancy rates. This is a remarkably simple policy prescription. I take it as my task in this comment to highlight under which circumstances this policy prescription seems reasonable and when it might be a bit of an oversimplification.

THREE ASSUMPTIONS BEHIND THE SIMPLE POLICY PRESCRIPTION The simple policy prescription in this paper follows three assumptions made. To put these assumptions in the context of the Beveridge space that I will focus on in the rest of my discussion, consider figure 1.

The first assumption behind the policy prescription is that it is desirable for a policymaker to choose the unemployment, u , and vacancy, v , rates to minimize their sum. This means that there is a one-for-one trade-off between these two rates reflected by the isocost curves in figure 1, which are lines with a slope equal to minus one.

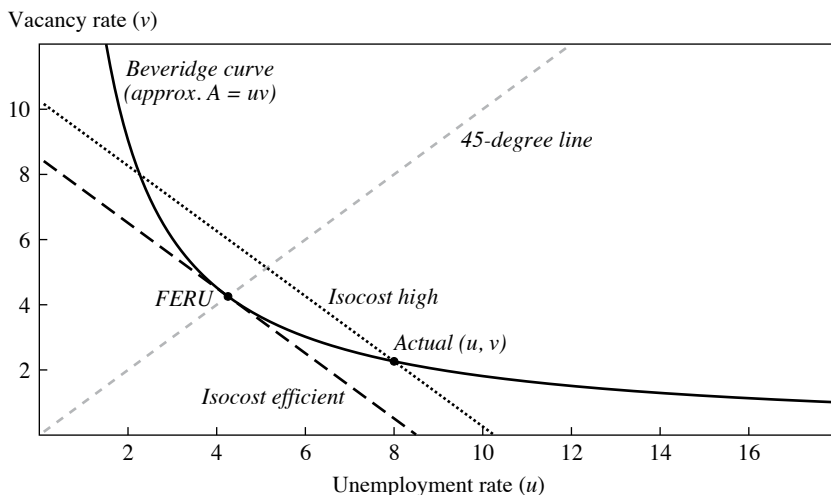
The policymaker cannot choose u and v independently. At the chosen unemployment rate, they need to choose the vacancy rate at which the unemployment rate remains constant. The combinations of u and v for which this is the case form the theoretical Beveridge curve. The second assumption made is that the policymaker's decisions keep the Beveridge curve fixed, that is, policies move the labor market equilibrium along the Beveridge curve and do not affect the position of the Beveridge curve.

The third assumption is that the Beveridge curve is accurately approximated by the hyperbola on which the product of the unemployment and vacancy rates is constant, that is, $A = uv$.

These three assumptions yield the cost minimization problem, illustrated in figure 1, which determines the efficient levels of the unemployment and vacancy rates.

$$(1) \quad \text{Choose } (u, v) \text{ to minimize } u + v \text{ subject to } A = uv.$$

1. Preparation of these comments has benefited from extensive discussions with Gadi Barlevy, Andre Kurmann, Tristan Potter, and Aysegül Şahin. Any opinions, findings, and conclusions or recommendations expressed in this material are my own and do not necessarily reflect the views of the Federal Reserve Bank of Chicago or the Federal Reserve Board of Governors.

Figure 1. Three Main Assumptions Behind Policy Prescription in Beveridge Space

Source: Author's calculations using illustrative parameter values.

The solution to this problem, (u^*, v^*) , is the FERU that satisfies $u^* = v^* = \sqrt{A} = \sqrt{uv}$. The latter is true if the current levels of the unemployment and vacancy rates, u and v , are on the Beveridge curve themselves. In the rest of this comment, I discuss the three underlying assumptions that drive this simple policy prescription, explain when they hold and when not, and illustrate the impact of deviations from these assumptions.

SUM OF UNEMPLOYMENT AND VACANCIES AS OBJECTIVE In terms of efficiency and desirability, the main assumption of this paper is that a policy-maker should aim to minimize the sum of the unemployment and vacancy rates. This implies a social marginal rate of transformation of one between unemployed persons and job openings that is constant over time. This is definitely a simple and transparent objective. But I don't think that its pursuit is socially desirable.

Ambiguous unit of measurement of vacancies. Considering the sum of unemployment and vacancies suggests that they have the same unit of measurement. But this is not the case. The unit of measurement of vacancies is ambiguous. The measure of vacancies used in the paper is the series by Petrosky-Nadeau and Zhang (2021), which merges the job openings measure for the Job Openings and Labor Turnover Survey (JOLTS) with a composite help wanted measure of online and newspaper job postings. These two merged series themselves already have different

units of measurement. But let me focus on the job openings measure from JOLTS.

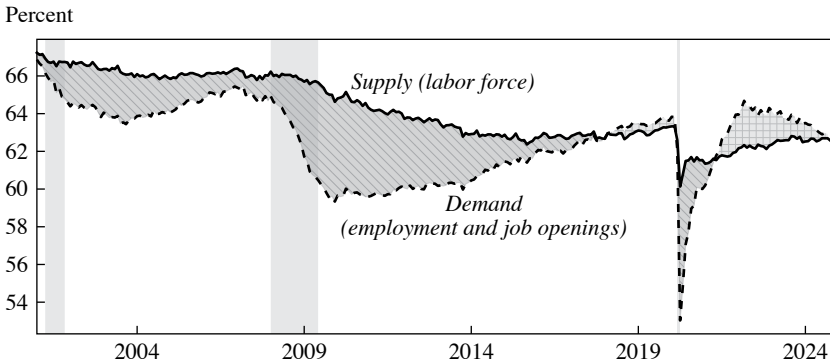
The Bureau of Labor Statistics defines job openings in JOLTS as “all positions that are open (not filled) on the last business day of the month. A job is ‘open’ only if it meets all three of the following conditions: (1) A specific position exists and there is work available for that position. The position can be full-time or part-time, and it can be permanent, short-term, or seasonal, and (2) the job could start within 30 days, whether or not the establishment finds a suitable candidate during that time, and (3) there is active recruiting for workers from outside the establishment location that has the opening.”²

This definition is important because it points to why summing the number of job openings and the number of unemployed is problematic. First of all, the definition clearly states that job openings are measured in terms of jobs while unemployment is measured in terms of persons. Second, job openings are not limited to ones used to recruit persons out of unemployment. During the Great Resignation of 2021 and 2022, many job openings were filled with workers poached from other employers. Moreover, many job openings are posted to hire high school and college graduates who are out of the labor force rather than unemployed. Importantly, the cost per hire estimates used in this paper don’t distinguish between costs for hiring an unemployed person versus someone from another employer or from out of the labor force.³ Third, job openings in JOLTS capture vacancies for jobs in which workers can start within thirty days. This requirement excludes many vacancies. If you are a government employee or academic reading this comment, you most likely applied for a vacancy that did not satisfy this criterion. Finally, many people are hired without a job opening. For example, Davis, Faberman, and Haltiwanger (2013) report that, in the JOLTS data from 2001–2006, 42 percent of hires occur at establishments that do not report any vacancies.⁴

2. Bureau of Labor Statistics, “Job Openings and Labor Turnover Survey,” under “How Does JOLTS Define Job Openings?” <https://www.bls.gov/jlt/jltdef.htm>.

3. Estimates of the distribution of cost per hire from the Society for Human Resource Management (SHRM) show that this distribution is severely right skewed with the mean being higher than the 75th percentile (Miller 2022). This suggests the average cost per hire estimates used in this paper are largely driven by hires of high-skilled workers from other employers, which are not part of the u versus v trade-off the paper considers. These costs are also countercyclical.

4. An alternative interpretation that Michaillat and Saez provide about the unemployment and vacancy trade-off that they focus on is that one unemployed worker is the equivalent of one worker spending time on filling a vacancy. But the evidence in Davis, Faberman, and Haltiwanger (2013) implies that a lot of recruitment efforts are not captured by vacancies. They also provide evidence that the intensity with which firms pursue filling vacancies varies over the business cycle and is not constant, as assumed in this paper.

Figure 2. Balance Between Supply and “Demand” in the Labor Market

Source: Bureau of Labor Statistics and author's calculations.

Note: Supply in the labor market is measured as the labor force and “demand” is measured as the number of persons employed plus the number of job openings for nonfarm payroll jobs. Both are measured as a percentage of the noninstitutional population of age 16 and over.

This means that I am not sure how to interpret the policy prescription of a vacancy to unemployment ratio, v/u , equal to one based on the JOLTS job openings measure in this paper. This is important for the conclusion in this paper that, on average, the unemployment rate in the United States has been inefficiently high.

To see this, consider figure 2, which has been used, in other contexts, to illustrate how labor supply and labor demand have come into better balance over the past year and a half. The labor demand measure in the figure interprets JOLTS job openings as the level of unmet labor demand, just as is done in this paper. By definition, the difference between supply and demand in the figure is the difference between the levels of unemployment and vacancies (as a share of the population). In this figure, balance in the labor market, that is, measured supply equals measured demand, happens when $u = v = \sqrt{uv}$ and the FERU condition introduced in this paper is satisfied.

Depending on how you interpret this figure, it indicates that, on average, there is excess supply in the labor market or that the unemployment rate is higher than the FERU, as Michaillat and Saez argue in this paper. But the interpretations rely on an incorrect interpretation of the JOLTS job openings measure. This imbalance is most likely the result of the JOLTS job openings measure not capturing all forms of unmet labor demand in the economy.

This does not mean the JOLTS measure of job openings is not useful. Many measures of unfilled vacancies and unmet labor demand move together. The high degree of comovement between the help wanted measure and JOLTS job openings is what inspired the splicing of help wanted with JOLTS data (Barnichon 2010, fig. 1). More recently, data on job postings by Indeed have moved similarly to the JOLTS measure as well. So, job openings in JOLTS are a good measure of the *fluctuations* in unmet labor demand but *not necessarily of its level*.

Theoretical desirability of minimizing sum of unemployment and vacancies. The paper includes (in section II.H) an example of how the FERU can be derived in the context of the model from Pissarides (2000, chap. 1) in the particular case where the vacancy posting cost moves one-for-one with the productivity level in the economy. It shows that, if that is the case and the government policy is to choose a combination (u, v) , then it would be optimal to choose the FERU. What the paper does not do is solve the equilibrium of a full version of the model with a labor demand part that generates a job creation curve.

A natural way to close the model to obtain such a full version is to impose the condition derived by Hosios (1990). It assures that all equilibrium points on the Beveridge curve are Pareto efficient.⁵ Closing the stylized model and imposing the Hosios condition yields that, if the vacancy posting cost is proportional to the level of productivity, then there are no equilibrium fluctuations in the unemployment and vacancy rates in response to productivity shocks.

The intuition for this is that the constant returns to scale production and matching technologies mean that the relative allocation of workers between u and v does not change when both technologies shift at the same proportion. At first glance, this result might seem more of a theoretical curiosity.⁶ But it illustrates that what is important for unemployment fluctuations in search and matching models is exactly the type of fluctuations in the relative cost of recruitment that the authors assume away when they argue the government should focus on minimizing the sum of the unemployment and vacancy rates.

Though the stylized model in the paper has counterfactual implications, it does provide a useful way to think about optimal policy in a search and

5. If the Hosios condition does not hold, points on the Beveridge curve are not (necessarily) Pareto efficient, and thus the assumption that optimal policy should choose a point on the Beveridge curve, made in this paper, is not necessarily correct.

6. This is a particular example of the importance of cyclical fluctuations in vacancy posting costs in these models emphasized by Pissarides (2009).

matching model. For example, one can think of the government choosing (p, u, v) to maximize steady-state welfare

$$(2) \quad (1 - u)p - cv - \phi(p),$$

subject to the Beveridge curve

$$(3) \quad u = \frac{1}{1 + \frac{\omega \left(\frac{v}{u}\right)^\alpha}{\lambda}},$$

and the job creation condition

$$(4) \quad cv = \omega \left(\frac{v}{u}\right)^\alpha J(p, u).$$

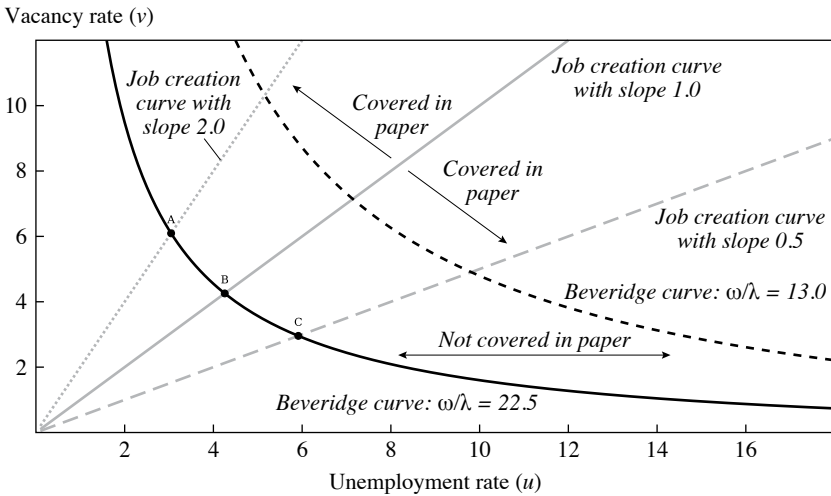
Here, equation (3) is the commonly used functional form for the Beveridge curve (e.g., Pissarides 2000, chap. 1), rather than the hyperbola approximation used by Michaillat and Saez. In particular, λ is the separation rate out of employment, ω is the level of match efficiency, and α is the elasticity of the matching function. The function $J(p, u)$ is the value of a job match for an employer.

The additional term in the welfare function (2), $\phi(p)$, is included to capture welfare costs of stimulating labor demand, that is, choosing a higher p , beyond its impact on output and unemployment. For example, it includes the cost of inflationary pressures resulting from increases in the v/u ratio due to rising labor demand (Ball, Leigh, and Mishra 2022; Barnichon and Shapiro 2024; Benigno and Eggertsson 2023).

Just like the welfare maximization problem solved by Michaillat and Saez, the above problem, through restriction—equation (3)—assumes that policies move the labor market equilibrium along the Beveridge curve. I consider this assumption next.

POLICIES THAT KEEP THE BEVERIDGE CURVE FIXED AND THAT SHIFT IT The simple policy prescription in this paper results from the policymaker choosing the optimal combination of unemployment and vacancies on a fixed Beveridge curve. Though the paper is not explicit about how this would be achieved, it suggests that this is done using policies that pin down job creation at the level at which it results in the desired equilibrium outcome. Michaillat and Saez call this outcome the FERU, because they interpret

Figure 3. Policies Covered and Not Covered by Prescription in Beveridge Space



Source: Author's calculations using illustrative parameter values.

it as the choice that the federal government and the central bank should implement to maintain the economy at “full employment,” or “maximum employment,” as mandated in the Employment Act of 1946, the Federal Reserve Reform Act of 1977, and the Full Employment and Balanced Growth Act of 1978.

But even if one thinks the social objective is to minimize the sum of unemployment and vacancies, the policy prescription in this paper is only applicable to policies that affect job creation and do not shift the Beveridge curve. Among the many policies that the federal government implements there are certainly some for which this might be the case. But many policies are specifically aimed at moving the Beveridge curve, and the optimal policy derived in this paper has little to say about the desirable outcomes for these programs.

To put this in context, consider figure 3. It shows a textbook diagram in Beveridge space. For a given Beveridge curve, for example, the solid curve in the diagram, the equilibrium combination of unemployment and vacancies is given by the intersection of this curve with the job creation curve. The figure shows three such curves. The policy prescription in this paper covers policies that solely move the job creation curve and do not affect the position of the Beveridge curve. It does not cover policies that shift the Beveridge curve.

Policies that move the economy along the Beveridge curve. The position and shape of the Beveridge curve reflect unemployment inflows as well as search and matching technology in the labor market, as captured, for example, in the functional form used in equation (3). So, policies that do not affect these things will keep the Beveridge curve fixed.

Probably the most prominent example of this is monetary policy. For example, Figura and Waller (2022) consider the likelihood of a “hard landing” after the monetary tightening that started in 2022 by studying the shape of the Beveridge curve and how much job creation can decline, and the vacancy rate come down, before the Beveridge curve flattens out. Their analysis assumes that the Beveridge curve remains fixed in response to the monetary tightening.⁷

Policies that move the Beveridge curve. The policy prescription in this paper is not meant to be solely applicable to monetary policy but to the federal government’s maximum employment objective in general. It is important to realize, however, that many federal policies affect the position and shape of the Beveridge curve. The simple recipe provided in this paper is of limited use for evaluating the efficiency of those policies. For example, there is an extensive literature about the incentive effects of federal extensions of unemployment insurance (e.g., Farber, Rothstein, and Valletta 2015; Chodorow-Reich, Coglianese, and Karabarbounis 2019, among many). Though most studies find only limited disincentive effects, they do suggest that unemployment insurance extensions result in a small decline in match efficiency that shifts the Beveridge curve rightward. The method applied in this paper would spuriously interpret this rightward shift as an increase in the FERU. The question about the optimal level of unemployment insurance weighs the costs of benefit payments and longer unemployment spells, and a corresponding rightward shift in the Beveridge curve, against the benefits of a reduction in income uncertainty and a potential increase in the quality of matches—a very different policy trade-off than the unemployment versus vacancies choice considered in this paper.

The policy that is probably most directly focused on affecting the position of the Beveridge curve is the Workforce Innovation and Opportunity Act (WIOA).⁸ Its main purpose is “to increase, for individuals in the United States, particularly those individuals with barriers to employment, access to and opportunities for the employment, education, training, and support

7. But “hard landings” after monetary tightening episodes often involve increases in the separation rate, which at least temporarily shift the Beveridge curve outward.

8. Workforce Innovation and Opportunity Act, Pub. L. 113–128, 128 Stat. 1425 (July 22, 2014).

services they need to succeed in the labor market” (sect. 2.1). As part of it, local workforce development boards work together with federal, state, and local governments, employers, educational institutions, and nonprofits to enhance economic opportunity for workers, including matching job searchers with local employers. It provides a decentralized nationwide institutional framework for active labor market policies in the United States.

Cross-country variation in institutions like WIOA, labor market regulations, education systems, and taxation has been studied extensively to understand differences in the locations and persistent shifts of Beveridge curves across countries. This has been particularly true when it relates to “Eurosclerosis,” that is, the lasting increase in unemployment rates in many European countries following the recessions of the early 1980s (e.g., Nickell 1997; Blanchard and Wolfers 2000). Though “Eurosclerosis” has disappeared from the radar screen of most academic researchers, many of the current cross-country differences in unemployment rates between advanced industrialized economies are rooted in that period.

A normative framework to assess the relative efficiency of labor market outcomes across countries would be very useful to shape the policy discussion about these differences. Though the policy prescription in this paper cannot be used as such a framework, it does remind us of the importance of thinking about efficiency and desirability of labor market outcomes, not only in the United States but across countries.

QUALITY OF BEVERIDGE CURVE APPROXIMATION The third main assumption made in the paper is that the Beveridge curve is well approximated by the hyperbola on which the product of the unemployment and vacancy rates is a constant, $A = uv$. But conceptually, the Beveridge curve is given by the combinations of the unemployment and vacancy rates at which the inflows into unemployment equal the outflows and the unemployment rate is in its flow steady state.

Equation (3) captures these combinations for a constant separation rate, λ , and a Cobb-Douglas matching function with match efficiency ω and elasticity α . This is the most commonly used functional form for the theoretical Beveridge curve. Taking the total differential of this curve, we obtain that along it

$$(5) \quad \left[1 - (1 - u)\alpha\right] d \ln u = - \left[(1 - u)\alpha\right] d \ln v.$$

Thus, when $(1 - u)\alpha \approx 1/2$ then $d \ln u \approx -d \ln v$. In that case the product of the unemployment and vacancy rates is approximately constant along the

curve. This is when the approximation of the Beveridge curve, $A = uv$, used in this paper works relatively well. As the authors point out in equation (4) of the paper, because u is relatively small in the United States, this condition holds in the United States when the elasticity of the matching function is approximately equal to $1/2$.

Of course, what is relevant for the policy recommendation is not how well the hyperbola, $A = uv$, approximates the Beveridge curve, but how close the policy prescription in the paper is to the one with the actual Beveridge curve. For the functional form in equation (3), the latter is the one that solves the problem

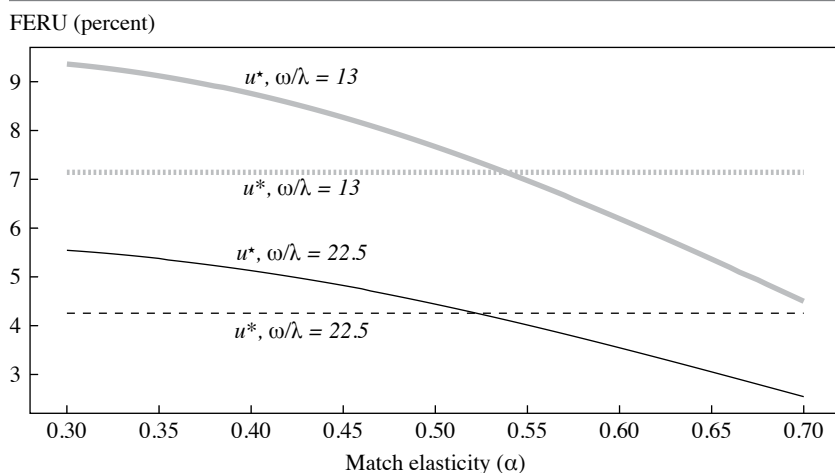
$$(6) \quad \text{Choose } (u, v) \text{ to minimize } u + v \text{ subject to equation (3).}$$

Its solution is given by

$$(7) \quad u^* = \frac{1}{1 + \frac{\omega \left(\frac{v^*}{u^*} \right)^\alpha}{\lambda}} \quad \text{and} \quad v^* = \frac{(1 - u^*)^\alpha}{1 - (1 - u^*)^\alpha} u^*.$$

This solution depends on two parameters: the ratio of match efficiency and the separation rate, ω/λ , which determines the location of the Beveridge curve, and the elasticity of the matching function, α , which determines its shape. The approximate solution in this paper, that is, the FERU solution, only depends on the location of the Beveridge curve and equals $u^* = 1/(1 + \omega/\lambda)$.

Figure 4 compares the actual solution, u^* , with the approximate solution, u^* , for different locations and shapes of the Beveridge curve. The black lines in the figure correspond to the location of Beveridge curve that is similar to recent estimates for the United States. In that case, $u^* = u^* = 4.25$ for $\alpha = 0.52$. At $\alpha = 0.5$, $u^* = 4.40$, which is a relatively small 0.15 percentage point deviation from the approximation. However, at the value $\alpha = 0.3$, used by Figura and Waller (2022), $u^* = 5.5$ and the approximation used in the paper is off by 1.25 percentage points—the equivalent of about two million unemployed persons. When the Beveridge curve is shifted further outward, u^* approximates u^* less well. This can be seen from the gray lines in figure 4, which illustrate this for a case where u^* is around 7 percent, as is approximately the case in France for example.

Figure 4. Policy Prescriptions for Different Assumptions of Shape of Beveridge Curve

Source: Author's calculations.

The point is that the hyperbola approximation that Michailat and Saez use only yields a solution that is close to the FERU obtained using the most commonly used functional form for the Beveridge curve under a very limited set of parameter values.

TO SUMMARIZE The formula $u^* = \sqrt{uv}$ might not be the “ $E = mc^2$ ” of economics. But the simple policy prescription in this paper serves as a useful starting point for furthering the discussion of how to think about the federal government’s task of maximizing employment and what types of conceptual and normative frameworks provide context for it. This type of discussion is important if we would like to be more specific about what maximum employment means than “a broad-based and inclusive goal that is not directly measurable” (Federal Reserve Board 2020).

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GENERAL DISCUSSION Robert Hall acknowledged the contributions of the authors to the growing literature demonstrating that, contrary to the characterization typical of monetary policy discussions, the natural rate of unemployment u^* is an endogenous variable rather than a constant. He noted that the amount of variability of u^* estimated by the authors is less

than other estimates, including those in his working paper with Marianna Kudlyak.¹

A number of participants commented on data and measurement issues regarding unemployment and vacancy rates. Hall described that during the COVID-19 pandemic, official statistics counted those workers who had kept their jobs but were only temporarily laid off as unemployed. Without accounting for this group, the Diamond-Mortensen-Pissarides (DMP) model appears to break down, when in reality, researchers are applying the model to the wrong group of people. Hall emphasized that this group is irrelevant to the matching process, as they remain attached to their employers. He pointed to his paper with Kudlyak that parses out this group in the unemployment data.²

Regarding vacancy rates specifically, Steven Davis reiterated the discussion Bart Hobijn's point that using the Bureau of Labor Statistics Job Openings and Labor Turnover Survey (JOLTS) vacancy data might underestimate the true level of unmet labor demand. He pointed to his previous work showing a systematic undercount of vacancies in the JOLTS data for two reasons.³ First, respondents in the JOLTS sample underreport vacancies relative to hires. Second, the JOLTS sample misses the new and very young employers. These employers have high vacancies relative to their share of employment, because they are new and growing.⁴ Based on his previous research, Davis speculated that the reported JOLTS vacancy rate is about 15–18 percent below the true vacancy rate. Taken together with Hall's point about measuring unemployment during the pandemic, Davis argued that there are significant first-order measurement issues to address in order to operationalize the formula the authors present.

Tara Sinclair stressed that if the authors' formulation is to be useful for real-time policymaking, it must take into account the distinct data revision processes of vacancies and unemployment. She also pointed out additional compositional challenges that arise when putting the unemployment and

1. Robert E. Hall and Marianna Kudlyak, "The Active Role of the Natural Rate of Unemployment," working paper 23117 (Stanford, Calif.: Hoover Institution, 2024).

2. Robert E. Hall and Marianna Kudlyak, "The Unemployed With Jobs and Without Jobs," *Labour Economics* 79 (2022): 102244.

3. Steven J. Davis, R. Jason Faberman, and John Haltiwanger, "The Establishment-Level Behavior of Vacancies and Hiring," *Quarterly Journal of Economics* 128, no. 2 (2013): 581–622.

4. More information on the methodology of JOLTS can be found at Bureau of Labor Statistics, "Job Openings and Labor Turnover Survey: Overview," <https://www.bls.gov/opub/hom/jlt/home.htm>.

vacancy rates together, further complicating the application of the authors' formula.

Pascal Michaillat clarified that the paper does not conceive of vacancies as a measure of unmet labor demand, as was originally suggested by Beveridge (1944).⁵ Instead, the authors view it as a way to measure non-productive labor in the form of recruitment, and it happens to be the case that it takes roughly one worker devoted to recruitment to fill one vacancy.

Participants also dug into various assumptions underlying the authors' model. Davis questioned the assumption that there are zero social benefits to unemployment. He pointed to an article in which he and Pawel Krolikowski survey individuals receiving unemployment insurance.⁶ Two-thirds of those surveyed indicated that they would not be willing to continue their old jobs with a 25 percent pay cut. In their sample, the unemployment replacement rate is about 38 percent on average. This means that for at least two-thirds of the sample, their value of time (net of the unemployment benefit) is at least 35 percent of the wage they just lost, possibly much higher. Given this, Davis wondered how sensitive the authors' results are to the assumption that unemployment has zero social value.

Valerie Ramey also questioned the authors' assumption. She pointed to evidence of procyclical death rates since the 1970s, including evidence from Ruhm (2000).⁷ Ramey also argued that the value of unemployment may be particularly high for those with dependents.

Michaillat responded by pointing to his previous work with Emmanuel Saez that relaxes many of the assumptions made in the current paper; for example, allowing for a social value of unemployment that is not zero and for varying costs of servicing a vacancy.⁸ This work also includes a sensitivity analysis of the full-employment rate of unemployment (FERU) to these parameters and others. Using the range of values found in the literature for these parameters, it finds that FERU stays within 1.2 percentage points of the baseline results.

Several participants brought up the importance of inflation in discussions of full employment. William English was struck by the authors' figure

5. William H. Beveridge, *Full Employment in a Free Society* (London: George Allen & Unwin, 1944).

6. Steven J. Davis and Pawel M. Krolikowski, "Sticky Wages on the Layoff Margin," *American Economic Review* 115, no. 2 (2025): 491–524.

7. Christopher J. Ruhm, "Are Recessions Good for Your Health?" *Quarterly Journal of Economics* 115, no. 2 (2000): 617–50.

8. Pascal Michaillat and Emmanuel Saez, "Beveridgean Unemployment Gap," *Journal of Public Economics Plus* 2 (2021): 100009.

showing that the non-accelerating inflation rate of unemployment (NAIRU) measures are much higher than the authors' FERU measures. This reminded him of the work by Barro and Gordon (1983) and the idea that the socially efficient amount of labor may be higher than the amount of employment that can be sustained without generating inflation.⁹ English wondered if this was captured in the authors' model.

Christina Romer took English's comments further, arguing that a concept of full employment that isn't consistent with stable inflation is not a sensible goal for policy. She noted that the marginal groups that get hired when unemployment is low are the same groups that lose their jobs when the Federal Reserve tightens monetary policy due to high inflation. She concluded that trying to divorce the concept of full employment from stable prices is not practical.

Joseph Gagnon discussed the usefulness of the authors' measured employment gap, despite the gap being almost always positive. He pointed to work by both Hobijn and Jón Steinsson showing that when there is downward wage rigidity, this bends the Phillips curve, giving it a flat segment.¹⁰ As a result, the economy tends to operate above the natural rate of unemployment. Moreover, it becomes difficult for the central bank to estimate the natural rate of unemployment. This is another factor leading the economy to operate above the natural rate with no downward pressure on inflation, which can go on for years or even decades. Thus, Gagnon cautioned against dismissing the authors' results based on the persistently high unemployment gap. Gagnon further added to Hobijn's discussion of labor market institutions shifting the Beveridge curve. He argued that accounting for changes in labor market institutions would improve the model, especially when applying it to other countries.

In response to participants' comments on inflation—in particular, Romer's argument that any concept of full employment should consider stable prices—Michaillat argued that there is value to a measure of full employment free from inflation considerations. He noted that nothing guarantees that NAIRU aligns with a socially efficient level of unemployment; it only promises stable prices. He highlighted that the Federal Reserve's dual mandate to promote full employment and stable prices demands a measure

9. Robert J. Barro and David B. Gordon, "Rules, Discretion and Reputation in a Model of Monetary Policy," *Journal of Monetary Economics* 12, no. 1 (1983): 101–21.

10. Mary C. Daly and Bart Hobijn, "Downward Nominal Wage Rigidities Bend the Phillips Curve," *Journal of Money, Credit and Banking* 46, no. S2 (2014): 51–93; Stéphane Dupraz, Emi Nakamura, and Jón Steinsson, "A Plucking Model of Business Cycles," working paper, 2024, <https://eml.berkeley.edu/~enakamura/papers.html>.

of socially efficient unemployment in addition to a measure of unemployment consistent with stable prices. When there is a difference between these two measures, the Federal Reserve must face this trade-off earnestly. Saez added that the United States experienced a long period of relatively stable prices between the 1990s and 2020. In such periods, the Federal Reserve needs a measure of full employment to guide policy toward an efficient rate of unemployment.

To Hobijn's discussion on policies that shift the Beveridge curve, such as labor market institutions, Michaillat again pointed to the aforementioned work with Saez, which relaxes assumptions and derives a more complex formula that allows for the Beveridge curve to change shape over time. He acknowledged that unemployment insurance would move the Beveridge curve, complicating the analysis. Michaillat referred to his work with Saez and Camille Landais that considers the case of unemployment insurance.¹¹

Oleg Itskhoki expressed surprise that the authors' measure of optimal unemployment did not require information about the nature of underlying shocks and frictions in the economy. He added that while much of the discussion has been framed around monetary policy, this is the relevant policy tool only if the underlying frictions are sticky prices or sticky wages. If unemployment was instead generated for other reasons, such as financial shocks or other structural issues, monetary policy would not be able to resolve these problems, he argued. In this case, the knowledge of u^* might be useful, but ultimately policymakers need to understand the frictions and shocks in the background to respond appropriately.

In response, Michaillat described how their model brings the concept of sufficient statistics, often used in public finance, to a macroeconomic setting. This allows the analysis to be consistent with many macroeconomic models irrespective of background shocks, so long as the model contains a Beveridge curve and so long as the Federal Reserve does not create other distortions.

John Haltiwanger appreciated that the authors' framework relied on the DMP model but also suggested it needs to be refined to match the data. In particular, he asserted that the standard matching function that specifies hires as a function of vacancies and unemployment does not perform well in matching the job-filling rate and the job-finding rate. Hence, he proposed that the authors consider a generalized matching function. He pointed

11. Camille Landais, Pascal Michaillat, and Emmanuel Saez, "A Macroeconomic Approach to Optimal Unemployment Insurance: Theory," *American Economic Journal: Economic Policy* 10, no. 2 (2018): 152–81.

to papers by Abraham, Haltiwanger, and Rendell (2020) and Hall and Schulhofer-Wohl (2018) that generalize the matching function to recognize heterogeneity among vacancies.¹² For example, recruiting and intensities vary across vacancies and across time, hires come from more than just the unemployed, and job search intensities vary dramatically within the unemployed.

12. Katharine G. Abraham, John C. Haltiwanger, and Lea E. Rendell, “How Tight Is the US Labor Market?” *Brookings Papers on Economic Activity*, Spring (2020): 97–138; Robert E. Hall and Sam Schulhofer-Wohl, “Measuring Job-Finding Rates and Matching Efficiency with Heterogeneous Job-Seekers,” *American Economic Journal: Macroeconomics* 10, no. 1 (2018): 1–32.